

Problem Set 4 : SDE-s

Solutions

4.1 Straightforward applications of Itô's formula.

4.2
$$\begin{pmatrix} U(t) \\ V(t) \end{pmatrix} = \begin{pmatrix} \cos B(t) \\ \sin B(t) \end{pmatrix}$$

by applying Itô's formula one gets

$$d \begin{pmatrix} U(t) \\ V(t) \end{pmatrix} = -\frac{1}{2} \begin{pmatrix} U(t) \\ V(t) \end{pmatrix} dt + \begin{pmatrix} -V(t) \\ U(t) \end{pmatrix} dB(t).$$

4.3

$$\textcircled{a} \quad dB(t) = e^t dX(t) + e^t X(t) dt$$

$$\stackrel{Ito}{=} d(e^t X(t))$$

$$X(t) = e^{-t} X(0) + e^{-t} B(t)$$

$$\textcircled{b} \quad \text{Let } \begin{pmatrix} X(t) \\ Y(t) \end{pmatrix} := \begin{pmatrix} r \cdot t + \alpha \int_0^t X(s) dB(s) \\ B(t) \end{pmatrix}$$

and apply Ito to $g(t; x, y) = x e^{-\alpha y + \frac{\alpha^2}{2} t}$

$$Z(t) = g(t; X(t), B(t))$$

$$dZ(t) = \frac{\alpha^2}{2} X(t) e^{-\alpha B(t) + \frac{\alpha^2}{2} t} dt$$

$$+ e^{-\alpha B(t) + \frac{\alpha^2}{2} t} dX(t)$$

$$- \alpha X(t) e^{-\alpha B(t) + \frac{\alpha^2}{2} t} dB(t)$$

$$+ \frac{\alpha^2}{2} X(t) e^{-\alpha B(t) + \frac{\alpha^2}{2} t} dt$$

$$- 2 \frac{\alpha^2}{2} X(t) e^{-\alpha B(t) + \frac{\alpha^2}{2} t} dt =$$

$$e^{-\alpha B(t) + \frac{\alpha^2}{2} t} (dX(t) - \alpha X(t) dB(t)) =$$

$$r e^{-\alpha B(t) + \frac{\alpha^2}{2} t} dt$$

We arrived at

$$d \left(X(t) e^{-\alpha B(t) + \frac{\alpha^2}{2} t} \right) = r e^{-\alpha B(t) + \frac{\alpha^2}{2} t} dt$$

$$X(t) = X(0) e^{\alpha B(t) - \frac{\alpha^2}{2} t}$$

$$+ r \int_0^t e^{\alpha (B(t) - B(s)) - \frac{\alpha^2}{2} (t-s)} ds$$

$$\textcircled{c} \quad d(e^{-\mathcal{F}t} X(t)) = e^{-\mathcal{F}t} A dB(t)$$

$$X(t) = e^{\mathcal{F}t} X(0) + \int_0^t e^{\mathcal{F}(t-s)} A dB(s)$$

Note that: $e^{\mathcal{F}t} = \begin{pmatrix} \cos t & + \sin t \\ -\sin t & \cos t \end{pmatrix}$

$$\begin{pmatrix} X_1(t) \\ X_2(t) \end{pmatrix} = \begin{pmatrix} \cos t X_1(0) + \sin t X_2(0) \\ -\sin t X_1(0) + \cos t X_2(0) \end{pmatrix}$$

$$+ \int_0^t \begin{pmatrix} \cos(t-s) & + \sin(t-s) \\ -\sin(t-s) & \cos(t-s) \end{pmatrix} \begin{pmatrix} \alpha & 0 \\ 0 & \beta \end{pmatrix} \begin{pmatrix} dB_1(s) \\ dB_2(s) \end{pmatrix}$$

$$\begin{pmatrix} \alpha \cos(t-s) + \beta \sin(t-s) \\ -\alpha \sin(t-s) + \beta \cos(t-s) \end{pmatrix}$$

(4.4)

$$(a) \quad X(t) = X(0) e^{-\gamma t} + a \int_0^t e^{-\gamma(t-s)} dB(s)$$

$$(b) \quad E(X|t) = x_0 e^{-\gamma t}$$

let $0 \leq s \leq t < \infty$:

$$\text{Cov}(X(t), X(s)) =$$

$$a^2 E \left(\int_0^t e^{-\gamma(t-u)} dB(u) \cdot \int_0^s e^{-\gamma(s-v)} dB(v) \right) =$$

$$a^2 \int_0^s e^{-\gamma(t+s-2u)} du =$$

$$\frac{a^2}{2\gamma} \left(e^{-\gamma|t-s|} - e^{-\gamma(t+s)} \right)$$

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$$E(Y_{k+1}^{(n)} - Y_k^{(n)} | Y_k^{(n)} = y) = -\frac{2}{n} \left(y - \frac{n}{2} \right)$$

$$\text{Var}(Y_{k+1}^{(n)} - Y_k^{(n)} | Y_k^{(n)} = y) = 1 - \frac{4}{n^2} \left(y - \frac{n}{2} \right)^2$$

May think about it as:

$$Y_{k+1}^{(n)} - Y_k^{(n)} = -\frac{2}{n} \left(Y_k^{(n)} - \frac{n}{2} \right) + \underbrace{\epsilon_{k+1}^{(n)}}_{\text{Small error}} +$$

Small error

random variable independent of the past

$$E(\epsilon_{k+1}^{(n)}) = 0$$

$$\text{Var}(\epsilon_{k+1}^{(n)}) = 1$$

Now, let $X^{(n)}(t) := \frac{1}{\sqrt{n}} \left(Y_{[nt]}^{(n)} - \frac{t}{2} \right)$ (7)

$$dt \approx \frac{1}{N}$$

to get

$$dX^{(n)}(t) \approx -2X^{(n)}(t) dt + dB(t) \dots$$

(4.5) $t \mapsto X(t)$ $t \in [0, 1]$
Gaussian, continuous sample paths

$$\mathbb{E}(X(t)) = 0$$

$$\mathbb{E}(X(t)X(s)) = (t \wedge s) (1 - (t \vee s))$$

(a) $t \mapsto X(t)$ clearly Gaussian, continuous

$$\mathbb{E}(X(t)) = \mathbb{E}(B(t)) - t\mathbb{E}(B(1)) = 0$$

(P)

$$0 \leq \lambda \leq t \leq 1$$

$$E(X(t)X(\lambda)) = E(B(t)B(\lambda))$$

$$= t E(B(1)B(\lambda)) - \lambda E(B(t)B(1))$$

$$+ \lambda t E(B(1)B(1)) = \dots = \lambda(1-t)$$

(b) similarly:

$$E(Y(t)) = (1-t) E(B(\frac{t}{1-t})) = 0$$

$$0 \leq \lambda \leq t \leq 1$$

$$E(Y(\lambda)Y(t)) =$$

$$(1-t)(1-\lambda) E(B(\frac{t}{1-t})B(\frac{\lambda}{1-\lambda})) =$$

$$\dots (1-t) \cdot \lambda.$$

(c) similarly:

$$E(Z(t)) = 0 \checkmark$$

$$E(Z(s)Z(t)) = \dots$$

$$= \int_0^s \frac{(1-t)}{(1-u)} \cdot \frac{(1-s)}{(1-u)} du =$$

$$= (1-t)(1-s) \int_0^s \frac{du}{(1-u)^2} = \dots = \Delta(t)$$

(d)

differentiate \rightarrow

$$\frac{Z(t)}{1-t} = \int_0^t \frac{dB(s)}{1-s}$$

Continuity at $t \rightarrow 0$:

Argument on blackboard. ...

(10)

Continuity of $t \mapsto Y(t)$ (and $t \mapsto Z(t)$)
at $t=1$:

$t \mapsto M(t) := (1-t)^{-1} Y(t)$ is a
continuous martingale for $t \in [0, 1)$

~~Doob's maximal inequality:~~

for any $0 \leq t \leq u < 1$

$$P\left(\sup_{t \leq s \leq u} |Y(s)| \geq \lambda\right) \leq$$

$$P\left(\sup_{t \leq s \leq u} |M(s)| \geq (1-u)^{-1} \lambda\right) \leq$$

by Doob's
maximal
ineq

$$\frac{(1-u)^2}{\lambda^2} E(M(u)^2) = \frac{u(1-u)}{\lambda^2} \quad (*)$$

Let $t_n := 1 - 2^{-n}$ $n=0, 1, 2, \dots$

$$A_n := \left\{ \sup_{t_{n+1} \leq s \leq t_n} |Y(s)| > 2^{-\frac{n}{4}} \right\}$$

(1)

Then by (*)

$$P(A_n) \leq 2^{-\frac{n}{2}} \quad \text{summable}$$

Thus, by Borel-Cantelli, almost surely
 $\exists n_0 < \infty$ (random, but a.s. finite!)

such that

($\forall n > n_0$):

$$\sup_{1-2^{-(n+1)} \leq \Delta \leq 1-2^{-n}} |Y(\Delta)| \leq 2^{-\frac{n}{4}}$$

This implies: $\lim_{t \rightarrow 1} Y(t) = 0$, a.s.

□

4.6

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$$t \mapsto \Theta(t) = (\Theta_1(t), \dots, \Theta_n(t))$$

deterministic, piece-wise continuous,
bdd.

$$M_\Theta(t) := \exp \left\{ \int_0^t \Theta(s) \cdot dY(s) - \frac{1}{2} \int_0^t |\Theta(s)|^2 ds \right\}$$

is a martingale.

Since $s \mapsto \Theta(s)$ is deterministic

we get:

$$\mathbb{E} \left(e^{\int_0^t \Theta(s) dY(s)} \right) = e^{\frac{1}{2} \int_0^t |\Theta(s)|^2 ds}$$

let $\Theta_i(s) = \chi_{[a_i, b_i]}(s) \cdot \lambda_i$
e.g.

$$E \left(e^{\sum_{i=1}^n \lambda_i (Y_i(b_i) - Y_i(a_i))} \right) = \quad (13)$$
$$e^{\frac{1}{2} \sum_{i=1}^n \lambda_i^2 (b_i - a_i)}$$

these are exactly the exponential moments of Gaussians ...

4.7 ...