

# Noise sensitivity questions in percolation-like models

Gábor Pete (Rényi Institute and TU Budapest)  
<http://www.math.bme.hu/~gabor>

Joint works with Zolt Bartha (UC Berkeley)  
and Pál Galicza (Central European University, Budapest)  
and Christophe Garban (U Lyon)

# Noise sensitivity of Boolean functions

$f : \{-1, 1\}^N \rightarrow \{-1, 1\}$  a **Boolean function**, usually monotone.  
Input is i.i.d. Bernoulli( $p$ ).

Take **critical density**  $p = p_c(N)$ , where  $\mathbf{P}_p[f(\omega) = 1] = 1/2$ .

**Resample** each input bit with probability  $\epsilon$ , independently, get  $\omega^\epsilon$ .

Given a typical  $\omega$ , can we predict what  $f(\omega^\epsilon)$  will be? What is the **correlation** between  $f(\omega^\epsilon)$  and  $f(\omega)$ ? Three simple examples:

**Dictator** $_n(\omega_1, \dots, \omega_n) := \omega_1$ . Here  $p_c(n) = 1/2$ .

Here  $\text{Corr}[\text{Dic}_n(\omega), \text{Dic}_n(\omega^\epsilon)] = 1 - \epsilon$ , hence noise-stable.

**Majority** $_n(\omega_1, \dots, \omega_n) := \text{sgn}(\omega_1 + \dots + \omega_n)$ . Again,  $p_c(n) = 1/2$ .

Here  $\text{Corr}[\text{Maj}_n(\omega), \text{Maj}_n(\omega^\epsilon)] = 1 - O(\sqrt{\epsilon})$ , hence noise-stable.

**Parity** $_n(\omega_1, \dots, \omega_n) := \omega_1 \cdots \omega_n$ . Again,  $p_c(n) = 1/2$ .

Here  $\text{Corr}[\text{Par}_n(\omega), \text{Par}_n(\omega^\epsilon)] = (1 - \epsilon)^n$ , very sensitive to noise.

## Noise sensitivity of Boolean functions

A sequence of Boolean functions  $f_k : \{-1, 1\}^{N_k} \rightarrow \{-1, 1\}$  is called **noise sensitive** at density  $p$  if

$$\forall \epsilon > 0 : \text{Corr}[f_k(\omega^\epsilon), f_k(\omega)] \rightarrow 0 \quad \text{as } k \rightarrow \infty,$$

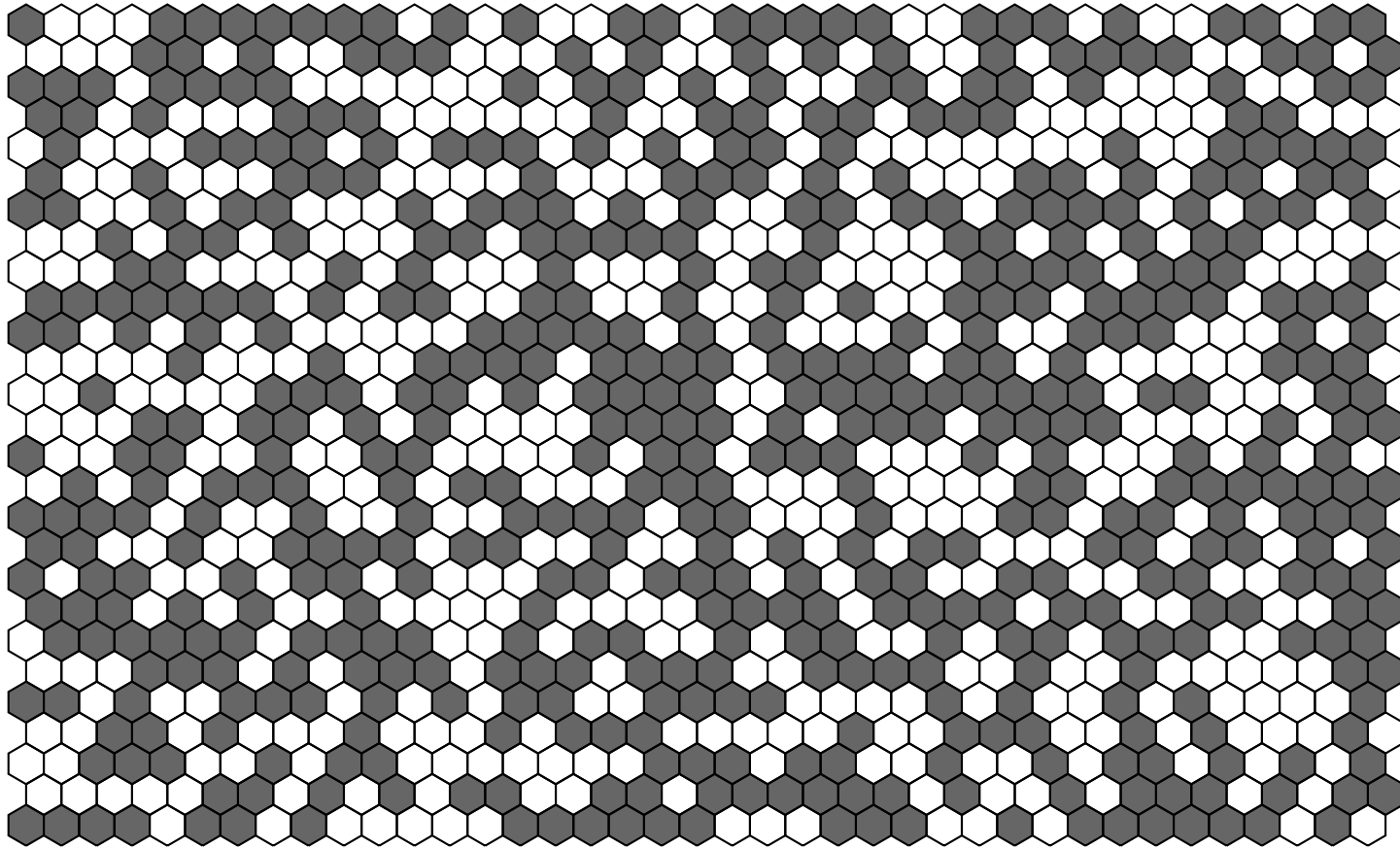
and **noise stable** if

$$\lim_{\epsilon \rightarrow 0} \sup_k \mathbf{P}_p[f_k(\omega^\epsilon) \neq f_k(\omega)] = 0.$$

Could be insensitive but not stable.

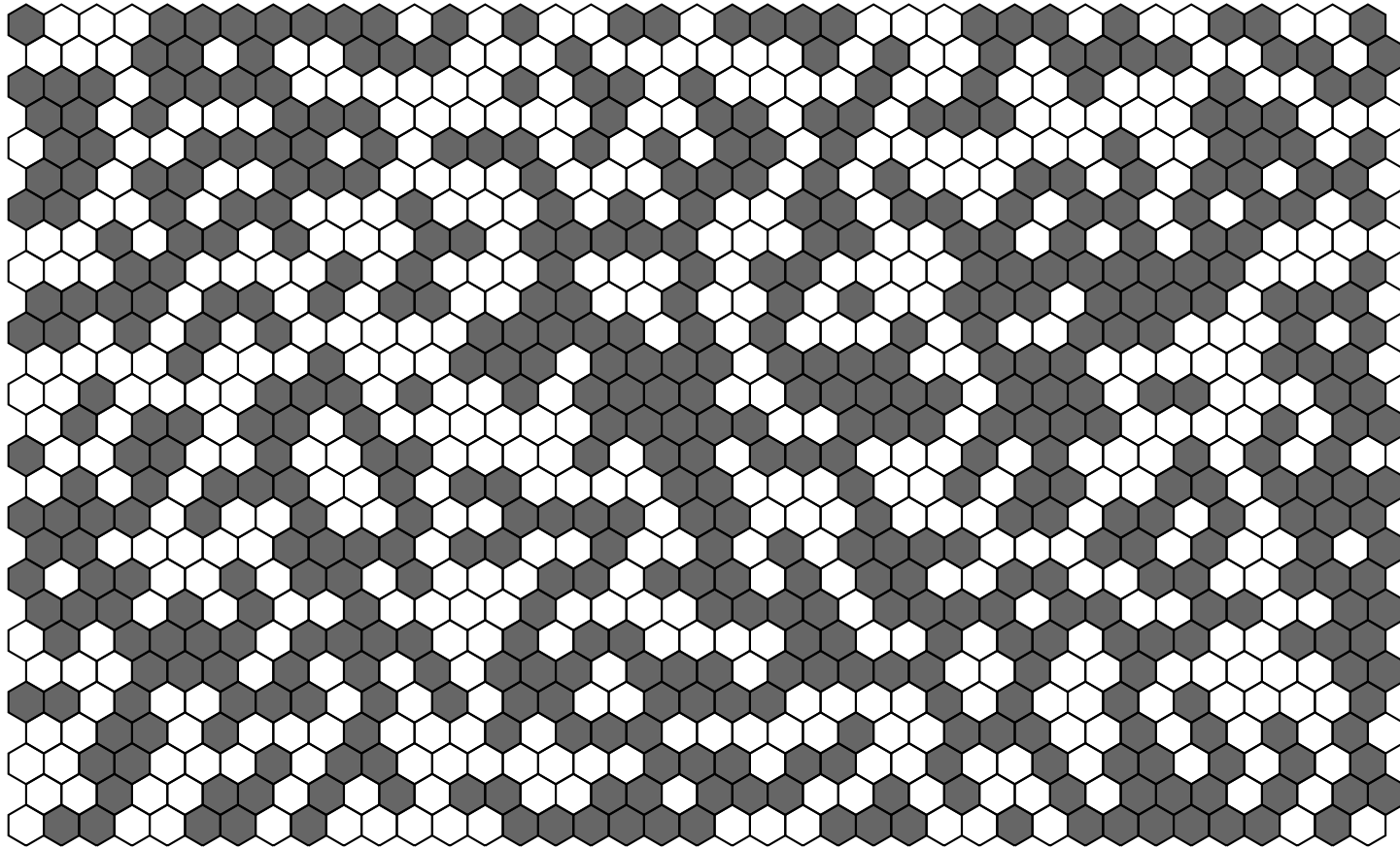
Give a monotone noise sensitive example!

## Percolation and noise



At  $p_c = 1/2$ , **left-right crossing** has non-trivial probability.

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At  $p_c = 1/2$ , **left-right crossing** has non-trivial probability.

## Naive idea: how many pivotals are there?

A bit is **pivotal** for  $f$  in  $\omega$  if flipping it changes the output.

Do the  $\epsilon$ -noise by switching bits one-by-one. In order to change the output, need at least one **pivotal switch**; in fact, need an odd number of them.

Complete decorrelation  $\Leftrightarrow$  so many pivotal switches that you don't know their parity.

Naively, “the more pivotals there are, the more noise sensitive the function should be”.

First issue: **Maj**<sub>2k+1</sub> typically has no pivotal bits at all; with probability  $\asymp 1/\sqrt{k}$ , it has  $k + 1$ , hence  $\mathbf{E}|\text{Piv}_{\text{Maj}_{2k+1}}| \asymp \sqrt{k}$ . This matters for **sharp thresholds** by the Russo-Margulis formula,

$$\frac{d}{dp} \mathbf{P}_p [ f(\omega) = 1 ] = \mathbf{E}_p [ |\text{Piv}_f| ],$$

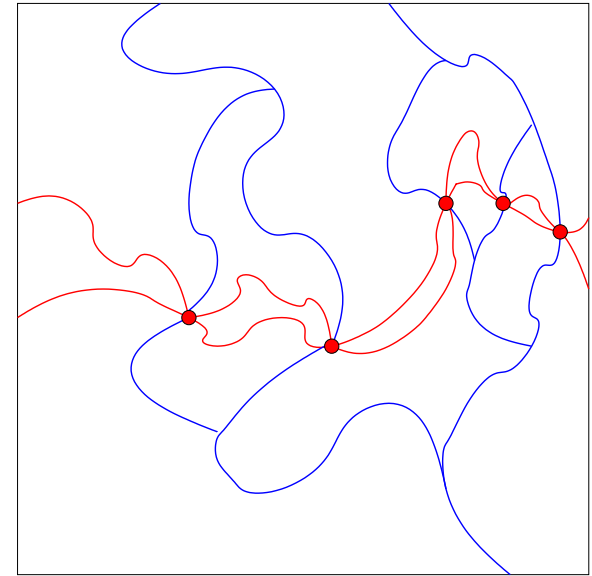
but **apparently not for noise sensitivity**.

## Naive idea: how many pivotals are there?

A site is **pivotal** for left-right crossing in  $\omega$  if it has the **alternating 4-arm event** to the sides.  $\mathbf{E}|\mathbf{Piv}_n| \asymp n^2 \alpha_4(n) \quad (= n^{3/4+o(1)})$ .

Furthermore,  $\mathbf{E}[|\mathbf{Piv}_n|^2] \leq C (\mathbf{E}|\mathbf{Piv}_n|)^2$ .  
So,  $\mathbf{P}[|\mathbf{Piv}_n| > \lambda \mathbf{E}|\mathbf{Piv}_n|] < C/\lambda^2$ , any  $\lambda$ .

And not only  $\exists \epsilon \mathbf{P}[|\mathbf{Piv}_n| > \epsilon \mathbf{E}|\mathbf{Piv}_n|] > \epsilon$ ,  
but  $\mathbf{P}[0 < |\mathbf{Piv}_n| < \epsilon \mathbf{E}|\mathbf{Piv}_n|] \asymp \epsilon^{11/9+o(1)}$ ,  
as  $\epsilon \rightarrow 0$  (exponent only for  $\Delta$ ).



If  $\epsilon_n \mathbf{E}[|\mathbf{Piv}_n|] \rightarrow 0$ , then  $\mathbf{E}[\text{number of pivotal switches}] \rightarrow 0$   
 $\implies$  asymptotically full correlation

If  $\epsilon_n \mathbf{E}[|\mathbf{Piv}_n|] \rightarrow \infty$ , then  $\mathbf{E}[\text{number of pivotal switches}] \rightarrow \infty$   
 $\not\implies \mathbf{P}[\text{hit (many) pivotals}] \rightarrow 1 \not\implies$  asymptotic independence!!

## Dynamical 2nd Moment Method for pivotal switches

Make  $m = t/\alpha_4(n)$  switches,  $\omega = \omega_0, \dots, \omega_m$ , so that

$$\mathbf{E}[S_t] = \mathbf{E}[\text{number of pivotal switches}] \asymp t.$$

Want to show  $\mathbf{P}[S_t = \text{odd}] > c_t > 0$ , uniformly in  $n$ , because then  $\text{Corr}[f(\omega), f(\omega_m)] < 1 - \tilde{c}_t$ .

Will prove  $\mathbf{E}[S_t^2 - S_t] = O(t^2)$ .

Then note  $\mathbf{E}[S_t^2] \geq \mathbf{P}[S_t = 1] + 2(\mathbf{E}S_t - \mathbf{P}[S_t = 1])$ . Rearranging gives  $\mathbf{P}[S_t = 1] \geq 2\mathbf{E}[S_t] - \mathbf{E}[S_t^2] \asymp t - O(t^2) > 0$  for  $t > 0$  small enough.

Also done for all  $t$ , since correlation is monotone decreasing in  $t$ .

And the second moment calculation:



$$\begin{aligned}
\mathbf{E}[S_t^2 - S_t] &\asymp \sum_{\substack{i,j=1 \\ i \neq j}}^m \sum_{\substack{x,y \in V_n \\ x \neq y}} n^{-4} \mathbf{P}[x \in \text{Piv}(\omega_i); y \in \text{Piv}(\omega_j)] \\
&\leq n^{-4} \sum_{i,j=1}^m \sum_{r=0}^{\lceil \log_2 n \rceil} \sum_{\substack{x,y \\ 2^r \leq d(x,y) < 2^{r+1}}} \mathbf{P}[A_x^r(\omega_i)] \mathbf{P}[B_y^r(\omega_j) \mid A_x^r(\omega_i)] \\
&\leq n^{-4} m^2 n^2 \sum_{r=0}^{\lceil \log_2 n \rceil} O(1) 2^{2r} \alpha_4(1, 2^r)^2 \alpha_4(2^r, n) \\
&\leq n^{-2} m^2 \alpha_4(n) \sum_{r=0}^{\lceil \log_2 n \rceil} O(1) 2^{2r} \alpha_4(2^r), \text{ recall } k^2 \alpha_4(k) = k^{3/4+o(1)}, \\
&\asymp O(1) \alpha_4(n)^2 m^2 = O(t^2),
\end{aligned}$$

where

$A_x^r(\omega_i) := \{\text{alternating 4 arms in } A_x(1, 2^{r-1}) \text{ and in } A_x(2^{r+2}, n) \text{ in } \omega_i\},$

$B_y^r(\omega_j) := \{\text{alternating 4 arms in } A_y(1, 2^{r-1}) \text{ in } \omega_j\}.$

## Same for weakly dependent input?

This was a robust argument! **Critical FK-Ising** and **critical spin-Ising on  $\mathbb{Z}^2$**  also satisfy

$$\mathbf{P} [ B_y^r(\omega_j) \mid A_x^r(\omega_i) ] \asymp \alpha_4(1, 2^r),$$

for their natural Glauber / heat-bath / Gibbs sampler dynamics.

Moreover, the exponents  $\alpha_4^{\text{FK-Ising}}(n) = n^{-35/24+o(1)}$  and  $\alpha_4^{\text{spin-Ising}}(n) = n^{-21/8+o(1)}$  are known (**Chelkak, Duminil-Copin, Hongler, Garban**).

**For FK-Ising, because of  $35/24 < 2$** , which means many pivotal points in the discrete world and self-touches of SLE(16/3) in the continuum, the previous argument works fine. Hence  $t n^{-13/24+o(1)}$  is the good space-time scaling to watch macroscopic connections **start changing**, and **Garban-P (2015+)** proves that there is a **Markovian scaling limit of the dynamics**.

**For spin-Ising Glauber dynamics, because of  $21/8 > 2$** , we do not know **anything**. Btw, the **mixing time of the entire system** is known to be polynomial (**Lubetzky-Sly 2010**), but the exponent is not known.

# Noise sensitivity of percolation

All results use **Fourier analysis of Boolean functions**:

**Theorem (Benjamini, Kalai & Schramm 1998)**. If  $\epsilon > 0$  is fixed, and  $f_n$  is the indicator function for a left-right percolation crossing in an  $n \times n$  square, then as  $n \rightarrow \infty$

$$\mathbf{E} [ f_n(\omega) f_n(\omega^\epsilon) ] - \mathbf{E} [ f_n(\omega) ]^2 \rightarrow 0.$$

This holds for all  $\epsilon = \epsilon_n > c / \log n$ .

**Theorem (Schramm & Steif 2005)**. Same if  $\epsilon_n > n^{-a}$  for some positive  $a > 0$ . If triangular lattice, may take any  $a < 1/8$ .

**Theorem (Garban, P & Schramm 2008)**. Same holds if and only if  $\epsilon_n \mathbf{E} [ |\text{Piv}_n| ] \rightarrow \infty$ . For triangular lattice, this threshold is  $\epsilon_n = n^{-3/4+o(1)}$ .

## What is the Fourier spectrum and why is it useful?

$f_n : \{\pm 1\}^{V_n} \longrightarrow \{\pm 1\}$  indicator of left-right crossing,  $V = V_n$  vertices.

$(N_\epsilon f)(\omega) := \mathbf{E}[f(\omega^\epsilon) \mid \omega]$  is the **noise operator**, acting on the space  $L^2(\Omega, \mu)$ , where  $\Omega = \{\pm 1\}^V$ ,  $\mu$  uniform measure, inner product  $\mathbf{E}[fg]$ .

**Correlation:**  $\mathbf{E}[f(\omega^\epsilon)f(\omega)] - \mathbf{E}[f(\omega)]\mathbf{E}[f(\omega^\epsilon)] = \mathbf{E}[f(\omega)N_\epsilon f(\omega)] - \mathbf{E}[f(\omega)]^2$ . So, we would like to **diagonalize** the noise operator  $N_\epsilon$ .

Let  $\chi_i$  be the function  $\chi_i(\omega) = \omega(i)$ ,  $\omega \in \Omega$ .

For  $S \subset V$ , let  $\chi_S := \prod_{i \in S} \chi_i$ , the **parity inside  $S$** . Then

$$N_\epsilon \chi_i = (1 - \epsilon) \chi_i; \quad N_\epsilon \chi_S = (1 - \epsilon)^{|S|} \chi_S.$$

Moreover, the family  $\{\chi_S, S \subseteq V\}$  is an **orthonormal basis** of  $L^2(\Omega, \mu)$ .

Any function  $f \in L^2(\Omega, \mu)$  in this basis (**Fourier-Walsh series**):

$$\hat{f}(S) := \mathbf{E}[f\chi_S]; \quad f = \sum_{S \subseteq V} \hat{f}(S) \chi_S.$$

The correlation:

$$\begin{aligned} \mathbf{E}[fN_\epsilon f] - \mathbf{E}[f]^2 &= \sum_S \sum_{S'} \hat{f}(S) \hat{f}(S') \mathbf{E}[\chi_S N_\epsilon \chi_{S'}] - \mathbf{E}[f\chi_\emptyset]^2 \\ &= \sum_{\emptyset \neq S \subseteq V} \hat{f}(S)^2 (1 - \epsilon)^{|S|} = \sum_{k=1}^{|V_n|} (1 - \epsilon)^k \sum_{|S|=k} \hat{f}(S)^2. \end{aligned}$$

By Parseval,  $\sum_S \hat{f}(S)^2 = \mathbf{E}[f^2] = 1$ . So can define probability measure  $\mathbf{P}[\mathcal{S}_f = S] := \hat{f}(S)^2 / \mathbf{E}[f^2]$ , the **spectral sample**  $\mathcal{S}_f \subseteq V$ .

If, for some functions  $f_n$  and numbers  $k_n$ , we have  $\mathbf{P}[0 < |\mathcal{S}_n| < tk_n] \rightarrow 0$  as  $t \rightarrow 0$ , uniformly in  $n$ , then  $(1 - \epsilon)^k \sim \exp(-\epsilon k)$  implies that for  $\epsilon_n \gg 1/k_n$  we have **asymptotic independence**. Maybe with  $k_n = \mathbf{E}|\mathcal{S}_n|$ ?

## Pivotal versus spectral sample

$\nabla_i f(\omega) := f(\sigma_i(\omega)) - f(\omega) \in \{-2, 0, +2\}$  gradient.

$\nabla_i f(\omega) = \sum_S \hat{f}(S) [\chi_S(\sigma_i(\omega)) - \chi_S(\omega)]$ , hence  $\widehat{\nabla_i f}(S) = -2\hat{f}(S)\mathbf{1}_{i \in S}$ .

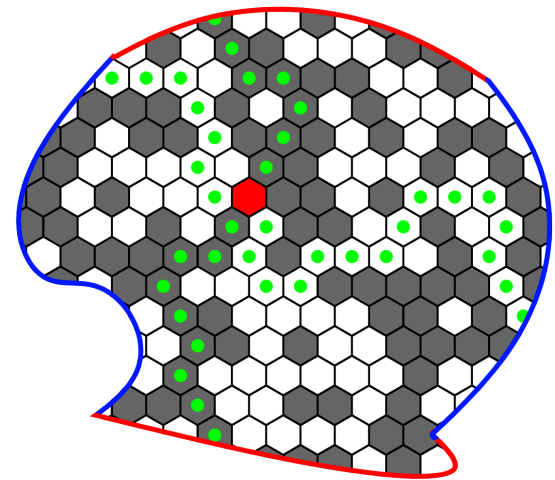
$\mathbf{P}[i \in \text{Piv}_f] = \frac{1}{4} \|\nabla_i f\|_2^2 = \frac{1}{4} \sum_S \widehat{\nabla_i f}(S)^2 = \sum_{S \ni i} \hat{f}(S)^2 = \mathbf{P}[i \in \mathcal{S}_f]$ .

Thus,  $\mathbf{E}|\mathcal{S}_f| = \mathbf{E}|\text{Piv}_f|$ . So, the pivotal upper bound for noise sensitivity is sharp if there is tightness around  $\mathbf{E}|\mathcal{S}|$ .

Alos,  $\mathbf{P}[i, j \in \text{Piv}_f] = \mathbf{P}[i, j \in \mathcal{S}_f]$ , hence  $\mathbf{E}|\mathcal{S}_f|^2 = \mathbf{E}|\text{Piv}_f|^2$ .

Not for more points and higher moments!  
Both random subsets measure the “influence” or “relevance” of bits, but in different ways.

For percolation,  $\mathbf{E}[|\text{Piv}_n|^2] \leq C (\mathbf{E}|\text{Piv}_n|)^2$ , hence  $\exists c > 0$  s.t.  $\mathbf{P}[|\mathcal{S}_n| > c\mathbf{E}|\mathcal{S}_n|] > c$ . That's why one hopes for tightness around mean.



## The earlier simple examples

**Dictator** $_n(\omega_1, \dots, \omega_n) := \omega_1$ . Here  $p_c(n) = 1/2$ .

Here  $\text{Corr}[\text{Dic}_n(\omega), \text{Dic}_n(\omega^\epsilon)] = 1 - \epsilon$ , hence noise-stable. And

$\mathbf{P}[\mathcal{S}_n = \{x_1\}] = 1$ .

**Majority** $_n(\omega_1, \dots, \omega_n) := \text{sgn}(\omega_1 + \dots + \omega_n)$ . Again,  $p_c(n) = 1/2$ .

Here  $\text{Corr}[\text{Maj}_n(\omega), \text{Maj}_n(\omega^\epsilon)] = 1 - O(\sqrt{\epsilon})$ , hence noise-stable.

And  $\mathbf{P}[\mathcal{S}_n = \{x_i\}] \asymp 1/n$ , most of the weight is on singletons.

On the other hand,  $\mathbf{E}|\mathcal{S}_n| = \mathbf{E}|\text{Piv}_n| \asymp \frac{1}{\sqrt{n}} n \asymp \sqrt{n}$ .

**Parity** $_n(\omega_1, \dots, \omega_n) := \omega_1 \cdots \omega_n$ . Again,  $p_c(n) = 1/2$ .

Here  $\text{Corr}[\text{Par}_n(\omega), \text{Par}_n(\omega^\epsilon)] = (1 - \epsilon)^n$ , very sensitive to noise.

And  $\mathbf{P}[\mathcal{S}_n = \{x_1, \dots, x_n\}] = 1$ .

## Benjamini, Kalai & Schramm 1998

Using hypercontractivity of the noise operator  $N_\epsilon$ :

**Theorem 1.** A sequence  $f_n$  of monotone Boolean functions is **noise sensitive**, iff it is asymptotically uncorrelated with all **weighted majorities**  $\text{Maj}_w(\omega_1, \dots, \omega_n) = \text{sign} \sum_{i=1}^n \omega_i \omega_i$ .

**Theorem 2.** A sequence  $f_n$  of monotone Boolean functions is **noise sensitive** at density  $p = 1/2$ , iff  $\sum_x \mathbf{P}[x \in \text{Piv}_n]^2 \rightarrow 0$  as  $n \rightarrow \infty$ .

Extended to  $p_c(n) \asymp 1/\text{poly}(\log n)$  by **Keller-Kindler '10** and **Bouyrie '14**.

**Corollary.** The left-right percolation crossing in an  $n \times n$  square is noise sensitive, even with  $\epsilon = \epsilon_n > c/\log n$ .



## Schramm & Steif 2005

**Theorem.** If  $f : \Omega \rightarrow \mathbb{R}$  can be computed with a randomized algorithm with **revelment**  $\delta$  (each bit is read only with probability  $\leq \delta$ ), then

$$\sum_{S:|S|=k} \hat{f}(S)^2 \leq \delta k \|f\|_2^2.$$

For left-right crossing in  $n \times n$  box on the hexagonal lattice, **exploration interface** with random starting point gives revelment  $n^{-1/4+o(1)}$  (it has length  $n^{7/4+o(1)}$ , given by 2-arm exponent), while  $\sum_{k \leq m} k \asymp m^2$ , thus:

**Corollary.** Left-right crossing on the triangular lattice is noise sensitive under  $\epsilon_n > n^{-a}$ , with any  $a < 1/8$ . Even on square lattice, can take some positive  $a > 0$ .

The revelment is at least  $n^{-1/2+o(1)}$  for *any* algorithm computing the crossing (**O'Donnell-Servedio '07**, plus  $n^{3/4}$  pivots), hence this method can give only  $n^{-1/4+o(1)}$ -sensitivity, far from the conjectured  $\epsilon_n = n^{-3/4+o(1)}$ .

## Garban, P. & Schramm 2008

Although goal is to understand size, Gil Kalai suggested trying to understand entire distribution of  $\mathcal{S}_f$ . A strange random set of bits.

[Smirnov '01] + [Tsirelson '04] + [Schramm-Smirnov '10] implies that it has a **conformally invariant scaling limit**.

How to prove tightness for the size of strange random fractal-like sets?

$$\text{For } A \subseteq V: \mathbf{E}[\chi_S \mid \mathcal{F}_A] = \begin{cases} \chi_S & S \subseteq A, \\ 0 & \text{otherwise.} \end{cases}$$

Therefore,  $\mathbf{E}[f \mid \mathcal{F}_A] = \sum_{S \subseteq A} \hat{f}(S) \chi_S$ , a nice projection.

$$\mathbf{P}[\mathcal{S}_f \subset U] = \sum_{S \subset U} \hat{f}(S)^2 = \mathbf{E} \left[ \left( \sum_{S \subset U} \hat{f}(S) \chi_S \right)^2 \right] = \mathbf{E} \left[ \mathbf{E}[f \mid \mathcal{F}_U]^2 \right].$$

From this, for disjoint subsets  $A$  and  $B$ , can try to give percolation meaning to  $\mathbf{P}[\mathcal{S}_f \cap B \neq \emptyset = \mathcal{S}_f \cap A]$ . **Very restricted independence**.

## Influence notions of subsets

Besides resampling a random subset  $U_\epsilon$ , could also use any fixed subset  $U$ .

**Influence:**

$$I(U) := \mathbf{P}[U \text{ is pivotal}] = \mathbf{P}[U^c \text{ does not determine the value of } f].$$

**Significance:** the amount of information missing if the bits in  $U^c$  are known,

$$\text{sig}(U) := \frac{\mathbf{E}[\text{Var}[f \mid \mathcal{F}_{U^c}]]}{\text{Var} f} = \frac{\mathbf{P}[\mathcal{S} \cap U \neq \emptyset]}{\text{Var} f}.$$

**Clue:** the amount of information we gain from the bits of  $U$ ,

$$\text{clue}(U) := \frac{\text{Var}[\mathbf{E}[f \mid \mathcal{F}_U]]}{\text{Var} f} = \frac{\mathbf{P}[\emptyset \neq \mathcal{S} \subseteq U]}{\text{Var} f}.$$

Clearly,  $I(U) \geq \text{sig}(U) \geq \text{clue}(U)$ . Also,  $\text{clue}(U) = 1 - \text{sig}(U^c)$ .

## Influence notions of subsets

A few examples:

$U_n$  is all the vertical edges in  $\mathbb{Z}^2$  bond percolation.

Then  $\text{sig}(U_n) \rightarrow 1$  and  $\text{clue}(U_n) \rightarrow 0$ .

$U_n$  has a scaling limit of Hausdorff-dimension  $\gamma$ :

If  $\gamma < 5/4$ , then  $\text{sig}(U_n) \rightarrow 0$ , or  $\text{clue}(U_n^c) \rightarrow 1$ .

If  $5/4 < \gamma < 2$ , then usually  $\text{sig}(U_n) \rightarrow 1$ , but also  $\text{sig}(U_n^c) \rightarrow 1$ , hence  $\text{clue}(U_n) \rightarrow 0$ .

For Majority $_n$ , any subset  $U$  of size  $\epsilon n$  has  $\text{sig}(U) \asymp \sqrt{\epsilon}$ .

**Benjamini:** does  $|U_n| = o(n^2)$  imply also for percolation that  $\text{clue}(U_n) \rightarrow 0$ ?

If we can choose the revealed bits *adaptively*, then, using the exploration interface,  $n^{7/4+o(1)}$  bits suffice.

## Small subsets are clueless (Galicza & P.)

If  $f : \{\pm 1\}^V \rightarrow \{\pm 1\}$  is *transitive*, and  $U \subset V$ , then

$$\text{clue}(U) \cdot \text{Var} f = \mathbf{P}[\emptyset \neq \mathcal{S} \subseteq U] \leq \mathbf{P}[X \in U] = \sum_{u \in U} \mathbf{P}[X = u] = \frac{|U|}{|V|}.$$

Percolation left-right crossing is *not transitive*, but if we consider an  $n \times n$  **torus**, then the indicator  $F$  of the event that  $\left\{ \text{there is a translate of the square on the torus where the left-right crossing happens} \right\}$  is transitive.

Now  $F_\epsilon$ , the indicator of  $\left\{ \text{there is a good translate by some vector } (k, \ell) \in n, k, \ell \in \{0, 1, \dots, 1/\epsilon\} \right\}$  is a good approximation to  $F$ , because of low probability of half-plane 3-arm events etc.

If a small subset  $U$  had a clue about left-right crossing in a square, then its  $1/\epsilon \times 1/\epsilon$  translates, using FKG, would have a clue about  $F_\epsilon$ , so also about  $F$ , which is impossible, since  $1/\epsilon^2 |U|$  is still small.

## The clue of random subsets

It Ain't over till it's over (**Mossel, O'Donnell, Oleszkiewicz '10**). For any density  $0 < \rho < 1$  and  $\epsilon > 0$  there is a  $\delta$  and a  $\tau$  such that for every Boolean function  $f$  satisfying  $\mathbf{P}[i \in \text{Piv}_f] \leq \tau$  for all  $i \in V$ , we have

$$\mathbf{P}[\text{clue}(U_\rho) \geq 1 - \delta] < \epsilon.$$

This is a deep result, though obvious for noise sensitive functions. Our easy method gives the result when  $\sum_i \mathbf{P}[i \in \text{Piv}_f]^2 = o(1)$ , via an **Azuma-Hoeffding concentration argument**:

$$\mathbf{E}[\text{clue}(U_\rho)] \cdot \text{Var} f = \mathbf{P}[\emptyset \neq \mathcal{S} \subseteq U_\rho] \leq \mathbf{P}[X \in U_\rho] = \rho.$$

And if  $U$  and  $U'$  differ only in bit  $i$ , then

$$|\text{clue}(U) - \text{clue}(U')| \leq \mathbf{P}[i \in \mathcal{S}] = \mathbf{P}[i \in \text{Piv}].$$

Hence we get concentration around the mean on the scale  $\sum_i \mathbf{P}[i \in \text{Piv}]^2$ .

## Bootstrap percolation on infinite graphs

Infinite transitive graph. Start with Bernoulli( $p$ ) percolation of **occupied sites**. If a site has **at least  $k$**  occupied neighbors, then gets occupied. Repeat ad infinitum. For what  $p$  do we occupy the entire graph, with probability 1?

**van Enter '87, Schonmann '90.**  $p_c(\mathbb{Z}^d, k) = 0$  for  $k \leq d$ , and 1 for  $k > d$ .

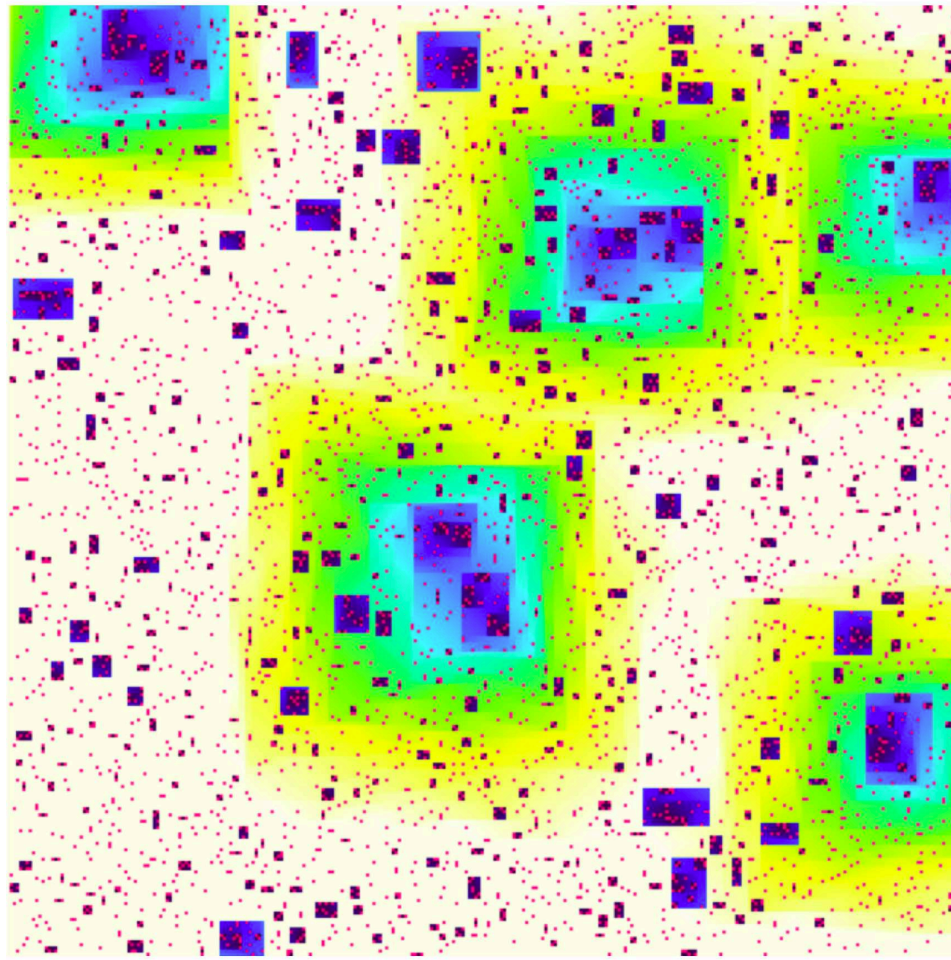
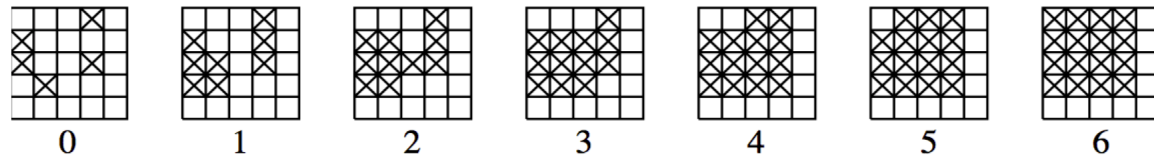
Probably also true for the Heisenberg group (observation by **Rob Morris**).

**Balogh, Peres & P. '06.** For any non-amenable group  $\Gamma$  that has a free subgroup  $F_2$ , there exists a symmetric finite generating set and a  $k$ -neighbor rule such that  $G = \text{Cayley}(\Gamma, S)$  has  $0 < p_c(G, k) < 1$ .

On  $d$ -regular tree  $p_c(T_d, k)$  can be explicitly calculated, since unsuccessful occupation is equivalent to having a  $d - k + 1$ -regular vacant subtree in the initial configuration.

**Question.** What about non-amenable graph without free subgroups? Conversely, on amenable transitive graphs, does  $p_c < 1$  imply  $p_c = 0$ ?

# Bootstrap percolation on infinite graphs





# Bootstrap percolation on finite graphs

**Aizenman, Lebowitz '88.**  $p_c(\mathbb{Z}_n^d, 2) \asymp \frac{1}{\log^{d-1} n}$ .

More general and sharper results by **Cerf, Cirillo, Manzo, Holroyd, Balogh, Bollobás, Morris, Duminil-Copin, . . . .**

**Balogh, Pittel '07.** For the random  $d$ -regular graph on  $n$  vertices,  $p_c(G_{n,d}, k) \rightarrow p_c(T_d, k)$  for any  $2 \leq k \leq d - 2$ . The threshold window around  $p_c(n)$  is of size  $O(1/\sqrt{n})$ .

**Bartha & P. '14.**

- (1) At  $p_c(\mathbb{Z}_n^d, 2)$ , the process is noise sensitive.
- (2) At  $p_c(G_{n,d}, k)$ , the process is noise insensitive.

## Rough sketch of sensitivity on $\mathbb{Z}_n^2$

Will use the condition that  $\sum_i \mathbf{P}[i \in \text{Piv}]^2$  is small.

Complete occupation is roughly equivalent to having an internally spanned rectangular **seed** of side lengths  $\asymp \log n$ , because this will keep growing.

Hence, again using an “almost transitivity” argument,

$$\mathbf{P}[i \in \text{Piv and } \omega_i = 1] \leq \frac{C \log^2 n}{n^2}.$$

And we have  $\mathbf{P}[i \in \text{Piv and } \omega_i = 1] \asymp \frac{1}{\log n} \mathbf{P}[i \in \text{Piv and } \omega_i = -1]$ . So, altogether,

$$\mathbf{P}[i \in \text{Piv}] \leq \frac{C \log^3 n}{n^2}.$$

Should hold for  $k > 2$ , too, but seeds are much more complicated.

Revelment does *not* work: for negative result, need to reveal a lot.

## Sketch of stability on $G_{n,d}$

The **very narrow threshold window**  $O(1/\sqrt{n})$  implies that it is correlated with a weighted majority: if **majority at level**  $p_c(n)$  is fulfilled, then with positive probability we **overshoot** by density  $O(1/\sqrt{n})$ , but that raises the probability of complete occupation noticeably.

But this used the Balogh-Pittel explicit ODE calculation, relying on the lot of independent randomness in  $G_{n,d}$ .

A more intuitive reason for (2) is that it should hold for many **expander graphs**: for  $2d$ -regular expanders, with  $k \geq d$ , by the **perimeter trick**, all positive witnesses have size at least  $cn$ , and this is often also true for the negative witnesses, **which already should imply insensitivity** — that would be a converse to the revelation method.

However, there are expanders where each vertex is part of large piece of a square lattice, hence  $p_c(G_n, 2) \rightarrow 0$ . I do not know whether to expect noise stability or not.