

**Probability 1**  
**CEU Budapest, fall semester 2016**  
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**Homework sheet 4 – solutions**

4.1 (**homework**) *Poisson approximation of the binomial distribution.* Fix  $0 < \lambda \in \mathbb{R}$ . Show that if  $X_n$  has binomial distribution with parameters  $(n, p)$  such that  $np \rightarrow \lambda$  as  $n \rightarrow \infty$ , then  $X_n$  converges to  $Poi(\lambda)$  weakly.

**Solution:** Set  $q_n = 1 - p_n$ , so  $X_n$  has characteristic function

$$\psi_{X_n}(t) = (q_n + p_n e^{it})^n = \left[ \left( 1 + \frac{e^{it} - 1}{1/p_n} \right)^{1/p_n} \right]^{np_n}.$$

The base of the power converges to  $\exp(e^{it} - 1)$  as  $p_n \rightarrow 0$  by standard elementary calculus, while the exponent converges to  $\lambda$ , so

$$\psi_{X_n}(t) \rightarrow e^{\lambda(e^{it} - 1)},$$

which is exactly the characteristic function of the  $Poi(\lambda)$  distribution, so the continuity theorem ensures that  $X_n$  converges to  $Poi(\lambda)$  weakly.

4.2 (**homework**) Let  $X$  be uniformly distributed on  $[-1; 1]$ , and set  $Y_n = nX$ .

- a.) Calculate the characteristic function  $\psi_n$  of  $Y_n$ .
- b.) Calculate the pointwise limit  $\lim_{n \rightarrow \infty} \psi_n(t)$ , if it exists.
- c.) Does (the distribution of)  $Y_n$  have a weak limit?
- d.) How come?

**Solution:**

a.) The characteristic function of  $X$  is

$$\psi_1(t) = \int_0^1 e^{itx} \frac{1}{2} dx = \frac{1}{2} \left[ \frac{e^{itx}}{it} \right]_0^1 = \frac{\sin t}{t},$$

so

$$\psi_n(t) = \psi_1(nt) = \frac{\sin(nt)}{nt}$$

(with  $\psi_n(0) = 1$ , of course).

b.) So for every  $t \neq 0$  we have  $|\psi_n(t)| \leq \frac{1}{n|t|}$ , which goes to 0 as  $n \rightarrow \infty$ , so

$$\lim_{n \rightarrow \infty} \psi_n(t) = \begin{cases} 0, & \text{if } t \neq 0 \\ 1, & \text{if } t = 0. \end{cases}$$

c.) No:  $\mathbb{P}(Y_n < x) \rightarrow \frac{1}{2}$  as  $n \rightarrow \infty$  for every  $x \in \mathbb{R}$ , and the constant  $\frac{1}{2}$  is not a distribution function. Another possible reasoning is that for any continuous  $\phi : \mathbb{R} \rightarrow \mathbb{R}$  which is bounded by some  $K$  and supported on some bounded interval  $[a, b]$  we have

$$|\mathbb{E}\phi(Y_n)| \leq \mathbb{E}|\phi(Y_n)| \leq K\mathbb{P}(Y_n \in [a, b]) \leq K \frac{b-a}{2n} \xrightarrow{n \rightarrow \infty} 0,$$

so if  $Y_n$  would converge weakly to some  $Y$ , then we would have  $\mathbb{E}\phi(Y) = 0$  for every such  $\phi$ , but then the distribution of  $Y$  has to give zero weight to every interval, which is impossible.

d.) There is no contradiction with the continuity theorem, because the pointwise limit  $\psi(t) := \lim_{n \rightarrow \infty} \psi_n(t)$  of the sequence of characteristic functions is not continuous at 0 (and thus not a characteristic function).

4.3 Durrett [1], Exercise 3.3.1

4.4 Durrett [1], Exercise 3.3.3

4.5 Durrett [1], Exercise 3.3.9

4.6 (**homework**) Durrett [1], Exercise 3.3.10. Show also that independence is needed.

**Solution:**

a.) Denote the characteristic functions of  $X_n$ ,  $Y_n$  and  $X_n + Y_n$  by  $\psi_n$ ,  $\phi_n$  and  $\rho_n$ , respectively. Then the assumptions about independence give  $\rho_n(t) = \psi_n(t)\phi_n(t)$  for every  $t \in \mathbb{R}$  and  $1 \leq n \leq \infty$ , and the continuity theorem gives  $\psi_n(t) \rightarrow \psi_\infty(t)$  and  $\phi_n(t) \rightarrow \phi_\infty(t)$ , so we get  $\rho_n(t) \rightarrow \rho_\infty(t)$ . Using the continuity theorem again gives that  $X_n + Y_n \Rightarrow X_\infty + Y_\infty$ .

b.) To see that independence is needed, consider the following example. For  $1 \leq n < \infty$  let  $X_n \sim B(\frac{1}{2})$  and  $Y_n = 1 - X_n$ , so  $Y_n \sim B(\frac{1}{2})$  also. For  $n = \infty$  let  $X_\infty \sim B(\frac{1}{2})$  again, but set  $Y_\infty = X_\infty$ . Again, this implies  $Y_\infty \sim B(\frac{1}{2})$ . Clearly  $X_n \Rightarrow X_\infty$  and  $Y_n \Rightarrow Y_\infty$ , but  $X_n + Y_n \equiv 1 \not\Rightarrow X_\infty + Y_\infty$ , because e.g.  $\mathbb{P}(X_\infty + Y_\infty = 1) = 0$ .

4.7 Durrett [1], Exercise 3.3.11

4.8 (**homework**) Durrett [1], Exercise 3.3.12

**Solution:** Let  $\xi_1, \xi_2, \dots$  be independent and uniform on the two-element set  $\{-1; 1\}$ , and set  $X_n = \sum_{m=1}^n \frac{\xi_m}{2^m}$ . Then the characteristic function of the  $\xi_m$  is

$$\psi_\xi(t) = \frac{1}{2}e^{it(-1)} + \frac{1}{2}e^{it1} = \cos(t)$$

and the characteristic function of  $X_n$  is

$$\psi_{X_n}(t) = \prod_{m=1}^n \psi_\xi\left(\frac{t}{2^m}\right) = \prod_{m=1}^n \cos\left(\frac{t}{2^m}\right).$$

But notice that  $X_n$  is uniform on the  $2^n$ -element set

$$\left\{ \frac{k}{2^n} : k = -2^n + 1; -2^n + 3; -2^n + 5; \dots; 2^n - 3; 2^n - 1 \right\},$$

so  $X_n$  converges weakly to some  $X$  with the (continuous) uniform distribution on  $[-1; 1]$ . (This can easily be seen e.g. from the pointwise convergence of the distribution functions.) So the characteristic function of  $X$  is

$$\psi_X(t) = \int_{-1}^1 e^{itx} \frac{1}{2} dx = \frac{\sin t}{t},$$

so the continuity theorem states that

$$\frac{\sin t}{t} = \lim_{n \rightarrow \infty} \psi_{X_n}(t) = \prod_{m=1}^{\infty} \cos\left(\frac{t}{2^m}\right).$$

4.9 Durrett [1], Exercise 3.3.13

4.10 (**homework**) Let  $X_1, X_2, \dots$  be i.i.d. random variables with density (w.r.t. Lebesgue measure)  $f(x) = \frac{1}{\pi} \frac{1}{1+x^2}$ . (So they have the Cauchy distribution.) Find the weak limit (as  $n \rightarrow \infty$ ) of the average

$$\frac{X_1 + \dots + X_n}{n}.$$

*Warning: this is not hard, but also not as trivial as it may seem. Hint: a possible solution is using characteristic functions. Calculating the characteristic function of the Cauchy distribution is a little tricky, but you can look it up.*

**Solution:** The characteristic function of the Cauchy distribution is  $\psi_{X_k}(t) = e^{-|t|}$  (see e.g. Durrett [1], Example 3.3.9). So  $S_n = X_1 + \dots + X_n$  has characteristic function  $\psi_{S_n}(t) = (\psi_{X_k}(t))^n = e^{-n|t|}$  and  $\frac{S_n}{n}$  has characteristic function  $\psi_{\frac{S_n}{n}}(t) = \psi_{S_n}\left(\frac{t}{n}\right) = e^{-|t|}$ . This means that  $\frac{S_n}{n}$  has the same Cauchy distribution as the  $X_k$  for every  $n$ , so it also converges to the Cauchy distribution weakly.

Note that this does not contradict the weak law of large numbers, because our  $X_k$  do not have an expectation.

4.11 Durrett [1], Exercise 3.3.20

4.12 Durrett [1], Exercise 3.4.4

4.13 Durrett [1], Exercise 3.4.5

4.14 Durrett [1], Exercise 3.6.1

4.15 Durrett [1], Exercise 3.6.2

## References

[1] Durrett, R. *Probability: Theory and Examples*. Cambridge University Press (2010)