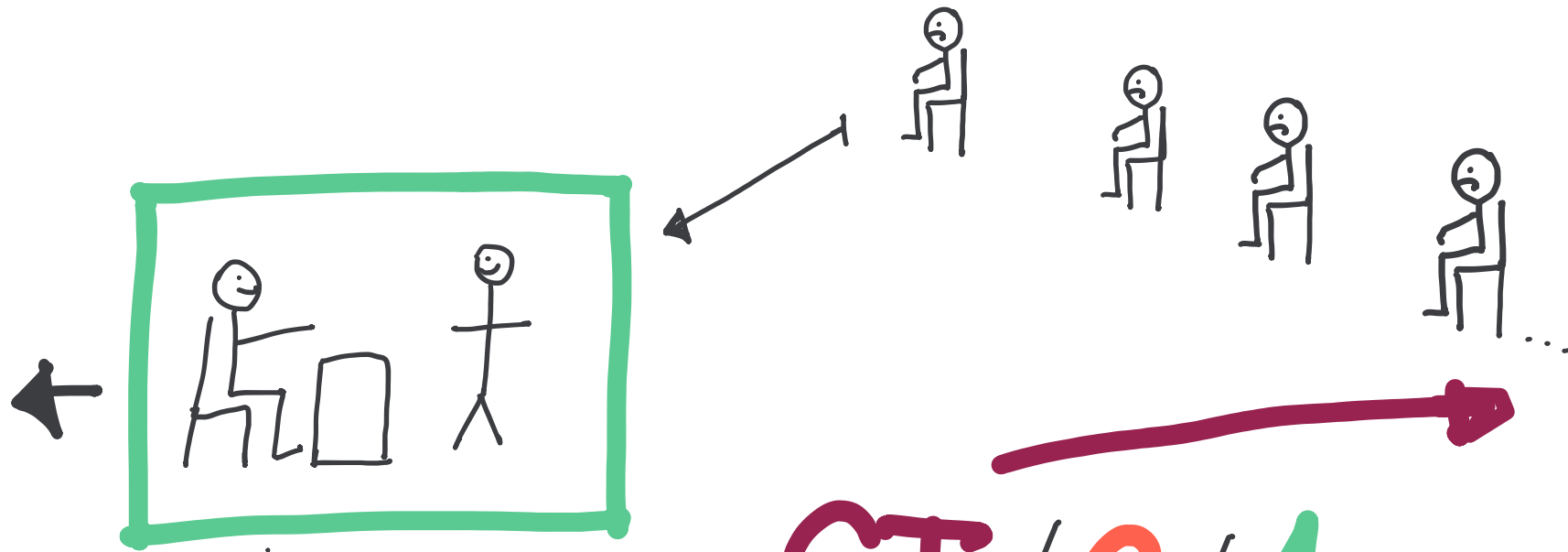


Contents:

- GI/G/1 Queues, Little's Law
- M/M/1 Queues: (∞-waiting room)
 - birth & death process
 - stationary distr.
 - waiting in the queue
 - other properties
- M/M/1 Queue: (finite waiting room)
 - stat. distr. for restricted proc.
 - stat. distr. for M/M/1 queues for N people
 - Barber
- M/M/s Queues:
 - stat. distribution
 - transience, recurrence
- Exercises + Summary

QUEUEING THEORY CONT.



GI / G / 1

↑
SERVICE TIME

$\tau_i \sim G, \text{ i.i.d. } E(G) = 1/\mu$
(rate of service μ)

ARRIVAL



$t_i \sim F, \text{ i.i.d.}$

$E(F) = 1/\lambda$

(rate of arrival is λ)

GI/G/1 Q

X_s = #customers in the system
at time s (people in queue
+ people being served)

L = the avg number of —||— = $\lim_{t \rightarrow \infty} \frac{1}{t} \int_0^t X_s ds$

W_m = time the m -th cust. spends in the system

W = avg. time spent in the system = $\lim_{m \rightarrow \infty} \frac{1}{m} \sum_{n=1}^m W_n$

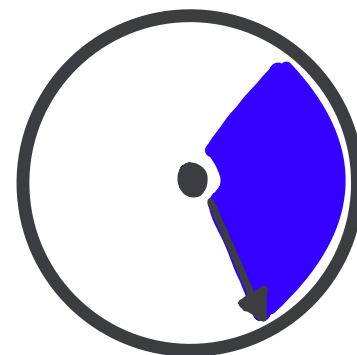
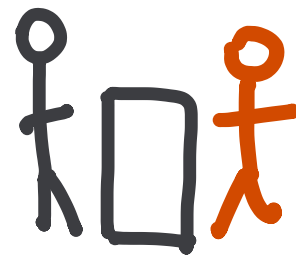
$N_a(t)$ = #people who arrive before t &
stand in the line

$\lambda_a = \lim_{t \rightarrow \infty} \frac{N_a(t)}{t}$

Little's law:

$$L = \lambda_a W$$

2

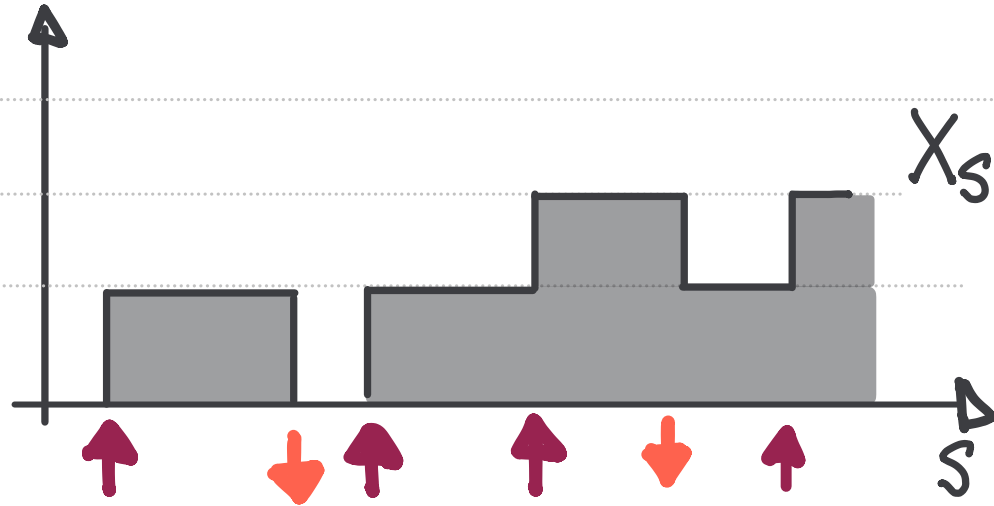


Little's law:

$$L = \lambda_a W$$

$$L = \lim_{t \rightarrow \infty} \frac{1}{t} \cdot \int_0^t X_s ds,$$

$$\lambda_a = \lim_{t \rightarrow \infty} \frac{N_a(t)}{t}, \quad W = \lim_{m \rightarrow \infty} \frac{1}{m} \sum_{k=1}^m W_k$$



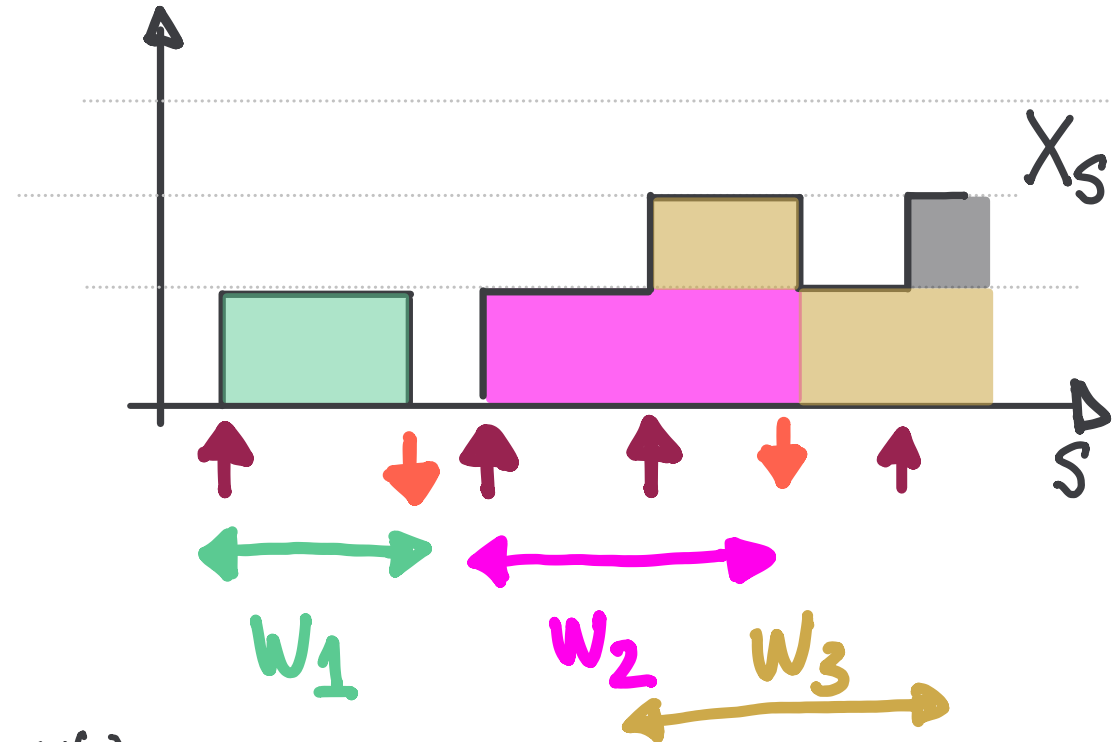
$$\int_0^t X_s ds =$$

$$\frac{1}{t} \int_0^t X_s ds$$

$\approx$

$$\approx \sum_{m=1}^{N(t)} W_m \approx N(t) \cdot W$$

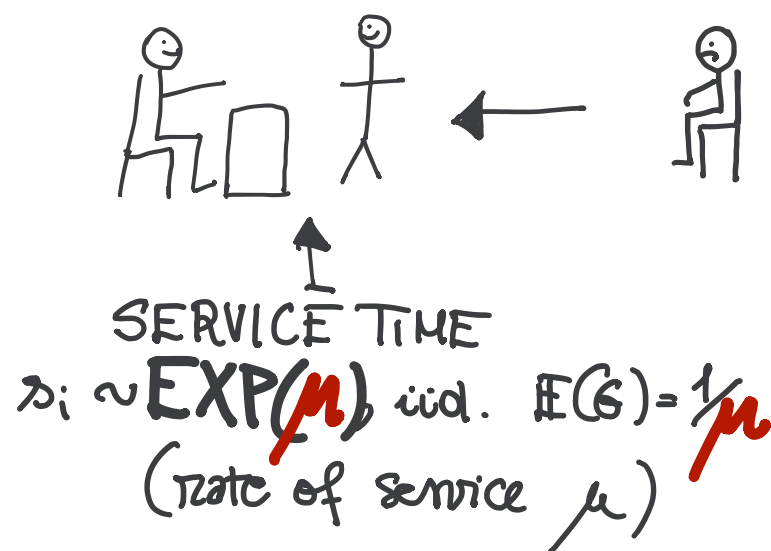
$$\frac{1}{t} N(t) \cdot W \approx \lambda_a \cdot W$$



MARKOVIAN Q-S - M/M/1 QUEUES

1 server

∞ - waiting room



ARRIVAL
according to
Poisson (λ) ($\Rightarrow$ interarrival $\sim \text{EXP}(\lambda)$)
the rate:
 $\lambda = \lim_{t \rightarrow \infty} \frac{N(t)}{t} = \frac{1}{E(G_i)} = \lambda$

Markovian

X_t = number of customers in the system. $\in \{0, 1, 2, \dots\}$ is a
with infinitesimal generator

continuous time
Markov process

$$q(n, n+1) = \lambda, \quad q(n, n-1) = \mu$$

Brief heuristic

X_t - # customers in the system

Markov prop. for $0 \leq s_0 < s_1 < \dots < s_n < \infty$ times, & $i_0, \dots, i_n, i, j$ states

$$\mathbb{P}(X_{t+s} = j \mid X_s = i, \underbrace{X_{s_n} = i_n, \dots, X_{s_0} = i_0}) = \mathbb{P}(X_t = j \mid X_0 = i)$$

does not matter because
of memoryless prop. of
EXP.

Brief heuristic

X_t - # customers in the system

Infinitesimal generator

$$q(i, j) = \lim_{h \rightarrow 0} \frac{\mathbb{P}(X_h = j \mid X_0 = i)}{h} \Rightarrow$$

$$q(n, n+1) = \lambda$$

$$q(n, n-1) = \mu$$

A/ more than 1 departure
in time h

$\Updownarrow$
Sum of two or more exp r.v. $< h$
probability $O(h)$

B/ More than 1 arrival in time h
prob $O(h)$

C/ one arrival, no dep: $(\lambda h e^{-\lambda h} + O(h))e^{-\mu h}$

D/ no arrival, one dep. $(\mu h e^{-\mu h} + O(h))e^{-\lambda h}$

M/M/1 Q-S ARE B&D PROCESSES

$$q(n, n+1) = \lambda$$

BIRTH & DEATH
PROCESS.

(with $\lambda_n = \lambda, \mu_n = \mu$)

$$q(n, n-1) = \mu$$

Assume $\lambda < \mu$ (the arrival rate < the departure rate)

Recall: (Thm. 3.1.)

Let X_t be a birth & death process with $q(n, n+1) = \lambda, q(n, n-1) = \mu$,
with infinite state space.

Then $\pi(n) = \frac{\lambda^n}{\mu^n} \cdot \pi(0)$ satisfies the detailed balanced condition,

Hence if $\sum \pi(i) = 1$ (attainable if $\sum_{n=1}^{\infty} \left(\frac{\lambda}{\mu}\right)^n < \infty$) $\Rightarrow \pi$ is a stationary distr.

M/M/1 Q-S, STATIONARY DISTR.

If $\pi(n) = \left(\frac{\lambda}{\mu}\right)^n \pi(0)$ &

$\sum \pi(i) = 1$ (attainable if $\sum_{n=1}^{\infty} \left(\frac{\lambda}{\mu}\right)^n < \infty$) $\Rightarrow \pi$ is a stationary distr.

$$\sum \pi(i) = \pi(0) \cdot \frac{1}{1 - \lambda/\mu} \Rightarrow \pi(0) = 1 - \lambda/\mu, \pi(n) = (1 - \lambda/\mu) \left(\frac{\lambda}{\mu}\right)^n$$

M/M/1 Q-S, OTHER PROPERTIES

We are in the stationary state

T_Q = time spent in the queue
(r.v.)

$$* T_Q | \{T_Q > 0\} \sim \text{EXP}(\mu - \lambda)$$

$$① \mathbb{P}(T_Q = 0) = \mathbb{P}(Q = 0) = 1 - \frac{\lambda}{\mu}$$

② on $(0, \infty)$ it has a density:
if there is n people already in the system one has to wait $\underbrace{G_1 + \dots + G_n}_{n\text{-service}}$, $G_i \sim \text{EXP}(\mu)$, hence

$$f_{T_Q}(x) = \sum_{n=1}^{\infty} \underbrace{f(x | n \text{ people in the system})}_{\Gamma(n, \mu)} \cdot \underbrace{\mathbb{P}(n \text{ people in the system})}_{\pi(n)} = \sum_{n=1}^{\infty} \frac{\mu^n \cdot x^{n-1} \cdot e^{-\mu x}}{(n-1)!} \cdot \left(\frac{\lambda}{\mu}\right)^n \cdot \left(1 - \frac{\lambda}{\mu}\right) =$$

$$= e^{-\mu x} \lambda \left(1 - \frac{\lambda}{\mu}\right) \sum_{n=1}^{\infty} \frac{(x \lambda)^{n-1}}{(n-1)!} = e^{-\mu x} \lambda \left(1 - \frac{\lambda}{\mu}\right) e^{\lambda x} = e^{-(\mu - \lambda)x} \cdot (\mu - \lambda) \cdot \frac{\lambda}{\mu} \quad \neq$$

M/M/1 Q-S, OTHER PROPERTIES

- $W_Q = \mathbb{E}(T_Q) = \mathbb{E}(T_Q \cdot \mathbb{1}_{\{T_Q > 0\}}) = \frac{1}{\mu} \cdot \frac{1}{\mu - \lambda}$
↑ average waiting time inside the queue

- $\mathbb{E}(W) = W_Q + \frac{1}{\mu} = \frac{1}{\mu - \lambda}$
↑ average overall waiting time (inside queue & in service)

- $L = \lim_{t \rightarrow \infty} \frac{1}{t} \int_0^t X_t dt$, **Little Law:** $L = W \cdot \lambda = \frac{\lambda}{\mu - \lambda}$
↑ the (time) average of the number of customers

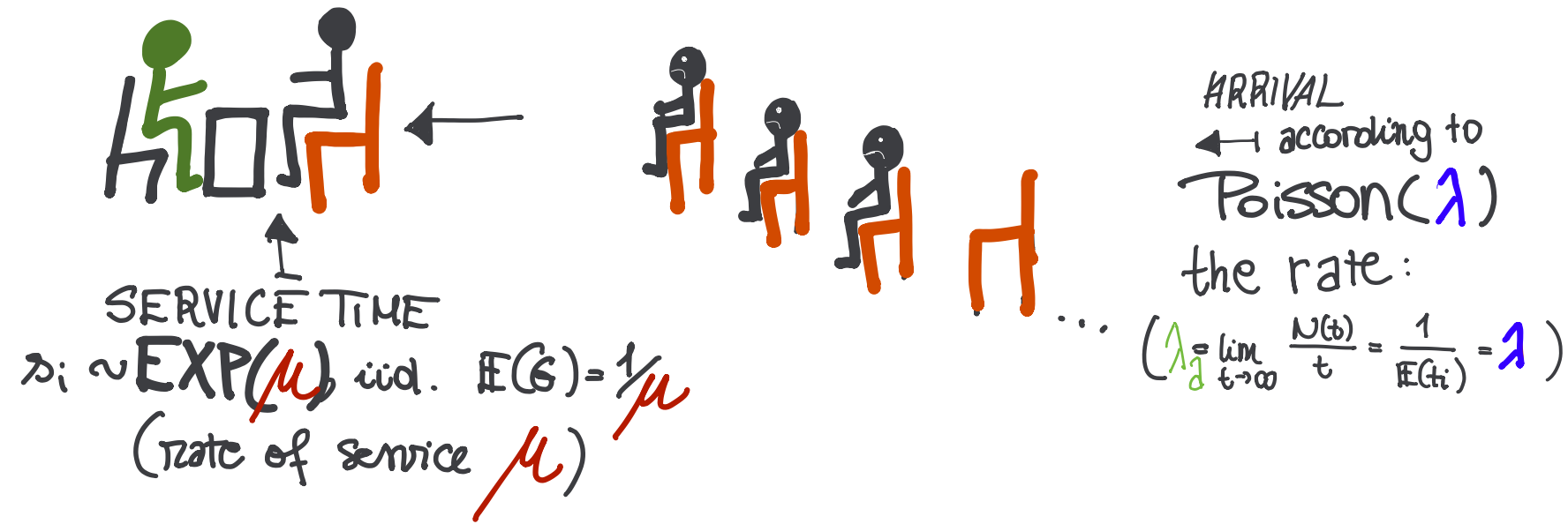
or: $L (= \text{time avg}) = \text{space avg} = \mathbb{E}(X_t) = \sum_{n=0}^{\infty} n \cdot \pi(n) = \frac{\lambda/\mu}{1 - \lambda/\mu} = \frac{\lambda}{\mu - \lambda}$

geom($1 - \lambda/\mu$)

M/M/1 Q-S, FINITE WAITING ROOM

There is only **N** space in the system.

1 server



e.g. barbershop: 2 chairs to wait.

M/M/1 Q-S, FWR, STATIONARY DISTR.

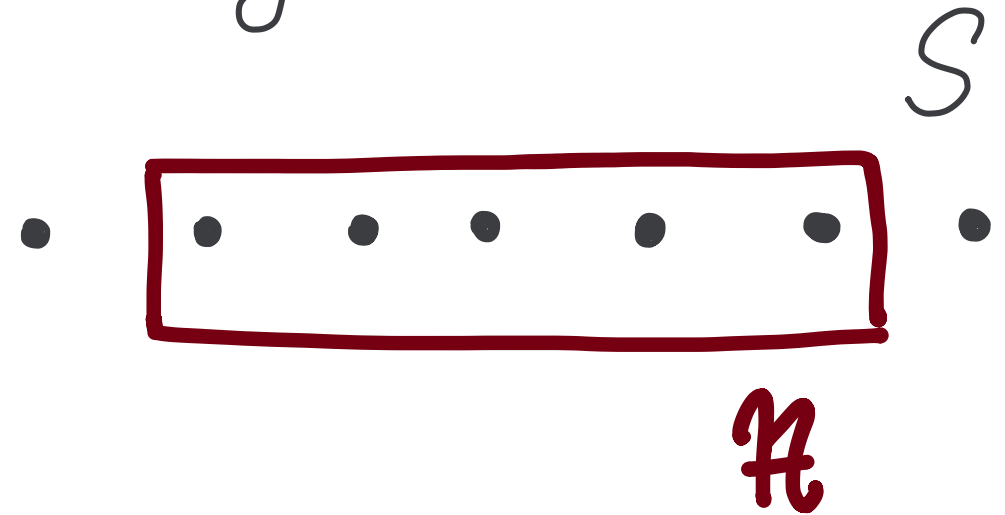
How to get the stationary distr. from the ∞ -waiting room case.

Lemma 4.2.

Given X_t , a M.C. on the state space S , with stationary distr. π satisfying the detailed balance condition.

Let $\{q(i,j)\}$ be the infinitesimal generator of X_t .

Let $A \subseteq S$, Y_t is the restriction of X_t to A .



Then $\underline{\nu} = \frac{1}{\sum_{x \in A} \pi(x)} \cdot \underline{\pi}$ is a stationary distr. for Y_t .

$$\tilde{q}(x,y) = \begin{cases} q(x,y), & \text{if } x,y \in A, \\ 0, & \text{otherwise.} \end{cases}$$

Proof: D.B.C. $x,y \in A$: $\nu(x) \tilde{q}(x,y) = \frac{\pi(x)}{\sum_{z \in A} \pi(z)} \cdot q(x,y) = \frac{\pi(y)}{\sum_{z \in A} \pi(z)} \cdot q(y,x) = \nu(y) \tilde{q}(y,x)$

$$\& \sum \nu(x) = \frac{\sum_{x \in A} \pi(x)}{\sum_{x \in A} \pi(x)} = 1$$

M/M/1 Q-S, FWR, STATIONARY DISTR.

From prev. lemma + stat. distr. for the ∞ -waiting room.

$$S = \{0, 1, \dots\} \Rightarrow \mathcal{H} = \{0, \dots, N\}$$

$$\pi(i) = (1 - \lambda/\mu) \left(\frac{\lambda}{\mu}\right)^i \quad i \in S$$

$$\Rightarrow \nu(i) = \frac{1 - \lambda/\mu}{1 - (\lambda/\mu)^{N+1}} \left(\frac{\lambda}{\mu}\right)^i, \quad i \in \{0, \dots, N\}$$

$$\sum_{k=0}^N (1 - \lambda/\mu) \left(\frac{\lambda}{\mu}\right)^k = (1 - \lambda/\mu) \frac{1 - (\lambda/\mu)^{N+1}}{1 - \lambda/\mu}$$

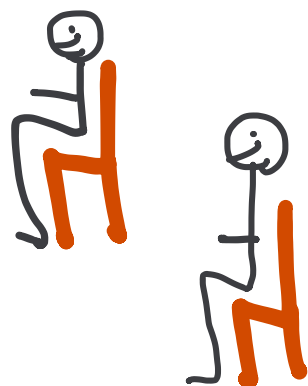
$$\underline{\nu} = \frac{1}{\sum_{x \in \mathcal{H}} \pi(x)} \cdot \underline{\pi}$$

$$\pi(i) = (1 - \lambda/\mu) \left(\frac{\lambda}{\mu}\right)^i$$

Remark. On this finite state space, this is a stat distr. even if $\lambda > \mu$.

For $\lambda = \mu$ we have: $\pi(i) \cdot \lambda = \pi(i+1) \cdot \lambda \Rightarrow \pi(i) = \pi(j) \quad \forall i, j \in \mathcal{H}$
hence $\pi(i) = \frac{1}{N+1} \quad \forall i \in \mathcal{H}$.
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BACK TO THE BARBER



waiting room: 3 chairs ($N=3$)
 Service: $\text{EXP}(3)$ ($\mu=3$)
 arrival: Poisson(2) ($\lambda=2$)

Little $\lambda_a W$:

$$L = \lambda_a \cdot W$$

$$W = \frac{66}{114} \quad \left(\begin{array}{l} \text{I would say} \\ \text{its about half} \\ \text{an hour} \end{array} \right)$$

$$\text{Stat. distr.: } \pi(i) = \frac{1 - \frac{\lambda}{\mu}}{1 - (\frac{\lambda}{\mu})^{N+1}} \left(\frac{\lambda}{\mu} \right)^i = \frac{1 - \frac{2}{3}}{1 - (\frac{2}{3})^4} \cdot \left(\frac{2}{3} \right)^i \Rightarrow \left(\frac{27}{65}, \frac{18}{65}, \frac{12}{65}, \frac{8}{65} \right) = \underline{\underline{\pi}}$$

$$L = \mathbb{E}_{\pi}(X_t) = \frac{18}{65} + 2 \cdot \frac{12}{65} + 3 \cdot \frac{8}{65} = \frac{66}{65}$$

$$\lambda_a = \text{rate of arrival} = \lim_{t \rightarrow \infty} \frac{N_a(t)}{t}$$

ratio of time
when the
waiting room
is full

$$\hookrightarrow \text{heuristics for finding } \lambda_a: \lambda_a = \pi(3) \cdot 0 + (1 - \pi(3)) \cdot \lambda = \frac{114}{65}$$

ratio of time, when its not!

BACK TO THE BARBER

average waiting time again:

$$W = \frac{1}{1 - \pi(3)} \left[\overset{\text{no waiting}}{\pi(0)} \cdot \frac{1}{3} + \pi(1) \cdot \frac{2}{3} + \pi(2) \cdot \frac{3}{3} \right] = \frac{66}{114}$$

↑
one can step inside
the line

↑ wait for **1**
haircut to finish
 $\frac{1}{3} = \frac{1}{\mu}$ amount of time
expectedly

wait for **2**
haircut to finish
 $\frac{1}{3} = \frac{1}{\mu}$ amount of time
expectedly

MINIS Q-S, STATIONARY DISTR.

For $\lambda < s \cdot \mu$ $\exists \pi$ stationary distribution, satisfying the detailed balance condition.

Proof.

d.b.c.

$$\pi(n) q(n, n+1) = \pi(n+1) q(n+1, n)$$

A/ $n+1 \leq s$

$$\pi(n) \cdot \lambda = \pi(n+1) \cdot (n+1) \cdot \mu \Rightarrow \pi(n+1) = \frac{\pi(n)}{n+1} \cdot \frac{\lambda}{\mu}$$

B/ $n+1 > s$

$$\pi(n) \cdot \lambda = \pi(n+1) \cdot s \cdot \mu \Rightarrow \pi(n+1) = \frac{\pi(n)}{s} \cdot \frac{\lambda}{\mu}$$

$$q(n, n+1) = \lambda$$

Keep in mind

$$q(n, n-1) = \begin{cases} n \cdot \mu, & 1 \leq n \leq s \\ s \cdot \mu & n \geq s \end{cases}$$

$\lambda < \infty \Leftrightarrow \lambda < \mu \cdot s \Leftrightarrow$
we have a stationary distr.

$$\pi(i) = \begin{cases} \left(\frac{\lambda}{\mu}\right)^k \cdot \frac{1}{k!} \cdot \frac{1}{s^k}, & k \leq s \\ \left(\frac{\lambda}{\mu}\right)^k \cdot \frac{1}{s! s^{k-s}} \cdot \frac{1}{s^k}, & k > s \end{cases}$$

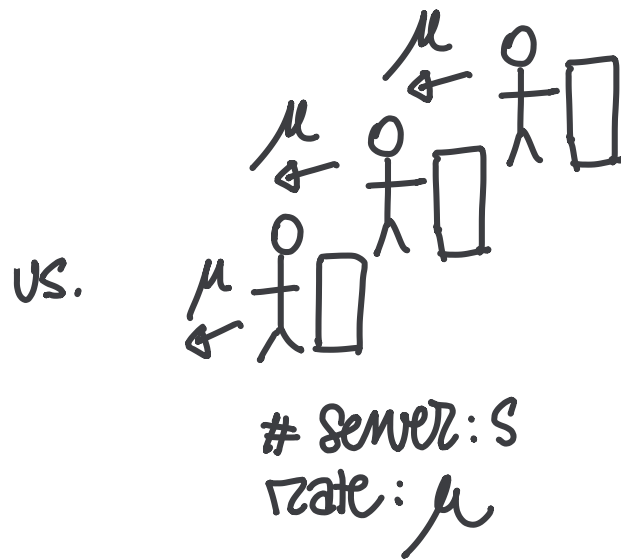
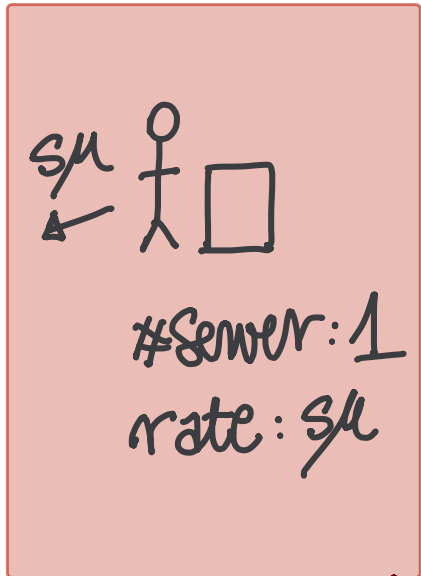
$$\sum_{k=0}^{\infty} \pi(i) = \pi(0) \sum_{k=0}^{s-1} \left(\frac{\lambda}{\mu}\right)^k \frac{1}{k!} + \pi(0) \sum_{k=0}^{\infty} \left(\frac{\lambda}{\mu}\right)^{s+k} \frac{1}{s!} \frac{1}{s^k} = \pi(0) \cdot e$$

M/M/S Q-S, TRANSIENCE & REC.

if $\lambda < s \cdot \mu \Rightarrow$ the chain is recurrent

if $\lambda > s \cdot \mu \Rightarrow$ the chain is transient

Why? Compare with the M/M/1 queue with 1 server with serving rate $s \cdot \mu$. This system is more efficient whenever not all servers are serving.



(e.g. 1 person is served at $\frac{1}{s \cdot \mu}$ or $\frac{1}{\mu}$ time)

If $\lambda > s \cdot \mu \Rightarrow$ the M/M/1 queue is already TRANSIENT, and it is more efficient than the M/M/s one.

more efficient

EXERCISES

1. Find the stationary distr. for the M/M/S queue for $\mu=1, \lambda=2, S=3$
2. — " — for the M/M/1 queue for $\mu=3, \lambda=2$.
3. Find the stationary distr. for the M/M/1 queue for $\mu=3, \lambda=2$, where only 5 people can be in the system.

Summary of stat. distr.

M/M/1 QUEUE
(∞ -waiting room)

$$\lambda < \mu: \pi(i) = (1 - \frac{\lambda}{\mu}) (\frac{\lambda}{\mu})^i$$

$$\lambda \geq \mu: \nexists$$

M/M/1 QUEUE, at most N ppl. in the system:

$$\lambda \neq \mu: \pi(i) = \frac{1 - \frac{\lambda}{\mu}}{1 - (\frac{\lambda}{\mu})^{N+1}} (\frac{\lambda}{\mu})^i$$

$$i \in \{0, \dots, N\}$$

$$\lambda = \mu: \pi(i) = \frac{1}{N+1}$$

$$i \in \{0, \dots, N\}$$

M/M/S Queue
(∞ waiting room)
 $\lambda < S \cdot \mu$:

$$\pi(i) = \begin{cases} (\frac{\lambda}{\mu})^k \cdot \frac{1}{k!} \cdot \frac{1}{S}, & k \leq S \\ (\frac{\lambda}{\mu})^k \cdot \frac{1}{S! S^{k-S}} \cdot \frac{1}{S}, & k > S \end{cases}$$

$$\rho = \sum_{k=0}^{S-1} (\frac{\lambda}{\mu})^k \frac{1}{k!} + (\frac{\lambda}{\mu})^S \frac{1}{S!} \sum_{k=0}^{\infty} (\frac{\lambda}{\mu S})^k$$

1. Find the stationary distr. for the M/M/S queue for $\mu=1, \lambda=2, S=3$

Recall.

$$\pi(k) = \begin{cases} \left(\frac{\lambda}{\mu}\right)^k \cdot \frac{1}{k!} \tilde{\rho}, & k \leq s \\ \left(\frac{\lambda}{\mu}\right)^k \cdot \frac{1}{s! s^{k-s}} \cdot \tilde{\rho}, & k > s \end{cases} \quad \text{where } \tilde{\rho} = \sum_{k=0}^{s-1} \left(\frac{\lambda}{\mu}\right)^k \frac{1}{k!} + \sum_{k=0}^{\infty} \left(\frac{\lambda}{\mu}\right)^{s+k} \frac{1}{s! s^k}$$

why? $\pi(0) \cdot \tilde{\rho} = \sum_{k=0}^{\infty} \pi(k) = 1$,
 $\frac{1}{\tilde{\rho}} = \pi(0)$

Find $\tilde{\rho}$: $1 + \frac{2}{1} + \frac{2^2}{2!} + \frac{2^3}{3!} + \sum_{k=0}^{\infty} \left(\frac{2}{3}\right)^k = 5 + \frac{2^3}{3!} \cdot 3 = 9$

$\pi(0) = \frac{1}{\tilde{\rho}} = \frac{1}{9}$, $\pi(1) = \frac{2}{9}$, $\pi(2) = \left(\frac{2}{3}\right)^2 \cdot \frac{1}{2!}$, for $k > 2$: $\pi(k) = \left(\frac{2}{3}\right)^k \cdot \frac{1}{3!}$

2. Find the stationary distr. for the M/M/1 queue, for $\mu=3, \lambda=2$

$$\pi(k) = \left(1 - \frac{2}{3}\right) \left(\frac{2}{3}\right)^k = \frac{1}{3} \left(\frac{2}{3}\right)^k$$

$$\pi(0) = \frac{1}{3}, \pi(1) = \frac{1}{3} \cdot \frac{2}{3} = \frac{2}{9}, \pi(2) = \frac{4}{27}, \dots$$

3. Find the stationary distr. for the M/M/1 queue for $\mu=3, \lambda=2$, when only 5 people can be in the system.

$$\pi(i) = \frac{1 - \frac{2}{3}}{1 - \left(\frac{2}{3}\right)^6} \left(\frac{2}{3}\right)^i = \frac{\frac{1}{3}}{1 - \frac{64}{729}} \left(\frac{2}{3}\right)^i = \frac{3^5}{665} \left(\frac{2}{3}\right)^i \quad i \in \{0, \dots, 5\}$$