

When the group means are unequal both T^2 and F increase as N_1 (and N_2) increases, while D^2 does not. Therefore D^2 is a better description of the distance between groups when distances are compared.

The above formulas assume that the data are available in each case for all the variables compared. If some data values are equal to missing value codes or are out of range, observations for all variables may not be present in each case. Then N_1 and N_2 are replaced by the harmonic means of the frequencies of variables of the first and second groups respectively; this provides an approximate test of the equality of means. The formula is given in Appendix A.3.

TEST no/prev.
 HOTELLING. no/prev.
 When HOTELLING is specified, Hotelling's T^2 and Mahalanobis D^2 are computed.

Example 3D.4 Restricting Analysis to Complete Cases

Usually an analysis of a single variable uses all the acceptable values for the variable whether or not the value of any other variable is acceptable. Conversely, the usual definition of Hotelling's T^2 requires that the data values be acceptable for all the variables for any case that is included in the computations.

Cases containing acceptable data values for all variables that are included in the analysis are called complete cases.

In P3D you can specify whether all the computations are to be based on COMPLETE cases only, or on all acceptable values. If COMPLETE is specified, the univariate statistics and the t statistics are also computed using only complete cases.

To illustrate the effect COMPLETE has on the results, we add the command COMPLETE to the TEST paragraph of Example 3D.3 as follows

 / TEST VARIABLES ARE CHOLSTRL, ALBUMIN,
 CALCIUM, URICACID.
 HOTELLING.
 COMPLETE.

The results for modified analysis are presented in Output 3D.4 and can be compared with the analysis in Output 3D.3. We do not display the results for ALBUMIN, CALCIUM and URICACID. The small differences between the two results are due to the fact that there are only eight cases containing unacceptable data. Note the different frequencies (sample sizes) and degrees of freedom used in the two analyses.

A large difference between the two analyses would indicate that the results may be biased due to the pattern of missing values or values out of range. If your analyses show a large difference, you may want to examine the data by using PAM (Section 12.2) to study the pattern of values excluded from the analysis.

TEST no/prev.
 COMPLETE. no/prev.
 When COMPLETE is specified, only complete cases are used in all the computations. Complete cases are cases in which the data are acceptable (not missing or out of range) for all variables specified in the USE statement of the VARIABLE paragraph (all variables if USE is not specified). COMPLETE or NO COMPLETE can be specified in only the first TEST paragraph of any problem. It cannot be altered until a new problem begins.

Example 3D.5 Correlation of Variables in Each Group

The CORRELATIONS between the variables in each group are printed when CORRELATION is specified in the TEST paragraph. If we submit the Control Language of Example 3D.1 with the added TEST paragraph

 / TEST VARIABLES ARE CHOLSTRL, ALBUMIN,
 CALCIUM, URICACID.
 CORRELATIONS.

we obtain the results shown in Output 3D.5.

Output 3D.5 Correlation matrices for each group

 TEST TITLE. WERNER BLOOD CHEMISTRY DATA
 INDEXES OF VARIABLES TO BE ANALYZED 6 7 8 9
 USE COMPLETE CASES ONLY NO
 PRINT GROUP CORRELATION MATRICES. YES
 COMPUTE HOTELLINGS T SQUARE NO
 INDEX OF GROUPING VARIABLE. 5

GROUPS USED IN COMPUTATIONS 1 2

CORRELATION MATRIX FOR GROUP 1 MOPILL

	CHOLSTRL 6	ALBUMIN 7	CALCIUM 8	URICACID 9
CHOLSTRL 6	1.0000			
ALBUMIN 7	0.0296	1.0000		
CALCIUM 8	0.2874	0.4452	1.0000	
URICACID 9	0.2739	0.0858	0.2009	1.0000

CORRELATION MATRIX FOR GROUP 2 PILL

	CHOLSTRL 6	ALBUMIN 7	CALCIUM 8	URICACID 9
CHOLSTRL 6	1.0000			
ALBUMIN 7	0.1160	1.0000		
CALCIUM 8	0.2153	0.4258	1.0000	
URICACID 9	0.2473	-0.0485	0.1916	1.0000

----- analyses of variables 6 to 9 as in Output 3D.1 -----

```

/ TEST   VARIABLES ARE CHOLSTRL, ALBUMIN,
        CALCIUM, URICACID.
        GROUPS ARE 1, 4.

/ END
    
```

These instructions are similar to Example 1D.1 except we now use AGE as the GROUPING variable; AGE is used to classify the cases into four groups. The results are presented in Output 3D.2.

Two TEST paragraphs are used. The first paragraph specifies that only the variable CHOLSTRL is to be analyzed and that three groups (with subscripts 1, 2 and 3) are to be compared. These groups are named '25ORLESS', '26 TO 35' and '36 TO 45'. There are three possible pairings of three groups, therefore three analyses of CHOLSTRL are computed, each using a different pair of groups.

The second TEST paragraph specifies that two groups (with subscripts 1 and 4) are to be compared using the data from the four blood chemistry measurements. This analysis follows the interpretation of the second TEST paragraph. We omit

the results for CALCIUM and URICACID.
 A title can be specified in each TEST paragraph to label the analysis.

```

TEST
Variable = v list.      all variables except the
                        GROUPING variable/prev.
Names or subscripts of the VARIABLES to be
analyzed. When USE is stated in the VARIABLE
paragraph, VARIABLES in the TEST paragraph must
be included in the USE list.
GROUP = g list.        all groups/prev.
The groups to be compared. Lists are the GROUP
NAMES or group subscripts. A group subscript is
the sequence number of the group in the list of
CODES or CUTPOINTS specified in the GROUP
paragraph, or, if not specified in a GROUP
paragraph, the rank order of the group. If more
than two GROUPS are specified, each possible
pair of groups is compared.
TITLE = 'c'.           < 80 char.           blank
A title for the analysis.
    
```

Output 3D.2 A subset of groups and variables are selected for analysis by P3D

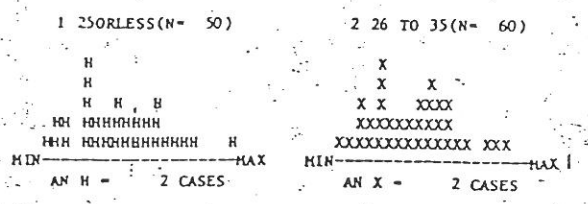
```

TEST TITLE. . . . . VERNER BLOOD CHEMISTRY DATA
INDEXES OF VARIABLES TO BE ANALYZED . . . . . 6
USE COMPLETE CASES ONLY . . . . . NO
PRINT GROUP CORRELATION MATRICES. . . . . NO
COMPUTE HOTELLINGS T SQUARE . . . . . NO
INDEX OF GROUPING VARIABLE. . . . . 2
    
```

GROUPS USED IN COMPUTATIONS 1 2 3

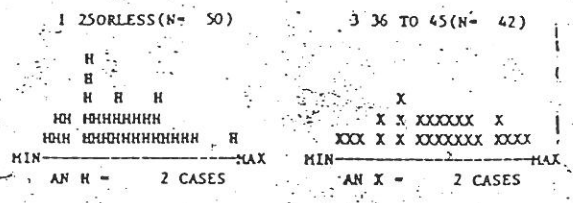
DIFFERENCES ON SINGLE VARIABLES

* CHOLSTRL *	VARIABLE NUMBER	6	GROUP	1 25ORLESS	2 26 TO 35
STATISTICS	P VALUE	DF	MEAN	222.1198	224.6331
			STD DEV	37.4441	35.7419
			S.E.M.	5.2954	4.6143
T (SEPARATE)	-0.36	0.721	SAMPLE SIZE	50	60
T (POOLED)	-0.36	0.720	MAXIMUM	330.0000	317.0000
F (FOR VARIANCES)			MINIMUM	155.0000	160.0000
LEVENE	0.01	0.912			



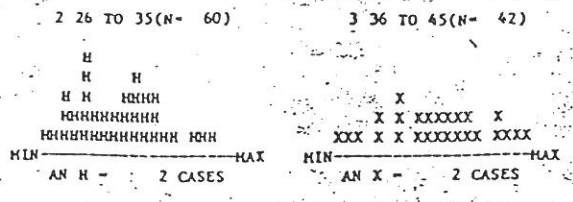
DIFFERENCES ON SINGLE VARIABLES

* CHOLSTRL *	VARIABLE NUMBER	6	GROUP	1 25ORLESS	3 36 TO 45
STATISTICS	P VALUE	DF	MEAN	222.1198	248.3332
			STD DEV	37.4441	44.8088
			S.E.M.	5.2954	6.9141
T (SEPARATE)	-3.01	0.003	SAMPLE SIZE	50	42
T (POOLED)	-3.06	0.003	MAXIMUM	330.0000	335.0000
F (FOR VARIANCES)			MINIMUM	155.0000	160.0000
LEVENE	1.87	0.174			



DIFFERENCES ON SINGLE VARIABLES

* CHOLSTRL *	VARIABLE NUMBER	6	GROUP	2 26 TO 35	3 36 TO 45
STATISTICS	P VALUE	DF	MEAN	224.6331	248.3332
			STD DEV	35.7419	44.8088
			S.E.M.	4.6143	6.9141
T (SEPARATE)	-2.85	0.006	SAMPLE SIZE	60	42
T (POOLED)	-2.97	0.004	MAXIMUM	317.0000	335.0000
F (FOR VARIANCES)			MINIMUM	160.0000	160.0000
LEVENE	2.60	0.110			



(output continued)

Example 3D.7 The One-Sample (or Matched Pairs) t Test

The Werner data (Table 5.1) consist of 94 pairs of age-matched women. In each pair the first woman is not on the pill and the second woman is. We perform a paired t test by reading each pair of data records as a single case. We then use BMDP transformations to form four new variables that represent differences in the blood measurement variables. The Control Language rules for the TRANSFORM paragraph are described in Chapter 6. Note that we state that we ADD four variables in the VARIABLE paragraph. We request a one-sample t test for each new variable (those representing differences) by not specifying a GROUPING variable. The Control Language is as follows

```

-----
/ PROBLEM  TITLE IS 'WERNER BLOOD CHEMISTRY DATA'.
/ INPUT    VARIABLES ARE 13.
           FORMAT IS '(A4, 5F4.0, 3F4.1/ 20X,
                     F4.0, 3F4.1)'.
    
```

```

/ VARIABLE NAMES ARE ID, AGE, HEIGHT, WEIGHT,
                    BRTHPILL, CHOL1, ALB1, CAL1,
                    URIC1, CHOL2, ALB2, CAL2,
                    URIC2, CHOLDIFF, ALBDIFF,
                    CALDIFF, URICDIFF.
MAXIMUMS ARE (6)400, (10)400.
MINIMUMS ARE (6)150, (10)150.
LABEL IS ID.
ADD IS 4.

/ TRANSFORM CHOLDIFF = CHOL1 - CHOL2.
           ALBDIFF  = ALB1  - ALB2.
           CALDIFF  = CAL1  - CAL2.
           URICDIFF = URIC1 - URIC2.

/ TEST    VARIABLES ARE 14 TO 17.
           HOTELLING.

/ END
-----
    
```

Output 3D.7 shows the results for CHOLDIFF and ALBDIFF; the results for ALBDIFF (the difference in

Output 3D.7 Paired t test by P3D

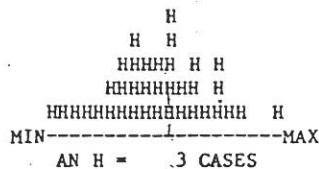
MAHALANOBIS D SQUARE	0.1303		
HOTELLING T SQUARE	11.8567		
F VALUE	2.8654	P VALUE	0.028
DEGREES OF FREEDOM	4,		87.0

WARNING - SINCE SPECIAL MISSING VALUE FORMULAS ARE USED,
THESE MULTIVARIATE STATISTICS ARE ONLY APPROXIMATE.

DIFFERENCES ON SINGLE VARIABLES

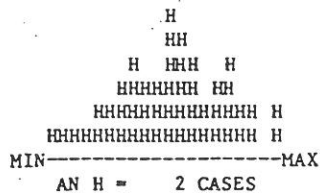
* CHOLDIFF * VARIABLE NUMBER 14

			MEAN	-6.1848	
T STATISTIC	P VALUE	DF	STD DEV	59.5390	
			S.E.M.	6.2074	
-1.00	0.322	91	SAMPLE SIZE	92	
			MAXIMUM	155.0000	
			MINIMUM	-145.0000	



* ALBDIFF * VARIABLE NUMBER 15

			MEAN	0.1804	
T STATISTIC	P VALUE	DF	STD DEV	0.5315	
			S.E.M.	0.0554	
3.26	0.002	91	SAMPLE SIZE	92	
			MAXIMUM	1.3000	
			MINIMUM	-1.2000	



--- similar analyses for CALDIFF and URICDIFF ---

Output 3D.3 Hotelling's T² and Mahalanobis D² - using all acceptable values

TEST TITLE. WERNER BLOOD CHEMISTRY DATA
INDEXES OF VARIABLES TO BE ANALYZED 6 7 8 9
USE COMPLETE CASES ONLY NO
PRINT GROUP CORRELATION MATRICES. NO
COMPUTE HOTELLINGS T SQUARE YES
INDEX OF GROUPING VARIABLE. 5

GROUPS USED IN COMPUTATIONS 1 2

DIFFERENCES AMONG GROUP MEANS USING ALL VARIABLES FOR THE FOLLOWING GROUPS

* NOPYLL *
* PILL *

MAHALANOBIS D SQUARE 0.2819
HOTELLING T SQUARE 13.0364
F VALUE 3.2057 P VALUE 0.014
DEGREES OF FREEDOM 4, 180.0

WARNING - SINCE SPECIAL MISSING VALUE FORMULAS ARE USED, THESE MULTIVARIATE STATISTICS ARE ONLY APPROXIMATE.

DIFFERENCES ON SINGLE VARIABLES

Table with columns: CHOLSTR, VARIABLE NUMBER, P VALUE, DF, GROUP (1 NOPYLL, 2 PILL), MEAN, STD DEV, S.E.M., SAMPLE SIZE, MAXIMUM, MINIMUM. Includes Levene test results and distribution plots for 1 NOPYLL (N=94) and 2 PILL (N=92).

--- similar analyses for variables 7 to 9 ---

Output 3D.4 Hotelling's T² and Mahalanobis D² - using complete cases only

TEST TITLE. WERNER BLOOD CHEMISTRY DATA
INDEXES OF VARIABLES TO BE ANALYZED 6 7 8 9
USE COMPLETE CASES ONLY YES
PRINT GROUP CORRELATION MATRICES. NO
COMPUTE HOTELLINGS T SQUARE YES
INDEX OF GROUPING VARIABLE. 5

GROUPS USED IN COMPUTATIONS 1 2

DIFFERENCES AMONG GROUP MEANS USING ALL VARIABLES FOR THE FOLLOWING GROUPS

* NOPYLL *
* PILL *

MAHALANOBIS D SQUARE 0.2864
HOTELLING T SQUARE 13.0284
F VALUE 3.2028 P VALUE 0.014
DEGREES OF FREEDOM 4, 177.0

DIFFERENCES ON SINGLE VARIABLES

Table with columns: CHOLSTR, VARIABLE NUMBER, P VALUE, DF, GROUP (1 NOPYLL, 2 PILL), MEAN, STD DEV, S.E.M., SAMPLE SIZE, MAXIMUM, MINIMUM. Includes Levene test results and distribution plots for 1 NOPYLL (N=90) and 2 PILL (N=92).

--- similar analyses for variables 7 to 9 ---

5

⑤ Correlation matrix.

⑥ Squared multiple correlation (SMC) of each variable with all other variables. The condition number is the ratio of the largest eigenvalue to the smallest eigenvalue and is of interest to see how nearly singular the correlation matrix might be. The condition number 0.492002 is read as 49.2.

⑦, ⑧ The eigenvalues of the factors in ⑧ are all listed (under the heading "Variance Explained"). The preassigned criterion for the number of factors

is the number of factors with eigenvalues greater than one (see third line of ①). Therefore, in communalities are obtained for three factors (those with eigenvalues greater than one). The communality of a variable is its squared multiple correlation with the factors extracted.

The cumulative proportion of total variance explained is the sum of the variance explained (eigenvalues) up to and including the factor divided by the sum of all the eigenvalues. A successful factor analysis explains a large proportion of variance with a very few factors.

Output 4M.1 (continued)

CORRELATION MATRIX ⑤

	CONCENTR 1	ANNOY 2	SHOKING1 3	SLEEPY 4	SHOKING2 5	TENSE 6	SHOKING3 7	ALERT 8	IRRITABL 9	TIRED 10	CONTENT 11	SHOKING4 12	
CONCENTR	1	1.000											
ANNOY	2	0.562	1.000										
SHOKING1	3	0.086	0.144	1.000									
SLEEPY	4	0.457	0.360	0.140	1.000								
SHOKING2	5	0.200	0.119	0.785	0.211	1.000							
TENSE	6	0.579	0.705	0.222	0.273	0.301	1.000						
SHOKING3	7	0.041	0.060	0.810	0.126	0.816	0.120	1.000					
ALERT	8	0.802	0.578	0.101	0.606	0.223	0.594	0.039	1.000				
IRRITABL	9	0.595	0.796	0.189	0.337	0.221	0.725	0.108	0.108	1.000			
TIRED	10	0.512	0.413	0.199	0.798	0.274	0.364	0.139	0.605	0.428	1.000		
CONTENT	11	0.492	0.739	0.239	0.240	0.235	0.711	0.100	0.605	0.697	0.394	1.000	
SHOKING4	12	0.228	0.122	0.775	0.277	0.813	0.214	0.845	0.201	0.156	0.271	0.171	1.000

SQUARED MULTIPLE CORRELATIONS (SMC) OF EACH VARIABLE WITH ALL OTHER VARIABLES

⑥

1 CONCENTR	0.70351
2 ANNOY	0.74250
3 SHOKING1	0.73312
4 SLEEPY	0.68377
5 SHOKING2	0.78201
6 TENSE	0.66472
7 SHOKING3	0.82062
8 ALERT	0.80208
9 IRRITABL	0.71437
10 TIRED	0.72627
11 CONTENT	0.69130
12 SHOKING4	0.80294

CONDITION NUMBER = 0.4921950 02

COMMUNALITIES OBTAINED FROM 3 FACTORS AFTER 1 ITERATIONS.

THE COMMUNALITY OF A VARIABLE IS ITS SQUARED MULTIPLE CORRELATION (OR COVARIANCE) WITH THE FACTORS.

⑦

1 CONCENTR	0.6601
2 ANNOY	0.7956
3 SHOKING1	0.8391
4 SLEEPY	0.8474
5 SHOKING2	0.8561
6 TENSE	0.7804
7 SHOKING3	0.8941
8 ALERT	0.9258
9 IRRITABL	0.7978
10 TIRED	0.8453
11 CONTENT	0.7715
12 SHOKING4	0.8698

FACTOR VARIANCE EXPLAINED CUMULATIVE PROPORTION OF TOTAL VARIANCE

⑧

1	5.425688	0.452141
2	2.996636	0.701860
3	1.360520	0.815237
4	0.560300	0.861929
5	0.363261	0.892200
6	0.302254	0.917388
7	0.240804	0.937455
8	0.199752	0.954101
9	0.158162	0.967281
10	0.145653	0.979419
11	0.136736	0.990814
12	0.110235	1.000000

THE VARIANCE EXPLAINED BY EACH FACTOR IS THE EIGENVALUE FOR THAT FACTOR.

TOTAL VARIANCE IS DEFINED AS THE SUM OF THE DIAGONAL ELEMENTS OF THE CORRELATION (COVARIANCE) MATRIX.



⑨ Unrotated factor loadings (pattern) for principal components. These loadings are the eigenvectors of the correlation matrix multiplied by the square roots of the corresponding eigenvalues. They are the correlations of the principal components with the original variables. The eigenvalues (VP) are printed at the bottom of each column.

⑪ Rotated factor loadings (pattern) -- coefficients of the factors after rotation. The sum of squares of the coefficients are printed below each column (VP). When the rotation is orthogonal, as in this example, VP is the variance explained by the factor and the rotated loadings are the correlations of the variables with the factors.

⑩ Orthogonal rotation is performed. Gamma is preassigned to 1 because varimax rotation is performed. At each iteration the simplicity criterion G (p. 488) is printed.

⑫ Plots of the rotated factor loadings. The loadings for one factor are plotted against those of another factor.

UNROTATED FACTOR LOADINGS (PATTERN)

FOR PRINCIPAL COMPONENTS

		FACTOR 1	FACTOR 2	FACTOR 3
CONCENTR	1	0.742	-0.309	0.117
ANNOY	2	0.755	-0.361	-0.309
SHOKING1	3	0.491	0.763	-0.124
SLEEPY	4	0.611	-0.117	0.679
SHOKING2	5	0.561	0.735	-0.030
TENSE	6	0.770	-0.232	-0.366
SHOKING3	7	0.417	0.847	-0.055
ALERT	8	0.808	-0.337	0.244
IRRITABL	9	0.783	-0.302	-0.306
TIRED	10	0.702	-0.138	0.577
CONTENT	11	0.748	-0.256	-0.382
SHOKING4	12	0.540	0.757	0.070
VP		5.426	2.997	1.361

THE VP FOR EACH FACTOR IS THE SUM OF THE SQUARES OF THE ELEMENTS OF THE COLUMN OF THE FACTOR LOADING MATRIX CORRESPONDING TO THAT FACTOR. THE VP IS THE VARIANCE EXPLAINED BY THE FACTOR.

ORTHOGONAL ROTATION, GAMMA = 1.0000

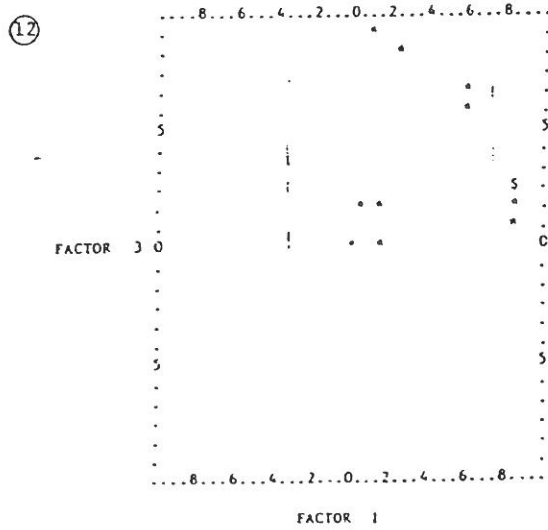
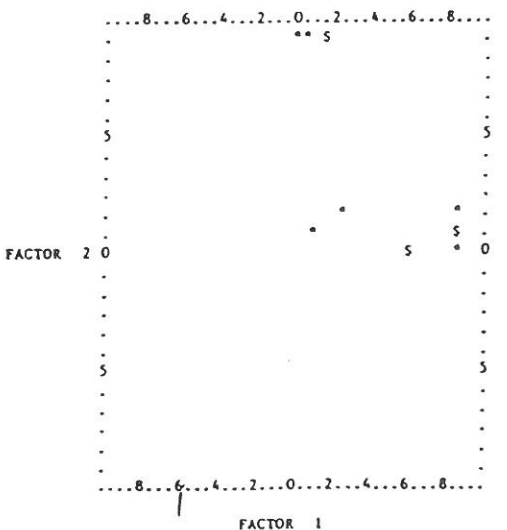
ITERATION	SIMPLICITY CRITERION
0	-1.900373
1	-6.017688
2	-6.019553
3	-6.019557

ROTATED FACTOR LOADINGS (PATTERN)

		FACTOR 1	FACTOR 2	FACTOR 3
CONCENTR	1	0.601	0.034	0.546
ANNOY	2	0.867	0.021	0.209
SHOKING1	3	0.131	0.907	0.007
SLEEPY	4	0.117	0.116	0.906
SHOKING2	5	0.141	0.905	0.128
TENSE	6	0.859	0.147	0.144
SHOKING3	7	0.005	0.945	0.010
ALERT	8	0.590	0.030	0.691
IRRITABL	9	0.863	0.085	0.214
TIRED	10	0.249	0.143	0.873
CONTENT	11	0.862	0.117	0.125
SHOKING4	12	0.061	0.910	0.195
VP		3.802	3.443	2.538

THE VP FOR EACH FACTOR IS THE SUM OF THE SQUARES OF THE ELEMENTS OF THE COLUMN OF THE FACTOR PATTERN MATRIX CORRESPONDING TO THAT FACTOR. WHEN THE ROTATION IS ORTHOGONAL, THE VP IS THE VARIANCE EXPLAINED BY THE FACTOR.

ROTATED FACTOR LOADINGS



OVERLAP IS INDICATED BY A DOLLAR SIGN. SCALE IS FROM -1 TO +1.

--- we omit the plot of factor 3 versus factor 2 ---

(output continued)

~~7~~ (7)

ITERATION FOR MAXIMUM LIKELIHOOD (21)

ITERATION	MAXIMUM CHANGE IN SQRT(UNIQUENESS)	LIKELIHOOD CRITERION TO BE MINIMIZED	STEP BALVINGS
		0.969191	
		0.900708	2
1	0.120209	0.836862	0
2	0.061833	0.834094	0
3	0.012304	0.834089	0
4	0.000609		
5	0.000002		

AN ASTERISK (IF ANY) AFTER THE ITERATION NUMBER INDICATES THAT APPROXIMATE DERIVATIVES WERE USED.

CANONICAL CORRELATIONS

1	0.9790
2	0.9668
3	0.9082

(22)

COMMUNALITIES OBTAINED FROM 3 FACTORS AFTER 5 ITERATIONS.

THE COMMUNALITY OF A VARIABLE IS ITS SQUARED MULTIPLE CORRELATION (OR COVARIANCE) WITH THE FACTORS.

1 CONCENTR	0.5753
2 ANNOY	0.7457
3 SHOKING1	0.7630
4 SLEEPY	0.7596
5 SHOKING2	0.8011
6 TENSE	0.7110
7 SHOKING3	0.8784
8 ALERT	0.7561
9 IRRITABL	0.7551
10 TIRED	0.8043
11 CONTENT	0.6993
12 SHOKING4	0.8352

FACTOR	VARIANCE EXPLAINED	CUMULATIVE PROPORTION OF TOTAL VARIANCE
1	4.793273	0.399439
2	3.162091	0.662947
3	1.128832	0.757016

TOTAL VARIANCE IS DEFINED AS THE SUM OF THE DIAGONAL ELEMENTS OF THE CORRELATION (COVARIANCE) MATRIX.

UNROTATED FACTOR LOADINGS (PATTERN)

FOR MAXIMUM LIKELIHOOD CANONICAL FACTORS

		FACTOR 1	FACTOR 2	FACTOR 3
CONCENTR	1	0.529	0.543	0.027
ANNOY	2	0.527	0.599	-0.329
SHOKING1	3	0.732	-0.466	-0.097
SLEEPY	4	0.518	0.389	0.583
SHOKING2	5	0.790	-0.419	-0.023
TENSE	6	0.579	0.490	-0.369
SHOKING3	7	0.722	-0.596	-0.040
ALERT	8	0.587	0.622	0.157
IRRITABL	9	0.574	0.562	-0.332
TIRED	10	0.590	0.448	0.506
CONTENT	11	0.552	0.507	-0.370
SHOKING4	12	0.789	-0.456	0.066
VP		4.793	3.162	1.129

THE VP FOR EACH FACTOR IS THE SUM OF THE SQUARES OF THE ELEMENTS OF THE COLUMN OF THE FACTOR LOADING MATRIX CORRESPONDING TO THAT FACTOR. THE VP IS THE VARIANCE EXPLAINED BY THE FACTOR.

ORTHOAGONAL ROTATION, GAMMA = 1.0000

ITERATION	SIMPLICITY CRITERION
0	-0.611441
1	-5.852844
2	-5.864646
3	-5.864750
4	-5.864750

ROTATED FACTOR LOADINGS (PATTERN)

		FACTOR 1	FACTOR 2	FACTOR 3
CONCENTR	1	0.595	0.051	0.468
ANNOY	2	0.839	0.030	0.204
SHOKING1	3	0.128	0.864	0.023
SLEEPY	4	0.164	0.116	0.848
SHOKING2	5	0.144	0.874	0.127
TENSE	6	0.818	0.142	0.146
SHOKING3	7	0.007	0.937	0.011
ALERT	8	0.597	0.039	0.631
IRRITABL	9	0.840	0.090	0.204
TIRED	10	0.283	0.137	0.840
CONTENT	11	0.817	0.111	0.142
SHOKING4	12	0.068	0.893	0.183
VP		3.604	3.264	2.216

THE VP FOR EACH FACTOR IS THE SUM OF THE SQUARES OF THE ELEMENTS OF THE COLUMN OF THE FACTOR PATTERN MATRIX CORRESPONDING TO THAT FACTOR. WHEN THE ROTATION IS ORTHOGONAL, THE VP IS THE VARIANCE EXPLAINED BY THE FACTOR.

--- the remainder of the results is analogous to (12) to (18) in Output 4M.1 ---



Output 1R.1 Multiple linear regression. Circled numbers correspond to those in the text

--- the BMDP instructions read by PLR are printed and interpreted ---

```

REGRESSION INTERCEPT. . . . .NON-ZERO
GROUPING VARIABLE . . . . .
WEIGHT VARIABLE . . . . .
PRINT COVARIANCE MATRIX . . . . . NO
PRINT CORRELATION MATRIX. . . . . NO
PRINT CORRELATION OF REGRESSION COEFFICIENTS. . . . . NO
PRINT RESIDUALS . . . . . NO
PRINT NORMAL PROBABILITY PLOT . . . . . NO
PRINT DETRENDED NORMAL PROBABILITY PLOT . . . . . NO

NUMBER OF CASES READ. . . . . 188
CASES WITH DATA MISSING OR BEYOND LIMITS . . . . . 8
REMAINING NUMBER OF CASES . . . . . 180
    
```

VARIABLE	MEAN	STANDARD DEVIATION	COEFFICIENT	MINIMUM	MAXIMUM
			OF VARIATION		
2 AGE	33.53819	9.89836	0.29514	19.00000	55.00000
3 HEIGHT	64.46597	2.48213	0.03850	57.00000	71.00000
4 WEIGHT	131.09384	20.49977	0.15637	94.00000	215.00000
5 BRTHPILL	1.50551	0.50136	0.33302	1.00000	2.00000
6 CHOLSTRL	235.83821	42.74364	0.18124	155.00000	390.00000
7 ALBUMIN	4.12052	0.35871	0.08706	3.20000	5.00000
8 CALCIUM	9.96773	0.47279	0.04743	8.80000	11.10000
9 URICACID	4.75551	1.12111	0.23575	2.20000	9.90000

```

REGRESSION TITLE. . . . .WERNER BLOOD CHEMISTRY DATA
DEPENDENT VARIABLE. . . . . 6 CHOLSTRL
TOLERANCE . . . . . 0.0100
    
```

ALL DATA CONSIDERED AS A SINGLE GROUP

```

MULTIPLE R . . . . .
MULTIPLE R-SQUARE . . . . .
STD. ERROR OF EST. . . . . 39.1698
    
```

ANALYSIS OF VARIANCE

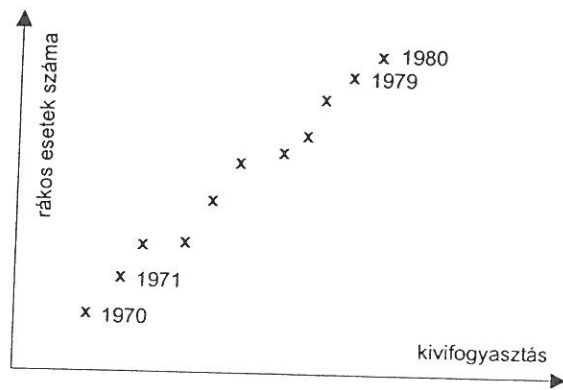
	SUM OF SQUARES	DF	MEAN SQUARE	F RATIO	P(TAIL)
REGRESSION	57004.242	3	19001.414	12.385	0.0
RESIDUAL	270032.000	176	1534.273		

VARIABLE	COEFFICIENT	STD. ERROR	STD. REG COEFF	T	P(2 TAIL)	TOLERANCE
INTERCEPT	151.42036					
AGE	1.38971	0.309	0.322	4.497	0.000	0.915924
WEIGHT	0.00289	0.153	0.001	0.019	0.985	0.869294
URICACID	7.87099	2.769	0.206	2.843	0.005	0.889443

magában. A felismert összefüggés látszólagos lehet, ha az analízis mögött nem állnak elméleti megfontolások.

Ok és okozat

Saville és Wood könyvéből [1] vettük az alábbi példát. A 4.1. ábra az Egyesült Államokban megfigyelt rákos esetek számát mutatja a kivífogasztás függvényében. Mivel 1970 és 1980 között mindkét mennyiség növekedett, ezek évente megfigyelt értékei *korreláltak*. Jóllehet ez matematikai bizonyosság, mégsem állíthatjuk, hogy a rákos esetek számának a növekedését az *okozta*, hogy az emberek több kivit ettek. A ténylegesen talált (és statisztikailag bizonyított) korrelációt csak akkor szabad *ok-okozati* kapcsolatnak tekinteni, ha erre *elméleti indok* van.



4.1. ábra. Kapcsolat az Egyesült Államokban megfigyelt rákos esetek száma és a kivífogasztás között

Hasonló példákat lehet az élet legkülönbözőbb területén találni. Például határozottan pozitív korreláció van a Duna vízállása és a BME területén tartózkodó hallgatók száma között. Nyilván épeszű ember nem tételez fel ezek között ok-okozati kapcsolatot. A matematikai statisztika, vagy inkább az azt rosszul alkalmazó áltudomány iránt bizalmatlan emberek gyakran köszörülik szellemességüket az ilyen korrelációkon. Akkor mire vezethetők vissza ezek a látszólagos összefüggések? A válasz egyszerű. Az ilyen példákban általában lehet találni egy közvetítő mennyiséget, ami legtöbbször az idő. Mikor magas ugyanis a Duna vízszintje? Koratavasszal és késő ősszel. Éppen ezek az időszakok előzik meg a vizsgaidőszakokat, amikor a hallgatók a legszorgalmasabban járnak az egyetemre. Hasonlóan az idő a közvetítő a 4.1. ábrán mutatott példában is.

Az extrapoláció veszélyei

Nem csak a lineáris regresszióban, hanem – általánosabban – a polinom-illesztésben (vö. 4.2. alfejezet) is nagyon veszélyes az illesztésben kapott függvényt a vizsgált valószínűségi változók mérési tartományán túl extrapolálni. Súlyos tévedések forrása az ilyesmi. A probléma hangsúlyozottan főleg a polinomillesztésnél merül fel, ugyanis többnyire akkor fordulunk eh-

ANOVA táblázatok

Egyszempontos varianciaanalízis

A szóródás oka	Négyzetösszeg	Szabadsági fok	Empírikus szórásnégyzet
Csoportok között	$Q_1 = \sum_{i=1}^k n_i (\bar{X}_i - \bar{X}_{..})^2$	$k - 1$	$s_1^2 = \frac{Q_1}{k-1}$
Csoportokon belül	$Q_2 = \sum_{i=1}^k \sum_{j=1}^{n_i} (X_{ij} - \bar{X}_i)^2$	$n - k$	$s_2^2 = \frac{Q_2}{n-k}$
Teljes	$Q = \sum_{i=1}^k \sum_{j=1}^{n_i} (X_{ij} - \bar{X}_{..})^2$	$n - 1$	-

Kétszempontos varianciaanalízis (interakció nélkül)

A szóródás oka	Négyzetösszeg	Szabadsági fok	Empírikus szórásnégyzet
a-hatások	$Q_1 = p \sum_{i=1}^k (\bar{X}_i - \bar{X}_{..})^2$	$k - 1$	$s_1^2 = \frac{Q_1}{k-1}$
b-hatások	$Q_2 = k \sum_{j=1}^p (\bar{X}_j - \bar{X}_{..})^2$	$p - 1$	$s_2^2 = \frac{Q_2}{p-1}$
Véletlen hiba	$Q_3 = \sum_{i=1}^k \sum_{j=1}^p (X_{ij} - \bar{X}_i - \bar{X}_j + \bar{X}_{..})^2$	$(k - 1)(p - 1)$	$s_3^2 = \frac{Q_3}{(k-1)(p-1)}$
Teljes	$Q = \sum_{i=1}^k \sum_{j=1}^p (X_{ij} - \bar{X}_{..})^2$	$kp - 1$	-

Kétszempontos varianciaanalízis (Interakcióval)

A szóródás oka	Négyzetösszeg	Szabadsági fok	Empírikus szórásnégyzet
a-hatások	$Q_1 = pn \sum_{i=1}^k (\bar{X}_{i..} - \bar{X}_{...})^2$	$k - 1$	$s_1^2 = \frac{Q_1}{k-1}$
b-hatások	$Q_2 = kn \sum_{j=1}^p (\bar{X}_{.j.} - \bar{X}_{...})^2$	$p - 1$	$s_2^2 = \frac{Q_2}{p-1}$
ab-interakció	$Q_3 = n \sum_{i=1}^k \sum_{j=1}^p (\bar{X}_{ij.} - \bar{X}_{i..} - \bar{X}_{.j.} + \bar{X}_{...})^2$	$(k - 1)(p - 1)$	$s_3^2 = \frac{Q_3}{(k-1)(p-1)}$
Véletlen hiba	$Q_4 = \sum_{i=1}^k \sum_{j=1}^p \sum_{l=1}^n (X_{ijl} - \bar{X}_{ij.})^2$	$kp(n - 1)$	$s_4^2 = \frac{Q_4}{kp(n-1)}$
Teljes	$Q = \sum_{i=1}^k \sum_{j=1}^p \sum_{l=1}^n (X_{ijl} - \bar{X}_{...})^2$	$kpn - 1$	-

not have an acceptable value. Two values are excluded from the first group in our example.

The midpoint for each interval is printed to the left of the histograms. Each interval includes its upper limit. For example, 210.0 and 225.0 are successive midpoints, so the value 217.5 would be classified into the interval with midpoint 210.0.

- ② For each group, P7D prints
- mean: \bar{x}
 - standard deviation: s based on sample variance
 - standard error of the mean (S.E.M.) based on sample variance
 - robust estimate of standard deviation based on mean deviation from the mean

- maximum and minimum observed value (not out of range)
- sample size (frequency): N

③ For all groups combined, P7D prints the mean, standard deviation, standard error of the mean, maximum, minimum and frequency. The standard deviation is computed from the overall mean for the variable (not from the group means).

④ A one-way analysis of variance (ANOVA) that tests the equality of group means.
Let x_{ij} represent the j th observation in the i th group and \bar{x}_i the mean and N_i the number of

Output 7D.1 Comparison of groups. Circled numbers correspond to those in the text.

--- the BMDP instructions read by P7D are printed and interpreted ---

```

*****
HISTOGRAM OF * CHOLSTRL * (VARIABLE 6). CASES DIVIDED INTO GROUPS BASED ON VALUES OF * AGE * (VARIABLE 2)
*****
25ORLESS      26 TO 35      36 TO 45      OVER 45
VAR 6
EXCLUDED
VALUES
**
TABELATIONS AND COMPUTATIONS WHICH FOLLOW EXCLUDE VALUES LISTED ABOVE
MIDPOINTS
435.000)
420.000)
405.000)
390.000)
375.000)
360.000)
345.000)
330.000)* ①
315.000)
300.000)*
285.000)**
270.000)***
255.000)****
240.000)*****
225.000)M*****
210.000)M*****
195.000)M*****
180.000)M****
165.000)M***
150.000)M**
135.000)M
120.000)M
GROUP MEANS ARE DENOTED BY M'S IF THEY COINCIDE WITH *'S, N'S OTHERWISE
MEAN      222.120      224.633      248.333      262.059
STD. DEV. 37.444      35.742      44.809      43.267
R. E. S. D. 38.044      37.427      46.613      42.448
S. E. M. 3.295      3.614      6.914      7.420
MAXIMUM 330.000      335.000      335.000      390.000
MINIMUM 155.000      160.000      160.000      190.000
SAMPLE SIZE 50      60      42      34
    
```

ALL GROUPS COMBINED
(EXCEPT CASES WITH UNUSED VALUES FOR AGE)

MEAN	236.150
STD. DEV.	42.556
S. E. M.	3.120
MAXIMUM	390.000
MINIMUM	155.000
SAMPLE SIZE	186

ANALYSIS OF VARIANCE TABLE

SOURCE	SUM OF SQUARES	DF	MEAN SQUARE	F VALUE	TAIL PROBABILITY
BETWEEN GROUPS	46857.3398	3	15619.1133	④ 9.86	0.0000
WITHIN GROUPS	288172.4290	182	1583.3650		
TOTAL	335029.7500	185			
LEVENE'S TEST FOR EQUAL VARIANCES		3, 182		⑤ 0.98	0.4054
ONE-WAY ANALYSIS OF VARIANCE					
TEST STATISTICS FOR WITHIN-GROUP					
VARIANCES NOT ASSUMED TO BE EQUAL					
WELCH		3, 90		⑥ 9.11	0.0000
BROWN-FORSYTHE		3, 151		9.42	0.0000

--- the analyses for variables 2 to 5 precede that for CHOLSTRL, those for variables 7 to 9 follow ---

(Handwritten signature)

pressure as a covariate or perform the latter analysis using the logarithm of blood pressure. Note that when there are only two grouping factors, a more detailed analysis can be obtained with P7D.

In the ANOVA table the main effect of each grouping factor is identified by the name of the grouping variable as specified in the VARIABLE paragraph. The first character of the grouping variable names are used to label interactions; therefore the two grouping variables are given names that begin with different letters. In BMDP instructions only the DESIGN paragraph is specific to P2V. The other BMDP instructions are explained in Chapter 5.

(See end of this P2V Section for organization of systems information, BMDP instructions and data)

If the FORM statement is used, the DESIGN paragraph is written

/ DESIGN FORM IS '2G,Y'.

2G specifies that the first two variables are grouping factors, and Y specifies that the third variable is the dependent variable.

The results are presented in Output 2V.1. Circled numbers below correspond to those in the output.

```
-----
/ PROBLEM TITLE IS 'KUTNER SYSTOLIC BLOOD PRESSURE
DATA'.
/ INPUT VARIABLES ARE 3.
      FORMAT IS '(3F3.0)'.
/ VARIABLE NAMES ARE TREATMNT, DISEASE, SYSINCR.
/ DESIGN DEPENDENT IS SYSINCR.
      GROUPING ARE TREATMNT, DISEASE.

/ GROUP CODES(1) ARE 1, 2, 3, 4.
      NAMES(1) ARE DRUG1, DRUG2, DRUG3, DRUG4.
      CODES(2) ARE 1, 2, 3.
      NAMES(2) ARE DISEASE1, DISEASE2, DISEASE3.

/ END
-----
```

① The DESIGN paragraph is interpreted by P2V.

② Number of cases read. Only cases containing acceptable values for all variables specified in the DESIGN paragraph are used in the analysis. An acceptable value is a value that is not missing or out of range. In addition, if CODES are specified for any GROUPING factors (variables), a case is included only if the value of the GROUPING factor is equal to a specified CODE.

③ The frequency (COUNT) of observations in each cell is printed.

Output 2V.1 A two-way analysis of variance by P2V. Circled numbers correspond to those in the text

--- the BMDP instructions read by P2V are printed and interpreted ---

DESIGN SPECIFICATIONS

GROUP = 1 2 ①
DEPEND = 3

VARIABLE NO. NAME	MINIMUM LIMIT	MAXIMUM LIMIT	MISSING CODE	CATEGORY CODE	CATEGORY NAME	GREATER THAN	LESS THAN OR = TO
1 TREATMNT				1.00000	DRUG1		
				2.00000	DRUG2		
				3.00000	DRUG3		
				4.00000	DRUG4		
2 DISEASE				1.00000	DISEASE1		
				2.00000	DISEASE2		
				3.00000	DISEASE3		

GROUP STRUCTURE ③

TREATMNT	DISEASE	COUNT
DRUG1	DISEASE1	6.
DRUG1	DISEASE2	4.
DRUG1	DISEASE3	5.
DRUG2	DISEASE1	5.
DRUG2	DISEASE2	4.
DRUG2	DISEASE3	6.
DRUG3	DISEASE1	3.
DRUG3	DISEASE2	5.
DRUG3	DISEASE3	4.
DRUG4	DISEASE1	5.
DRUG4	DISEASE2	6.
DRUG4	DISEASE3	5.

NUMBER OF CASES READ. 58 ②

(output continued)

13/6

Output 2V.1 (continued)

CELL MEANS FOR 1-ST DEPENDENT VARIABLE (4)

TREATMNT= DISEASE =	DRUG1 DISEASE1	DRUG1 DISEASE2	DRUG1 DISEASE3	DRUG2 DISEASE1	DRUG2 DISEASE2	DRUG2 DISEASE3	DRUG3 DISEASE1	DRUG3 DISEASE2	DRUG3 DISEASE3	DRUG4 DISEASE1
SYSINCR	29.33333	28.25000	20.40000	28.00000	33.50000	18.16667	16.33333	4.40000	8.50000	13.60000
COUNT	6	4	5	5	4	6	3	5	4	5

MARGINAL

TREATMNT= DISEASE =	DRUG4 DISEASE2	DRUG4 DISEASE3	MARGINAL
SYSINCR	12.83333	14.20000	18.87931
COUNT	6	5	58

STANDARD DEVIATIONS FOR 1-ST DEPENDENT VARIABLE

TREATMNT= DISEASE =	DRUG1 DISEASE1	DRUG1 DISEASE2	DRUG1 DISEASE3	DRUG2 DISEASE1	DRUG2 DISEASE2	DRUG2 DISEASE3	DRUG3 DISEASE1	DRUG3 DISEASE2	DRUG3 DISEASE3	DRUG4 DISEASE1
SYSINCR	13.01794	5.85235	13.37161	10.97725	2.08167	12.52863	14.18920	6.91375	9.00000	10.54988

TREATMNT= DISEASE =	DRUG4 DISEASE2	DRUG4 DISEASE3
SYSINCR	10.34247	8.92749

ANALYSIS OF VARIANCE FOR 1-ST DEPENDENT VARIABLE - SYSINCR (5)

SOURCE	SUM OF SQUARES	DEGREES OF FREEDOM	MEAN SQUARE	F	TAIL PROB.
MEAN	20037.61301	1	20037.61301	181.41	0.0000
TREATMNT	2997.47186	3	999.15729	9.05	0.0001
DISEASE	415.87305	2	207.93652	1.88	0.1637
TD	707.26626	6	117.87771	1.07	0.3958
1 ERROR	5080.81667	46	110.45254		

- (4) The mean, frequency and standard deviation of each cell for each dependent variable are printed.
- (5) An ANOVA table is printed.

The sums of squares in the one-way ANOVA are well known. The sums of squares in the two-way, or higher, ANOVA depend upon the hypothesis of interest unless each cell contains the same number of observations. The hypotheses tested by P2V are the same for equal or unequal cell size problems, and are not affected by losing some of the cases. Although the hypotheses tested are independent, the sums of squares for unequal cell size problems are not in general orthogonal. Orthogonal sums of squares methods (or "sequential" methods) test hypotheses that are functions of cell sizes; P2V does not use a sequential method. For more detailed discussions, see Kutner (1974) and Speed and Hocking (1976). (The hypotheses tested by P2V for the main effects are labelled A and B by Kutner and H1 and H2 by Speed and Hocking.) They, as well as others,

recommend these hypotheses for experimental data. Searle (1971, pp. 316-317) points out that sequential methods test hypotheses that depend on the cell sizes and cautions against their use. More general hypotheses can be tested in BMDP4V.

Hypotheses tested. In our example of a two-way ANOVA, let $E(Y_{ij}) = \mu_{ij}$ where Y_{ij} is an observation of the group (i, j). The test of equality of row means is the test that

$$\sum_j \mu_{ij} = \sum_j \mu_{kj} \quad \text{for all } i, k.$$

The test of equality of column means is the test that

$$\sum_i \mu_{ij} = \sum_i \mu_{il} \quad \text{for all } j, l.$$

NB! Az interakció tehát mást jelent a medicinában és mást a biometriában. Amit a biometriában interakciónak nevezünk, azt a medicinában potenciálásnak - esetleg blokkolásnak - hívjuk!

Az egyes Q_K -értékek mutatják az egyes gyógyszerek egyedi hatását, terminus technicus: főhatás. Az egyes Q_K összege és az össz- Q_K közötti különbség mutatja az interakciót. Mind az egyes Q_K -értéket, mind az interakciót (Q_I) a Q_B -hez hasonlítjuk. Ha a Q_I/Q_B -ből megfelelő módon számított F-érték szignifikáns, akkor van interakció, van potenciálás. Ha nem szignifikáns, akkor nincs vagy legalábbis nem jelentős. Ha nincs, akkor az interakcióra jutó négyzetes eltéréseket és szabadságfokot bele szoktuk olvasztani a Q_B -be. Ezt azért tesszük, mert így a „hiba” (Q_B) szabadságfokát növelve megbízhatóbbá tehetjük az analízist.

16.3 táblázat

Kétszemponos varianciaanalízis (a fehérjék minőségének és mennyiségének hatása a patkányok súlygyarapodására)

Nagy fehérjebevitel			Kis fehérjebevitel		
marha	disznó	gabona	marha	disznó	gabona
73	94	98	90	49	107
102	79	74	76	82	95
118	96	56	90	73	97
104	98	111	64	86	80
81	102	95	86	81	98
107	102	88	51	97	74
100	108	82	72	106	74
87	91	77	90	70	67
117	120	86	95	61	89
111	105	92	78	82	58
Σ 1000	995	859	792	787	839

$N = 6 \cdot 10 = 60$ $\Sigma \Sigma x = 5272$ $\bar{x} = 87,87$ $\Sigma \Sigma x^2 = 479\,435,7$

$\frac{(\Sigma \Sigma x)^2}{n} = \frac{27\,793\,984}{60} = 463\,233$ $Q_T = 16\,202,7$

$\frac{1000^2 + 995^2 + 859^2 + 792^2 + 787^2 + 839^2}{10} = 467\,846$
 $-463\,233$

$Q_K = 4\,613$

$\frac{(1000 + 995 + 859)^2 + (792 + 787 + 839)^2}{30} = 466\,401,3$
 $-463\,233$

$Q_M = 3\,168,3$

$\frac{(1000 + 792)^2 + (995 + 787)^2 + (859 + 839)^2}{20} = 463\,499,5$
 $-463\,233$

$Q_P = 266,5$

$Q_I = 4613 - (266,5 + 3168,3) = 1\,178,2$

$Q_B = 16\,200,7 - 4613 = 11\,585,7$

Q_P a három különböző fehérje között talált különbségek összegét jelenti, Q_M pedig ugyanezt a fehérje mennyiségére. N valamennyi volt a szó adat ki

16.4 táblázat

Varianciaanalízis

	Sz. f.	SSQ	MSQ	F
A fehérje eredete	2	266,5	133,2	0,6
Adag	1	3168,3	3168,3	14,8
Interakció	2	1179,2	589,6	2,7
Kezelés	5	4613,0	922,6	4,3
Hiba	54	11585,7	214,6 = s^2	
Összesen	59	16202,7	-	

valamennyi volt a szó adat ki A Q_{F-D} 10+10 ugyane szorzó A (16,4) juk lön' ter ös er

Éppen emiatt Finney-nek [4] egy példáját alakítottam át minimálisan: a sorok sorrendjét cseréltem meg. Ennek következtében nem változott meg sem a Q_T , sem a készítményekre (betű) jutó, sem a nyulakra (oszlopok) jutó, de a sorrendre (sorok) jutó lényegesen csökkent, és emiatt az „error”-ra jutó ugyanannyit nőtt (16.5 táblázat).

16.5 táblázat

Latin négyzet

(insulinkészítmények összehasonlítása, vércukor, mg%)

Napok	Nyulak				Számítások		
	I	II	III	IV	Σx	\bar{x}	$(\Sigma x)^2$
1.	B 47	A 90	C 79	D 50	266	66,50	70756
2.	D 46	B 61	A 87	C 66	260	65,00	67600
3.	A 62	C 74	D 58	B 59	253	63,25	64069
4.	C 76	D 63	B 63	A 69	271	67,75	73441
Σx	231	288	287	244	1050		275806
\bar{x}	57,75	72,00	71,75	61,00	$\bar{x} = 65,625$		

Insulinkészítmények	A	B	C	D
Σx	308	230	295	217
\bar{x}	77,0	57,5	73,8	54,2

$\Sigma \Sigma x^2 = 71452$

Korrektciós faktor = $1050^2/4 = 68906,26$

$Q_T = 2545,75$

A Q_T és a 3 darab Q_K hozzájárulását úgy kell kiszámítani, mint az előzőekben. Az „error”, mint a 12.4-ben volt, mert itt sem voltak paralelek. Itt nem látszik interakció. Ha gyanú lenne rá, akkor paralelekkel kellene megismételni a vizsgálatot. (Megjelent adatokat használók és az átszámítás S. I-re rontaná az áttekinthetőséget.)

Az $F_{13; 61}$ kritikus értéke 4,76.

16.6. táblázat

Az insulinkészítmények ezt erősen meghaladták, a nyulak majdnem elérték, a napok erősen elmaradtak tőle (16.6 táblázat). A napokra nem is számoltuk ki. Ugyanis, ha a beavatkozások – itt a napok – hatástalanok, akkor a csoportok közti eltérés akkora, mint amit az egyedek közti eltérés okoz, mint ezt a 12.4 pontban leírtuk. A véletlen ingadozások miatt azonban hatástalanság esetén sem kapunk „pont 1-et”, hanem ingadozik körülötte. (Viszont ha annyiszor nagyobb, hogy ezt már a véletlen csak ritkán okoz, szignifikánsnak minősítjük.) Természetesen ugyanígy eltérhet a tört értéke lefelé is. Itt a kérdés csak az lehet: „növelte a variabilitást”? Itt nem fordulhat elő az a nagyon de nagyon ritka eset, hogy csök-

Varianciaanalízis

	Sz. f.	SSQ	MSQ	$F_{13; 61}$
Nyulak	3	646,25	215,4	4,44
Napok	3	45,25	15,1	-
Insulin	3	1563,25	521,1	10,74
Hiba	6	291,00	48,5	6,96 = s
Összesen	15	2545,75	-	-

böző helyeken akarunk azonos kezelésekkel vizsgálni különböző lönböző stb. Azokat az eseteket, amikor szándékosan nem egyformán a létszám, később beszéljük meg, de már a kétmintás t-próbát ismertetve is tettünk erről említést.

Egy igen jó hatású, de csak parenteralisan adható készítményt igyekeztek orálisan is hatásossá tenni. Különböző időpontokban mérték a szer vérszintjét. Az egyik ilyen kritikus időpontban mért adatokat mutatja a 16.1 táblázat. (További részleteket

16.1 táblázat

Egyszempontos varianciaanalízis
(Vérszintértékek különböző adagolás mellett)

	im.	sc.	p. os (3×1)	p. os (3×2)	Összesen
Mért értékek	9 10 12 13	9 10 11 13 14 15	6 7 7 8 9 10 12 12 13 13	10 11 13 13 14 14 15 16 16 19	
n	4	6	10	10	30
Σx	44	72	97	139	352
\bar{x}	11,0	12,0	9,7	13,9	11,73
Σx^2	494	892	1005	1989	4380
$n\bar{x}^2$	484	864	940,9	1932,1	4221
Q_B	10	28	64,1	56,9	4130,13
s^2	3,33	5,60	7,21	69,2	$Q_K = 90,87$ 249,87

nem közölhetünk, mert a „védő anyag” toxicusnak bizonyult, és ezért további próbálkozások folynak.) A könnyebb áttekinthetőség céljából utólag nagyság szerint rendeztük az adatokat. A létszám különbözőségének oka: el kellett döntenünk, hogy a 3×1 p. os elegendő-e avagy 3×2-re van szükség, ezért vettek többet. Az injectiós adagolást már eléggé ismerték, egyformának is szánták a két csoport létszámát, de az im. csoportból verekedés miatt kettő elhullott. (Az előző vizsgálatok során nem tapasztaltak ilyen hatást az állatokon, emiatt tartották őket együtt.) Történetesen 32 megfelelő állatuk volt, és ezért tervezték 6–6–10–10 elosztásúnak.

A p. os készítményben egy új anyag is szerepel, amit gondosan vizsgálunk a hatására. Később valóban be is bizonyosodott.

A számítás menete logikája azonos a 12.4-ben ismertetettel, de kissé bonyolultabb. Az össz-négyzetes eltérés (Q_T) és a csoporton belüli (Q_B) kiszámítása változatlan.

16.2 táblázat

Varianciaanalízis

	Sz. f.	SSQ	MSQ
Csoporton belül	26	159,0	6,12
Csoportok között	3	90,9	30,30 = s^2
Összesen	29	249,9	—

$F_{(3,26)} = 4,95$

egy-egy fészkek létszáma. Itt azonban külön-külön kell elvégezni a szorzást, és csak azután szabad összegezni. Könnyebbéség, hogy az $n\bar{x}^2$ -értéket a csoporton belüli kiszámításához is ki kellett számítani. (A kerekítési hibák miatt előnyösebb a $\frac{(\Sigma x)^2}{n}$ értékkel számolni.)

Az SSQ a „négyzetes eltérés összege” (sum of squares) és az MSQ a „négyzetes eltérések átlaga” (mean squares) a nemzetközi irodalomban leggyakrabban használt rövidítések.

A csoportok között (Q_K) azonban mindegyik csoportnál más és más a szorzószám, mert más és más a létszám (16.2 táblázat). A 12.4 példájában a kezelések közöttit számítva, a középértékek négyzeteit összeadtuk, és csak az összegüket szoroztuk 5-tel, a csoportlétszámmal. A fészkek különbségét vizsgálva is előbb összeadhattuk a középértékek négyzeteit, de itt már 4-gyel szoroztuk az összeget, mert ennyi volt az

Tehát a két csoporton belüli jelentősége.

Ebben a példában a két csoport közötti különbség nem szignifikáns, azaz nem különbözik. Így a két csoport közötti különbség nem szignifikáns.

Minden szignifikáns különbséget egyetlen összehasonlításban is meg lehet vizsgálni, ha a két csoport között van, akkor az összehasonlítás a második csoportban történik.

$\binom{n}{2}$; itt tehát

Ezek száma — tehát a próbák száma — az esetén tehát 3 a fészkek között. A $P = 5\%$ arányú esetben kapunk 3 szignifikáns különbséget, vagyis a valószínűsége, hogy a két csoport között van, akkor az összehasonlítás a második csoportban történik. Tehát a két csoport közötti különbség nem szignifikáns.

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- Bartlett's (1947) test for the significance of the k smallest eigenvalues is printed, where k can be 1, 2, etc. The uppermost line (chi-square = 56.31) tests whether the eigenvalues differ significantly from zero; this is a test that the correlations between the two sets of variables are zero. A significant chi-square indicates that the two sets of variables are not independent. The next line (chi-square = 23.66) tests whether all eigenvalues but the largest differ significantly from zero; this is a test of whether the first canonical variable is sufficient to describe the dependence between the two sets of variables. The number of canonical variables of practical value is less than or equal to the smallest number of eigenvalues for which Bartlett's test for the remaining eigenvalues is nonsignificant.

⑨ Canonical variable loadings. These are the correlations of the canonical variables with the original variables. CNVRF1 is the name assigned by PGM to the 1st canonical variable in the first set; CNVRF2 to the 2nd, etc. CNVRS1 is the name assigned to the 1st canonical variable in the second set, etc. These correlations are analogous to unrotated factor loadings.

The canonical paragraph. The variables included in each set of variables must be specified in the CANONICAL paragraph. Each set should contain at least two variables; otherwise a regression program (Chapter 13) should be used.

The number of canonical variables to be obtained can be stated explicitly (NUMBER). If not stated the number is determined by the program as being all canonical variables whose correlations are greater than CONSTANT. (CONSTANT is preset to zero.)

In addition, you can specify the tolerance for matrix inversion and whether covariances and correlations are computed about the mean or about the origin.

CANONICAL

FIRST v list. required none
Names or subscripts of variables in the first set of variables. At least two variables must be specified.

SECOND v list. required none
Names or subscripts of variables in the second set of variables. At least two variables must be specified.

NUMBER = θ . # of vars. in smaller set
Maximum number of canonical variables to be obtained.

CONSTANT = θ . 0.0/prev.
Canonical variables obtained must have a canonical correlation that exceeds CONSTANT.

TOLERANCE = θ . between 0.0001/prev. 0.0 & 1.0
Tolerance for matrix inversion. Inversion is performed by stepwise pivoting. A variable is not pivoted if its squared multiple correlation with already pivoted variables exceeds 1 minus TOLERANCE, or if pivoting causes an already pivoted variable to have a squared multiple correlation with other pivoted variables that exceeds 1 minus TOLERANCE. Note that if a zero intercept model is used, then the R^2 is estimated under the assumption that all variables have zero means.

ZERO.
Covariances and correlations are computed about the origin and not about the mean. This is a rarely used option.

Example 6M.2 Printing the Coefficients of the Canonical Variables and the Canonical Variable Scores

In addition to the correlation matrix and the canonical variable loadings printed in 6M.1, PGM can print the covariance matrix, the canonical variables and the regression coefficients for the canonical variables. The number of cases for which the data

Output 6M.2 Scores and coefficients of the canonical variables

--- ② to ⑧ in Output 6M.1 are printed ---

COEFFICIENTS FOR CANONICAL VARIABLES FOR FIRST SET OF VARIABLES

	CNVRF1	CNVRF2	CNVRF3	CNVRF4
	1	2	3	4
SMOKING1	0.375543D-01	-0.976451D 00	-0.965493D 00	-0.900841D 00
SMOKING2	-0.109322D 01	-0.646536D 00	0.134105D 01	0.182999D 00
SMOKING3	0.119115D 01	-0.173899D 00	-0.333693D-01	0.145194D 01
SMOKING4	-0.704060D 00	0.128569D 01	-0.660196D 00	-0.214955D 00

⑩

STANDARDIZED COEFFICIENTS FOR CANONICAL VARIABLES FOR FIRST SET OF VARIABLES
(THESE ARE THE COEFFICIENTS FOR THE STANDARDIZED VARIABLES - MEAN ZERO, STANDARD DEVIATION ONE.)

	CNVRF1	CNVRF2	CNVRF3	CNVRF4	
	1	2	3	4	
SMOKING1	3	0.043	-1.104	-1.092	-1.019
SMOKING2	5	-1.160	-0.686	1.423	0.194
SMOKING3	7	1.383	-0.202	-0.039	1.686
SMOKING4	12	-0.898	1.641	-0.842	-0.274

Canonical correlations:
0.52229
0.37588
0.24090
0.12719

Number of canonical variables	χ^2	df	Significance
0	56.31	32	0.00502
1	23.66	21	0.20975
2	8.05	12	0.709125
3	1.91	5	0.85637

(output continued)

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Output 6M.2 (continued)

COEFFICIENTS FOR CANONICAL VARIABLES FOR SECOND SET OF VARIABLES

		CNVR51	CNVR52	CNVR53	CNVR54
		1	2	3	4
CONCENTR	1	-0.441692D 00	0.745510D 00	-0.470381D 00	-0.163811D 00
ANNOY	2	0.801410D 00	0.461495D 00	-0.605503D 00	-0.739549D 00
SLEEPY	4	-0.250790D 00	0.581216D 00	-0.685988D 00	0.615867D 00
TENSE	6	-0.692552D 00	-0.380734D 00	0.421877D 00	0.448775D 00
ALERT	8	0.140028D 00	0.204741D 00	0.150159D 01	-0.685341D 00
IRRITABL	9	0.900002D-01	-0.795294D 00	0.425982D 00	0.113746D 01
TIRED	10	-0.327905D 00	-0.616257D 00	-0.246355D 00	0.172116D 00
CONTENT	11	-0.402041D 00	-0.595032D 00	-0.971468D 00	-0.795208D 00

STANDARDIZED COEFFICIENTS FOR CANONICAL VARIABLES FOR SECOND SET OF VARIABLES
(THESE ARE THE COEFFICIENTS FOR THE STANDARDIZED VARIABLES - MEAN ZERO, STANDARD DEVIATION ONE.)

		CNVR51	CNVR52	CNVR53	CNVR54
		1	2	3	4
CONCENTR	1	-0.474	0.800	-0.505	-0.176
ANNOY	2	0.781	0.450	-0.590	-0.721
SLEEPY	4	-0.257	0.595	-0.702	0.630
TENSE	6	-0.687	-0.378	0.418	0.445
ALERT	8	0.143	0.208	1.529	-0.698
IRRITABL	9	0.070	-0.622	0.333	0.890
TIRED	10	-0.313	-0.588	-0.235	0.164
CONTENT	11	-0.339	-0.501	-0.818	-0.670

CANONICAL VARIABLES (CASE NUMBERS REFER TO DATA BEFORE DELETION OF CASES)

CASE										
LABEL	NO.	WEIGHT	CNVRF1	CNVRF2	CNVRF3	CNVRF4	CNVR51	CNVR52	CNVR53	CNVR54
	1	1.0000	-0.1954	1.8242	1.2321	-1.3620	0.1247	1.1330	-0.5973	-0.9229
	2	1.0000	-0.6712	-0.8597	-0.7208	0.7465	-2.5299	-1.3010	0.2394	-0.1880
(11)	3	1.0000	-1.1961	-0.9950	0.9382	0.4104	-2.5714	0.9062	1.3725	0.0831
	4	1.0000	-1.9002	0.2907	0.2780	0.1954	-2.1278	-0.7060	1.2109	0.6072
	5	1.0000	-0.1029	-0.3485	-0.4028	0.2274	-0.8327	0.8508	-0.0588	-1.4278

--- canonical variables for cases 6 to 105 ---

106	1.0000	-0.1029	-0.3485	-0.4028	0.2274	0.5286	-0.8461	1.3204	0.2293
107	1.0000	0.9525	1.2745	-0.7784	0.9452	1.0792	0.1966	0.0289	0.1056
108	1.0000	0.9958	1.6503	1.1997	0.0899	-0.7903	1.9462	0.4103	0.2740
109	1.0000	-0.2387	1.4484	-0.7450	-0.5067	-0.2426	1.8573	-0.3426	-0.1194
110	1.0000	-0.4109	-2.8813	2.6100	-1.1308	1.0021	-1.0009	0.4685	-0.3382

NUMERICAL CONSISTENCY CHECK

THE FOLLOWING VARIANCES OF CANONICAL VARIABLES SHOULD ALL BE EQUAL TO ONE

CANONICAL VARIABLE	VARIANCE	RELATIVE ERROR
CNVRF1	0.100000D 01	0.187350D-14
CNVRF2	0.100000D 01	0.301148D-14
CNVRF3	0.100000D 01	0.213718D-14
CNVRF4	0.100000D 01	0.337230D-14
CNVR51	0.100000D 01	-0.155431D-14
CNVR52	0.100000D 01	0.141553D-14
CNVR53	0.100000D 01	0.366374D-14
CNVR54	0.100000D 01	0.327516D-14

--- (9) in Output 6M.1 is printed ---

(Handwritten marks and circled number 19)

⑦ Summary table. This contains a one line summary of each step including the F-to-enter (or remove) for the variable entered (or removed), the Wilks' lambda U statistic and the approximate F statistic.

⑧ Classification of each case. For each case Mahalanobis D is computed to each group mean. The posterior probability for the distance of a case from a group is the ratio of $\exp(D^2)$ for the group

over the sum of $\exp(D^2)$ for all groups. Prior probabilities, if assigned, affect these computations (see Appendix A.23, step 4). Outliers can be identified as cases with large D^2 from their group means. For large samples from a multivariate normal distribution, the D^2 from a case to its group mean is approximately distributed as a chi-square with degrees of freedom equal to the number of variables selected.

Each case incorrectly classified is noted in the output (cases 5, 9 and 12).

Output 7M.1 (continued)

--- results for steps 2 and 3 ---

STEP NUMBER 4							
VARIABLE ENTERED				1 SEPALLEN			
VARIABLE	F TO REMOVE	FORCE LEVEL	TOLERANCE	VARIABLE	F TO ENTER	FORCE LEVEL	TOLERANCE
1 SEPALLEN	4.721	1	0.347993				
2 SEPALWID	21.936	1	0.608860				
3 PETALLEN	35.590	1	0.365126				
4 PETALWID	24.904	1	0.649314				

U-STATISTIC OR WILKS' LAMBDA 0.0234386 DEGREES OF FREEDOM 4 2 147
 APPROXIMATE F-STATISTIC 199.145 DEGREES OF FREEDOM 8.00 288.00

F - MATRIX DEGREES OF FREEDOM 4 144
 SETOSA VERSICOL
 VERSICOL 550.19
 VIRGINIC 1098.27 105.31

CLASSIFICATION FUNCTIONS

VARIABLE	GROUP - SETOSA	VERSICOL	VIRGINIC
1 SEPALLEN	23.54416	15.69820	12.44584
2 SEPALWID	23.58786	7.07252	3.68529
3 PETALLEN	-16.43063	5.21145	12.76655
4 PETALWID	-17.39839	6.43422	21.07909
CONSTANT	-86.30843	-72.85257	-104.36826

CLASSIFICATION MATRIX

GROUP	PERCENT CORRECT	NUMBER OF CASES CLASSIFIED INTO GROUP -		
		SETOSA	VERSICOL	VIRGINIC
SETOSA	100.0	50	0	0
VERSICOL	96.0	0	48	2
VIRGINIC	98.0	0	1	49
TOTAL	98.0	50	49	51

JACKKNIFED CLASSIFICATION

GROUP	PERCENT CORRECT	NUMBER OF CASES CLASSIFIED INTO GROUP -		
		SETOSA	VERSICOL	VIRGINIC
SETOSA	100.0	50	0	0
VERSICOL	96.0	0	48	2
VIRGINIC	98.0	0	1	49
TOTAL	98.0	50	49	51

SUMMARY TABLE

⑦

STEP NUMBER	VARIABLE ENTERED	F VALUE TO REMOVE	NUMBER OF VARIABLES INCLUDED	U-STATISTIC	APPROXIMATE F-STATISTIC	DEGREES OF FREEDOM
1	3 PETALLEN	1180.1597	1	0.0586	1180.161	2.00 147.00
2	2 SEPALWID	43.0353	2	0.0369	307.104	4.00 292.00
3	4 PETALWID	34.5686	3	0.0250	257.503	6.00 290.00
4	1 SEPALLEN	4.7211	4	0.0234	199.145	8.00 288.00

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⑨ Eigenvalues of the matrix $W^{-1}BW^{-1}$ are computed where B is the between-groups sums of cross products and W is the pooled (within-groups) sum of squares (see Appendix A.23 for a more precise definition). The eigenvalues, canonical correlations between the variables entered and dummy variables representing the groups, and the coefficients for the canonical variables are printed. The first canonical variable is the linear combination of variables entered that best discriminates among the groups (largest one-way

ANOVA F statistic), the second canonical variable is the next best linear combination orthogonal to the first one, etc. The canonical variables are adjusted so that the (pooled) within-group variances are one and their overall mean is zero. The canonical variables are evaluated at the group means.

⑩ The group means only are plotted in a scatter plot. The axes are the first two canonical variables. (This plot is not reproduced in Output 7M.1).

GROUP	INCORRECT CLASSIFICATIONS		MAHALANOBIS D-SQUARE FROM AND POSTERIOR PROBABILITY FOR GROUP -			
	SETOSA		SETOSA	VERSICOL	VIRGINIC	
CASE						
1			0.2 1.000	90.7 0.000	181.6 0.000	
6			1.3 1.000	84.0 0.000	170.1 0.000	⑧
10			2.3 1.000	113.7 0.000	210.0 0.0	
18			2.8 1.000	67.5 0.000	145.7 0.000	
26			4.0 1.000	113.2 0.000	210.2 0.0	

--- similar statistics for the remaining SETOSA cases ---

GROUP	VERSICOL	SETOSA	VERSICOL	VIRGINIC
CASE				
3		105.3 0.000	2.2 0.996	13.1 0.004
8		131.7 0.000	8.4 0.960	14.8 0.040
9	VIRGINIC	130.9 0.000	8.7 0.253	6.5 0.747
11		99.2 0.000	1.3 0.998	13.8 0.002
12	VIRGINIC	149.0 0.000	8.4 0.143	4.9 0.857

--- similar statistics for the remaining VERSICOL cases ---

GROUP	VIRGINIC	SETOSA	VERSICOL	VIRGINIC
CASE				
2		208.6 0.0	27.3 0.000	1.9 1.000
4		207.9 0.0	31.7 0.000	4.5 1.000
5	VERSICOL	133.1 0.000	5.3 0.729	7.2 0.271
7		173.2 0.000	26.6 0.000	11.0 1.000
13		159.0 0.000	12.8 0.003	1.2 0.997

--- similar statistics for the remaining VIRGINIC cases ---

EIGENVALUES			VARIABLE	COEFFICIENTS FOR CANONICAL VARIABLES	
32.19192	0.28539		1 SEPALLEN	0.82938	0.02410
			2 SEPALWID	1.53447	2.16452
			3 PETALLEN	-2.20121	-0.93192
			4 PETALWID	-2.81046	2.83919
			CONSTANT	2.10510	-6.66147
CUMULATIVE PROPORTION OF TOTAL DISPERSION			GROUP CANONICAL VARIABLES EVALUATED AT GROUP MEANS		
0.99121	1.00000		SETOSA	7.60760	0.21514
			VERSICOL	-1.82505	-0.72790
			VIRGINIC	-5.78255	0.51277
CANONICAL CORRELATIONS					
0.98482	0.47120				

--- ⑩ plot of group means ---

(output continued)

CLUSTER MEANS

	DOCNT	PHARM	NURSES	HOSPBEDS	ANIMAL	STARCH	LIFEEXP
1	3.8300	0.8600	3.0000	17.0867	14.6667	71.0000	53.6667
2	3.3329	0.8314	3.4614	23.5600	25.4285	55.0000	54.2857
3	1.2900	0.2300	3.5000	33.9200	21.0000	57.0000	35.0000
GRAND MEAN	3.2827	0.7845	3.3391	22.7363	22.0909	59.5454	52.3636

(5)

CLUSTER STANDARD DEVIATIONS

	DOCNT	PHARM	NURSES	HOSPBEDS	ANIMAL	STARCH	LIFEEXP
1	0.9571	0.3226	1.2826	4.3406	1.2472	1.6330	1.2472
2	2.6076	0.5289	1.5452	9.9481	4.4994	3.7417	2.9137
3	0.0	0.0	0.0	0.0	0.0	0.0	0.0

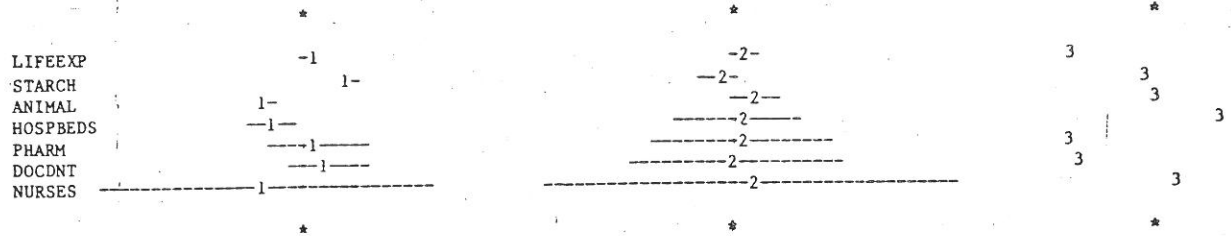
(6)

MEAN SQUARES

BETWEEN	2.4435	0.1700	0.2378	112.7889	122.2632	272.3643	166.2253
WITHIN	6.2934	0.2838	2.7061	93.6598	18.2976	13.2500	8.0119
D.F.-S	2, 8	2, 8	2, 8	2, 8	2, 8	2, 8	2, 8
F-RATIO	0.388	0.599	0.088	1.204	6.682	20.556	20.747
P-VALUE	0.765	0.633	0.965	0.369	0.014	0.000	0.000

CLUSTER PROFILES - VARIABLES ARE ORDERED BY F-RATIO SIZE

(7)



EACH COLUMN DESCRIBES A CLUSTER .
THE CLUSTER NUMBER IS PRINTED AT THE MEAN OF EACH VARIABLE
DASHES INDICATE ONE STANDARD DEVIATION ABOVE AND BELOW

POOLED WITHIN CLUSTER COVARIANCES

	DOCNT	PHARM	NURSES	HOSPBEDS	ANIMAL	STARCH	LIFEEXP
	3	4	5	6	7	8	9
DOCNT	3	6.29					
PHARM	4	1.17	0.28				
NURSES	5	2.90	0.38	2.71			
HOSPBEDS	6	16.27	2.07	12.41	93.66		
ANIMAL	7	9.80	1.89	3.51	22.75	18.30	
STARCH	8	-4.21	-0.73	-0.95	-12.72	-7.63	13.25
LIFEEXP	9	5.59	0.67	3.78	24.04	8.98	-4.12
							8.01

(8)

POOLED WITHIN CLUSTER CORRELATIONS

	DOCNT	PHARM	NURSES	HOSPBEDS	ANIMAL	STARCH	LIFEEXP
	3	4	5	6	7	8	9
DOCNT	3	1.0000					
PHARM	4	0.8757	1.0000				
NURSES	5	0.7039	0.4304	1.0000			
HOSPBEDS	6	0.6702	0.4017	0.7793	1.0000		
ANIMAL	7	0.9134	0.8276	0.4986	0.5495	1.0000	
STARCH	8	-0.4610	-0.3784	-0.1578	-0.3611	-0.4897	1.0000
LIFEEXP	9	0.7868	0.4425	0.8109	0.8777	0.7414	-0.4004
							1.0000

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each cluster for each value of k. The remaining results are printed for the largest value of k.

① For each cluster two histograms display the distance from the cluster center to each case: a) for cases in the cluster, and, b) for cases not in the cluster. The digits in the display indicate the cluster assignment for each case. The scale for each pair of histograms is set to cover the maximum distance from that cluster center.

② The cases in cluster 1 are listed with their weight and distance from the center of cluster 1. When case labels are not used, the case number is printed. The average distance for cases in cluster 1 is also printed.

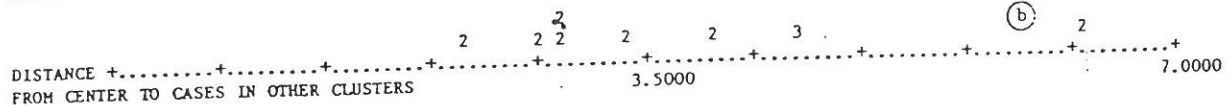
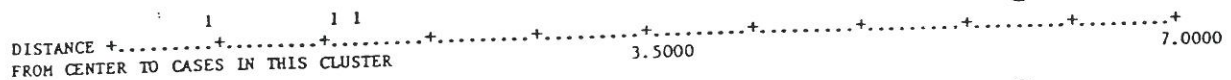
③ The program computes univariate statistics using the standardized data from the three countries in cluster 1: the center (mean), standard deviation and minimum and maximum values.

Output KM.1 K-means cluster analysis of health indicators. Circled numbers correspond to those in the text

--- the BMDP instructions are printed and interpreted ---

CLUSTER 1 OF 3 CONTAINS 3 CASES

STATISTICS ARE COMPUTED FROM THE STANDARDIZED DATA

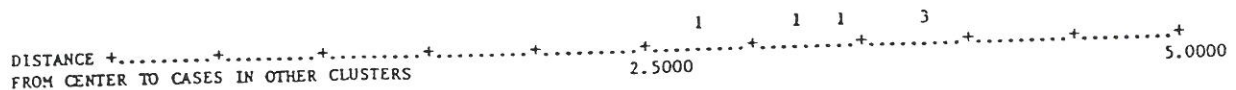
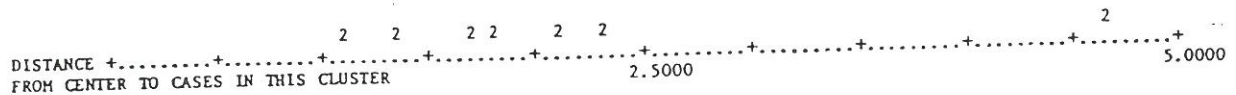


CASE	WEIGHT	DISTANCE	I	VARIABLE	MINIMUM	CENTER	MAXIMUM	ST.DEV.
			I	3 DOCNT	1.1331	1.7085	2.1546	0.5229
SYRIA	1.0000	1.5064	I	4 PHARM	1.1697	1.7648	2.6882	0.8108
TURKEY	1.0000	0.6877	I	5 NURSES	0.9868	2.1145	3.2000	1.1072
U.A.R.	1.0000	1.6138	I	6 HOSPBEDS	1.2349	1.8143	2.3626	0.5645
			I	7 ANIMAL	2.1799	2.4593	2.6829	0.2561
			I	8 STARCH	8.9675	9.2275	9.4874	0.2599
			I	9 LIFEEXP	8.6573	8.9348	9.1567	0.2543

AVERAGE DISTANCE 1.2693

CLUSTER 2 OF 3 CONTAINS 7 CASES

STATISTICS ARE COMPUTED FROM THE STANDARDIZED DATA



CASE	WEIGHT	DISTANCE	I	VARIABLE	MINIMUM	CENTER	MAXIMUM	ST.DEV.
IRAN	1.0000	1.7081	I	3 DOCNT	0.4193	1.4867	4.1620	1.2564
IRAQ	1.0000	1.1302	I	4 PHARM	0.5335	1.7062	3.9400	1.1722
JORDAN	1.0000	1.3758	I	5 NURSES	1.6423	2.4397	4.3136	1.1764
LEBANON	1.0000	4.6609	I	6 HOSPBEDS	1.1818	2.5017	4.3227	1.1410
LIBYA	1.0000	2.3137	I	7 ANIMAL	3.5213	4.2639	5.8689	0.8149
MOROCCO	1.0000	2.1473	I	8 STARCH	6.3683	7.1480	7.7979	0.5252
TUNISIA	1.0000	1.8280	I	9 LIFEEXP	8.4908	9.0378	9.9892	0.5240

AVERAGE DISTANCE 2.1663

(output continued)

CLUSTER 3 OF 3 CONTAINS 1 CASES
ALGERIA

choice of instrumental variable, since it is unlikely to be correlated with measurement errors for x or with the disturbance term in the regression.

Table 7.2.1 Capital-labour substitution data (Kmenta, 1971, p. 313)

Country	y	x	z
United States	0.7680	3.5459	3.4241
Canada	0.4330	3.2367	3.1748
New Zealand	0.4575	3.2865	3.1686
Australia	0.5002	3.3202	3.2989
Denmark	0.3462	3.1585	3.1742
Norway	0.3068	3.1529	3.0492
United Kingdom	0.3787	3.2101	3.1175
Colombia	-0.1188	2.6066	2.5681
Brazil	-0.1379	2.4872	2.5682
Mexico	-0.2001	2.4280	2.6364
Argentina	-0.3845	2.3182	2.5703

The values of z are given in Table 7.2.1. The instrumental variable estimates from (7.2.7) and (7.2.8) (with estimated standard errors) lead to the equation

$$y_i = -2.30 + 0.84 x_i$$

(0.10) (0.03)

It will be noted that IV estimates and OLS estimates are very similar. Thus, in this example, the measurement errors do not seem to be severe.

7.2.2 Two-stage least squares (2SLS) estimation

The instrumental variable matrix $Z(n \times k)$ in Section 7.2.1 is assumed to have the same dimension as the "independent" variable matrix $X(n \times q)$, i.e. $k = q$. However, if $k > q$ then an extension of IV estimation may be given using the method of *two-stage least squares* (2SLS). This method is defined as follows.

First, regress X on Z using the usual OLS' multivariate regression estimates to get a fitted value of X ,

$$\hat{X} = Z(Z'Z)^{-1}Z'X \tag{7.2.9}$$

Note that $\hat{X}(n \times q)$ is a linear combination of the columns of Z .

Second, substitute \hat{X} for X in the original equation (7.2.1) and use OLS

to estimate

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Table 7.4.1 Data for food consumption and prices model

Q_t	P_t	D_t	F_t	A_t
98.485	100.323	87.4	98.0	1
99.187	104.264	97.6	99.1	2
102.163	103.435	96.7	99.1	3
101.504	104.506	98.2	98.1	4
104.240	98.001	99.8	110.8	5
103.243	99.456	100.5	108.2	6
103.993	101.066	103.2	105.6	7
99.900	104.763	107.8	109.8	8
100.350	96.446	96.6	108.7	9
102.820	91.228	88.9	100.6	10
95.435	93.085	75.1	81.0	11
92.424	98.801	76.9	68.6	12
94.535	102.908	84.6	70.9	13
98.757	98.756	90.6	81.4	14
105.797	95.119	103.1	102.3	15
100.225	98.451	105.1	105.0	16
103.522	86.498	96.4	110.5	17
99.929	104.016	104.4	92.5	18
105.223	105.769	110.7	89.3	19
106.232	113.490	127.1	93.0	20

Table 7.4.2 Estimators (with standard errors) for food consumption and prices model

	True coefficient	OLS	2SLS	LIML	3SLS	FIML
<i>Demand equation</i>						
Constant	96.5	99.90	94.63 (7.9)	93.62 (8.0)	Same as 2SLS	Same as LIML
P	-0.25	-0.32	-0.24 (0.10)	-0.23 (0.10)		
D	0.30	0.33	0.31 (0.05)	0.31 (0.05)		
<i>Supply equation</i>						
Constant	62.5	58.28	49.53 (12.01)	Same as 2SLS	52.11 (11.89)	51.94 (12.75)
D	0.15	0.16	0.24 (0.10)		0.23 (0.10)	0.24 (0.11)
F	0.20	0.25	0.26 (0.05)		0.23 (0.04)	0.22 (0.05)
A	0.36	0.25	0.25 (0.10)		0.36 (0.07)	0.37 (0.08)

Values for D_t and the data is summ

The estimated Table 7.4.2. Note variable in each ϵ

Standard errors it can be seen h other procedures. equation are ider Exercises 7.4.1 a described in the n

7.5 System E

7.5.1 Seemingly

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where $X_j(n \times q_j)$ de equation. Note tha hand side of each e the assumptions of applied to each eq parameters. Howe between equations system as a whole.

Write the model

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Example 7.5.1 (Kmenta, 1971, p. 527) Consider data on the investment performance of two firms, General Electric Company and Westinghouse Electric Company, over the period 1935–1954. Each firm's investment (I) is related to the value of its capital stock (C), and the value of its shares (F). The assumed relationship is

$$I_t = \alpha C_t + \beta F_t + \gamma + u_t, \quad t = 1935, \dots, 1954.$$

The results for General Electric are as follows (standard errors in parentheses):

- (a) Using ordinary least squares,

$$I_t = 0.152C_t + 0.027F_t - 9.956.$$

(0.026) (0.016) (31.37)

- (b) Using Zellner's two-stage method,

$$I_t = 0.139C_t + 0.038F_t - 27.72.$$

(0.025) (0.015) (29.32)

The results for Westinghouse were as follows:

- (a) Using ordinary least squares,

$$I_t = 0.092C_t + 0.053F_t - 0.509.$$

(0.056) (0.016) (8.02)

- (b) Using Zellner's two-stage method,

$$I_t = 0.058C_t + 0.064F_t - 1.25.$$

(0.053) (0.015) (7.55)

It can be seen that in the case of each of the six coefficients, Zellner's estimate has a lower estimated standard error than does the ordinary least squares estimate.

7.5.2 Three-stage least squares (3SLS)

The method of three-stage least squares involves an application of Zellner's estimator to the general system of structural equations.

As in Section 7.4.1, write each of the structural equations as a regression-like equation,

$$y_{(j)} = Z_j \delta_{(j)} + u_{(j)}, \quad j = 1, \dots, p - r.$$

Here $Z_j = (Y_{*,j}, X_{0,j})$ denotes those endogenous and exogenous variables (other than $y_{(j)}$) appearing in the j th equation and $\delta_{(j)} = (-\beta'_{*,(j)}, -\gamma'_{0,(j)})'$ represents the corresponding structural coefficients. Also, r denotes the number of exact identities in the system which we omit from consideration.