

**Midterm Exam - April 20, 2023, Stochastic Analysis, Solutions**

1. (7 points) Let  $X_n$  denote the position of a one-dimensional simple symmetric random walker at time  $n$ . Let  $(\mathcal{F}_n)_{n \geq 0}$  denote the natural filtration of  $(X_n)$ .

(a) Given  $a_k \in \mathbb{R}_+, k \in \mathbb{N}$ , how to choose  $b_k \in \mathbb{R}_+, k \in \mathbb{N}$  if we want  $(M_n)$  to be a martingale, where

$$M_n = \prod_{k=1}^n \frac{(a_k)^{X_k - X_{k-1}}}{b_k}, \quad n = 0, 1, 2, \dots \quad (1)$$

(b) Write  $M_n - 1$  as the discrete stochastic integral  $(H \cdot X)_n$  of a predictable process  $(H_n)$  with respect to the martingale  $(X_n)$ . Write the formula for  $H_n$  in a way that shows that  $(H_n)$  is indeed predictable.

**Solution:**

(a)  $(M_n)$  is clearly an adapted process.  $M_0 = 1$  and  $M_n = M_{n-1} \frac{(a_n)^{X_n - X_{n-1}}}{b_n}$ , thus

$$\mathbb{E}[M_n | \mathcal{F}_{n-1}] \stackrel{(\diamond)}{=} M_{n-1} \mathbb{E} \left[ \frac{(a_n)^{X_n - X_{n-1}}}{b_n} \mid \mathcal{F}_{n-1} \right] \stackrel{(\blacklozenge)}{=} M_{n-1} \mathbb{E} \left[ \frac{(a_n)^{X_n - X_{n-1}}}{b_n} \right] = M_{n-1} \frac{\frac{1}{2} \frac{1}{a_n} + \frac{1}{2} a_n}{b_n}$$

where in  $(\diamond)$  we used that  $M_{n-1} \in \mathcal{F}_{n-1}$  and in  $(\blacklozenge)$  we used that the random walk has independent increments. Thus  $b_k = \frac{1}{2} \frac{1}{a_k} + \frac{1}{2} a_k$ .

(b) We want to identify the random variable  $H_n$  for which  $M_n - M_{n-1} = H_n \cdot (X_n - X_{n-1})$ . Now

$$\begin{aligned} M_n - M_{n-1} &= M_{n-1} \left( \frac{(a_n)^{X_n - X_{n-1}}}{b_n} - 1 \right) = \frac{M_{n-1}}{b_n} \left( (a_n)^{X_n - X_{n-1}} - \frac{1}{2} \frac{1}{a_n} - \frac{1}{2} a_n \right) \stackrel{(**)}{=} \\ &M_{n-1} \frac{a_n - \frac{1}{a_n}}{2b_n} (X_n - X_{n-1}), \end{aligned}$$

where  $(**)$  holds if  $X_n - X_{n-1} = 1$  and also if  $X_n - X_{n-1} = -1$  (these are the two possible values of  $X_n - X_{n-1}$ ). Thus  $H_n = M_{n-1} \frac{a_n - \frac{1}{a_n}}{2b_n}$ , which is indeed  $\mathcal{F}_{n-1}$ -measurable since  $M_{n-1} \in \mathcal{F}_{n-1}$  and  $\frac{a_n - \frac{1}{a_n}}{2b_n}$  is deterministic.

2. (8 points) Let  $0 = t_0 < t_1 < \dots < t_{n-1} < t_n = t$ . Let  $L_n = \sum_{k=1}^n e^{B(t_{k-1})} \cdot (B(t_k) - B(t_{k-1}))$  and  $\mathcal{I} = \int_0^t e^{B_s} dB_s$ . Show that

$$\mathbb{E}((\mathcal{I} - L_n)^2) = \sum_{k=1}^n e^{2t_{k-1}} \cdot \left[ \frac{1}{2} \left( e^{2(t_k - t_{k-1})} - 1 \right) - 4 \left( e^{\frac{1}{2}(t_k - t_{k-1})} - 1 \right) + (t_k - t_{k-1}) \right]. \quad (2)$$

*Help:* If  $X \sim \mathcal{N}(0, \sigma^2)$  then the moment generating function of  $X$  is  $M(\lambda) = e^{\frac{1}{2} \lambda^2 \sigma^2}$ .

**Solution:** Let  $X_t := \sum_{k=1}^n e^{B(t_{k-1})} \mathbb{1}[t_{k-1} < t \leq t_k]$ , thus  $(X_t)$  is a simple predictable process. We have

$$\mathbb{E}((\mathcal{I} - L_n)^2) = \mathbb{E} \left[ \left( \int_0^t (e^{B_s} - X_s) dB_s \right)^2 \right] \stackrel{(*)}{=} \mathbb{E} \left[ \int_0^t (e^{B_s} - X_s)^2 ds \right] = \sum_{k=1}^n \int_{t_{k-1}}^{t_k} \mathbb{E}[(e^{B_s} - e^{B_{t_{k-1}}})^2] ds,$$

where  $(*)$  holds by the Itô isometry. Now for any  $t_{k-1} \leq s \leq t_k$  we have

$$\begin{aligned} \mathbb{E}[(e^{B_s} - e^{B_{t_{k-1}}})^2] &= \mathbb{E}[e^{2B_{t_{k-1}}} (e^{B_s - B_{t_{k-1}}} - 1)^2] \stackrel{(\circ)}{=} \mathbb{E}[e^{2B_{t_{k-1}}}] \mathbb{E}[(e^{B_s - B_{t_{k-1}}} - 1)^2] = \\ &\mathbb{E}[e^{2B_{t_{k-1}}}] \mathbb{E}[e^{2(B_s - B_{t_{k-1}})} - 2e^{B_s - B_{t_{k-1}}} + 1] \stackrel{(\bullet)}{=} e^{\frac{1}{2} 2^2 t_{k-1}} \left( e^{\frac{1}{2} 2^2 (s - t_{k-1})} - 2e^{\frac{1}{2} (s - t_{k-1})} + 1 \right), \end{aligned}$$

where  $(\circ)$  holds since Brownian motion has independent increments and in  $(\bullet)$  we used the *Help* concerning moment generating functions as well as  $B_{t_{k-1}} \sim \mathcal{N}(0, t_{k-1})$  and  $(B_s - B_{t_{k-1}}) \sim \mathcal{N}(0, s - t_{k-1})$ . Now

$$\begin{aligned} \int_{t_{k-1}}^{t_k} e^{\frac{1}{2} 2^2 t_{k-1}} \left( e^{\frac{1}{2} 2^2 (s - t_{k-1})} - 2e^{\frac{1}{2} (s - t_{k-1})} + 1 \right) ds = \\ e^{2t_{k-1}} \cdot \left[ \frac{1}{2} \left( e^{2(t_k - t_{k-1})} - 1 \right) - 4 \left( e^{\frac{1}{2}(t_k - t_{k-1})} - 1 \right) + (t_k - t_{k-1}) \right], \end{aligned}$$

and putting together all of the above gives the desired (2).