

Applications of Stochastics — Renewal Equations Sketch

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Let ξ_1, ξ_2, \dots be the i.i.d. lifetimes in a renewal process, with non-arithmetic distribution function $F(s) = \mathbf{P}[\xi \leq s]$ and mean $\mathbf{E}\xi = \mu \in (0, \infty)$. Then $T_k := \sum_{i=1}^k \xi_i$ are the renewal times, $N_t := \min\{k : T_k \geq t\}$, and $U(t) := \mathbf{E}N_t$. The excess lifetime (or overshoot) is $\gamma_t := T_{N_t} - t$, the current lifetime is $\delta_t := t - T_{N_t-1}$, and the total lifetime is $\beta_t := \gamma_t + \delta_t$.

▷ **Exercise 1.**

- (a) Find the renewal equation $H(t) = h(t) + H * F(t)$ for $H(t) := \mathbf{P}[\beta_t > x]$, where $x \geq 0$ is fixed arbitrarily. (We actually did this in class.)
- (b) Find the renewal equation for $H(t) := \mathbf{P}[\gamma_t > x]$.
- (c) Using the Renewal Theorem, find the limit distributions of β_t and γ_t as $t \rightarrow \infty$.
- (d) Identify the limit distribution of the total lifetime β_t as the size-biased version of ξ , and the limit distribution of the overshoot γ_t as the size-biased version $\hat{\xi}$ multiplied with an independent $\text{Unif}[0, 1]$ variable. In order to avoid working with Stieltjes-integrals, you may assume that ξ has a density function.

Solution.

(b) Try to write a “recursion” for the distribution of γ_t , i.e., a renewal equation.

Let $H_x(t) := \mathbf{P}[\gamma_t > x]$. Condition on T_1 .

$$\mathbf{P}[\gamma_t > x \mid T_1] = \begin{cases} H_x(t - T_1) & \text{if } T_1 < t, \\ 0 & \text{if } t < T_1 < t + x, \\ 1 & \text{if } t + x < T_1. \end{cases}$$

Now average over the possible values of T_1 , i.e., take expectation. Recall that the cumulative distribution function of T_1 is $F(s)$.

$$\begin{aligned} \mathbf{P}[\gamma_t > x] &= \mathbf{E}[\mathbf{P}[\gamma_t > x \mid T_1]] = \int_0^t H_x(t - s) dF(s) \\ &\quad + 0 \cdot \mathbf{P}[t < T_1 < t + x] \\ &\quad + 1 \cdot \mathbf{P}[t + x < T_1], \end{aligned}$$

or, with the Stieltjes-convolution notation:

$$H_x(t) = (H_x * F)(t) + 1 - F(t + x). \tag{1}$$

(c) The Renewal Theorem says that, if F has non-arithmetic distribution with finite mean μ , and the $h(t)$ term in the renewal equation is directly integrable (e.g., positive and integrable), then

$$\lim_{t \rightarrow \infty} H(t) = \frac{1}{\mu} \int_0^\infty h(s) ds.$$

In (1), we have $h(t) = 1 - F(t + x)$, which is positive, and

$$\int_0^\infty h(s) ds = \int_x^\infty 1 - F(s) ds.$$

This is less than the same integral from 0 to ∞ , which is just $\mathbf{E}[T_1] = \mu < \infty$, so this is indeed directly integrable, and

$$x \mapsto \frac{1}{\mu} \int_x^\infty 1 - F(s) ds \in [0, 1] \tag{2}$$

is indeed the tail of a distribution.

(d) What is the tail of the distribution of the size-biased version $\widehat{\xi}$ multiplied with an independent $U \sim \text{Unif}[0, 1]$ variable? By first conditioning on the value of $\widehat{\xi}$:

$$\mathbf{P}[\widehat{\xi} \cdot U > x] = \int_x^\infty \mathbf{P}[sU > x] dF_{\widehat{\xi}}(s) = \int_x^\infty 1 - \frac{x}{s} dF_{\widehat{\xi}}(s) = \frac{1}{\mu} \int_x^\infty (s - x) dF_{\xi}(s),$$

since $dF_{\widehat{\xi}}(s) = \frac{s}{\mu} dF_{\xi}(s)$ is the size-biasing. And this formula is actually the same as (2), by the following picture:

