

[4.1] Burlett Ex. 3.2.11

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Let  $X_n$ ,  $1 \leq n \leq \infty$  be integer valued. Show that  $X_n \Rightarrow X_0$  iff  $P(X_n = m) \rightarrow P(X_0 = m)$  for all  $m$ .

Proof (one of the many possible proofs). Let  $F_n$  be the distr. fn of  $X_n$ .

a.) Assume  $X_n \Rightarrow X_0$  and fix  $m \in \mathbb{Z}$ . So  $m - \frac{1}{2}$  and  $m + \frac{1}{2}$  are continuity points of  $F_0$ , which means that

$$P(X_n = m) = F_n(m + \frac{1}{2}) - F_n(m - \frac{1}{2}) \rightarrow F_0(m + \frac{1}{2}) - F_0(m - \frac{1}{2}) = P(X_0 = m).$$

b.) Assume  $P(X_n = m) \rightarrow P(X_0 = m)$  for every  $m$ .

For any  $\varepsilon > 0$  we can choose  $N$  so big that

$$P(X_0 \notin [N, N]) = \sum_{m=-N}^N P(X_0 = m) > 1 - \varepsilon.$$

~~So for any  $\varphi: \mathbb{R} \rightarrow \mathbb{R}$  bounded and continuous,~~

This means that for  $n$  big enough,

~~$$\lim_{n \rightarrow \infty} P(X_n \notin [N, N]) = \lim_{n \rightarrow \infty} \left[ 1 - \sum_{m=-N}^N P(X_n = m) \right] = 1 - \varepsilon$$~~

$$\lim_{n \rightarrow \infty} P(X_n \notin [N, N]) = \lim_{n \rightarrow \infty} \left[ 1 - \sum_{m=-N}^N P(X_n = m) \right] \xrightarrow{\text{finite sum}} 1 - \varepsilon$$

$$= 1 - \sum_{m=-N}^N \lim_{n \rightarrow \infty} P(X_n = m) = 1 - \sum_{m=-N}^N P(X_0 = m) \leq 1 - (1 - \varepsilon) = \varepsilon.$$

So for  $\varphi: \mathbb{R} \rightarrow \mathbb{R}$  bounded and continuous, with  $|\varphi| \leq M$ ,

$$\begin{aligned} \text{we have } & |E(\varphi(X_n) - \varphi(X_0))| \leq M \underbrace{P(X_n \notin [N, N])}_{\leq \varepsilon} + M \underbrace{P(X_0 \notin [N, N])}_{\leq \varepsilon} \\ & + \sum_{m=-N}^N \varphi(m) \underbrace{[P(X_n = m) - P(X_0 = m)]}_{\rightarrow 0 \text{ as } n \rightarrow \infty} \leq 3M\varepsilon, \text{ if } n \text{ is big enough.} \end{aligned}$$

□

**4.2** Let  $X_n \sim \text{Bin}(n, p)$  so  $P(X_n=m) = \begin{cases} \binom{n}{m} p^m (1-p)^{n-m} & \text{for } m \in \{0, 1, \dots, n\} \\ 0 & \text{if not.} \end{cases}$

$X_n \sim \text{Poi}(\lambda)$  means  $P(X_n=m) = e^{-\lambda} \frac{\lambda^m}{m!}$  for  $m = 0, 1, 2, \dots$

So for every  $m \leq -1$ ,  $P(X_n=m)=0 \rightarrow 0 = P(X_n=m)$ .

For  $m \geq 0$ ,  $\boxed{P(X_n=m)} = \frac{n(n-1)\dots(n-m+1)}{m!} p^m (1-p)^{n-m} =$

$$= \frac{1}{m!} (1-p)^{-m} \left(1 - \frac{1}{p}\right)^{\frac{1}{p}} \left(\frac{(np)}{n}\right)^m \frac{n}{n} \frac{n-1}{n} \dots \frac{n-m+1}{n} =$$

$$\xrightarrow[n \rightarrow \infty]{p \rightarrow 0} \frac{1}{m!} e^{-\lambda} \lambda^m = \boxed{P(X_n=m)}.$$

The previous exercise gives  $X_n \Rightarrow X_0$   $\square$

**4.4** See Durrett, Thm 1.2.2 for a precise proof.

**4.8** 0.1 If  $X_n \rightarrow X$  in probability then  $X_n \Rightarrow X$

~~Proof: Let  $x$  be a continuity point of  $F$ , the dist. fn of  $X$ . Then for any  $\delta > 0$ ,  $F_n(x+\delta) \leq P(X_n \geq x+\delta) \leq P(X \geq x+\delta)$~~

4.9 Darrett 3.2.12.

Let  $F_n$  be the distr. fn of  $X_n$  and  $F$  the distr. fn of  $X$ .

a.) Assume that  $X_n \rightarrow X$  in probability and  $x$  is a continuity point of  $F$ .

Then for every  $\varepsilon > 0$   $\exists \delta > 0$  s.t.  $|F(x+\delta) - F(x)| < \frac{\varepsilon}{2}$   
and also  $|F(x-\delta) - F(x)| < \frac{\varepsilon}{2}$ .

This means that  ~~$X_n = x$~~ ,

$$\begin{aligned} F_n(x) = P(X_n \leq x) &\leq P(X \leq x + \delta/2 \text{ or } |X_n - x| \geq \delta/2) \\ &\leq P(X \geq x + \delta/2) + P(|X_n - x| \geq \delta/2) \\ &\quad \left[ \text{since } X_n \geq x \text{ implies that } X \leq x + \delta/2 \text{ or } |X_n - x| \geq \delta/2 \right] \end{aligned}$$

$$F_n(x) = P(X_n \leq x) \leq P(X \leq x + \delta) \text{ or } |X_n - x| \geq \delta \leq$$

$\left[ \text{since } X_n \geq x \text{ implies } X \leq x + \delta \text{ or } |X_n - x| \geq \delta \right]$

$$\leq P(X \leq x + \delta) + P(|X_n - x| \geq \delta) = F(x + \delta) + P(|X_n - x| \geq \delta).$$

Now if  $n$  is bigger than some  $n_0 = n_0(\delta, \varepsilon)$ , then

$P(|X_n - x| \geq \delta) < \frac{\varepsilon}{2}$  by our assumption that  $X_n \xrightarrow{\text{prob.}} X$ ,

so for  $n > n_0$  we have  $F_n(x) \leq F(x + \delta) + \frac{\varepsilon}{2} \leq F(x) + \frac{\varepsilon}{2} + \frac{\varepsilon}{2} = F(x) + \varepsilon$ .

Similarly,  $F_n(x) = P(X_n \leq x) \geq P(X \leq x - \delta \text{ and } |X_n - x| < \delta) \geq$

$\left[ \text{since if } X \leq x - \delta \text{ and } |X_n - x| < \delta, \text{ then } X_n \leq x \right]$

$\geq P(X \leq x - \delta) - P(|X_n - x| \geq \delta) \geq F(x) - \frac{\varepsilon}{2} - \frac{\varepsilon}{2} = F(x) - \varepsilon \text{ if } n \text{ is big enough,}$

so  $\forall \varepsilon > 0 \exists n_0: n > n_0 \text{ implies } |F_n(x) - F(x)| < \varepsilon$ .  ~~$\forall \varepsilon > 0 \exists n_0: n > n_0 \text{ implies } |F_n(x) - F(x)| < \varepsilon$~~

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We have just shown that  $\bar{F}_n(x) \rightarrow F(x)$  for every continuity point  $x$  of  $F$ , so  $X_n \Rightarrow X$ . ~~□~~ □

b.) Assume  $X_n \Rightarrow X$  where  $x=c \in \mathbb{R}$ . Then

$F(x) = \begin{cases} 0, & \text{if } x < c \\ 1, & \text{if } x \geq c \end{cases}$ , so any  $x \neq c$  is a continuity point of  $F$ .

So for any  $\delta > 0$

$$\begin{aligned} P(|X_n - X| \geq \delta) &= P(X_n \leq c-\delta \text{ or } X_n \geq c+\delta) \leq \\ &\leq P(X_n \leq c-\delta) + P(X_n \geq c+\delta) \leq P(X_n \leq c-\delta) + P(X_n > c + \frac{\delta}{2}) = \\ &= \bar{F}_n(c-\delta) + 1 - \bar{F}_n(c + \frac{\delta}{2}) \xrightarrow{n \rightarrow \infty} F(c-\delta) + 1 - F(c + \frac{\delta}{2}) = \\ &= 0 + 1 - 1 = 0. \end{aligned}$$

We have shown that  $X_n \rightarrow c$  in probability. □