

$$\textcircled{1} \text{ Let } a_n = \frac{1}{\sqrt{n}} \sum_{i=1}^n \sum_{j=1}^n \dots \sum_{k=1}^n \int \sqrt{x_1^2 + \dots + x_n^2} dx_1 dx_2 \dots dx_n = \sum_{i=1}^n \sum_{j=1}^n \dots \sum_{k=1}^n \int \sqrt{\frac{x_1^2 + \dots + x_n^2}{n}} dx_1 dx_2 \dots dx_n.$$

Then  $a_n = \mathbb{E} \sqrt{\frac{X_1^2 + \dots + X_n^2}{n}}$  where  $X_1, \dots, X_n$  are independent  $\mathcal{N}(0, 1)$ .

By the law of large numbers,  $\nu_n = \frac{X_1^2 + \dots + X_n^2}{n} \Rightarrow \mathbb{E} X_i^2 = \int x^2 dx = \frac{1}{3}$ .

The function  $f: [0, 1] \rightarrow \mathbb{R}$   $f(x) = \sqrt{x}$  is bounded and continuous,

$$\text{so } a_n = \mathbb{E} f(\nu_n) \xrightarrow{\text{def}} \mathbb{E} f\left(\frac{1}{3}\right) = \sqrt{\frac{1}{3}}.$$

\textcircled{2} i.) ~~Thm a)~~ If  $X_n$  has char. fn  $\psi_n$ ,  $X$  has char. fn  $\psi$  and  $X_n \Rightarrow X$ , then  $\psi_n(t) \rightarrow \psi(t)$  for every  $t \in \mathbb{R}$ .

b.) If  $X_n$  has char. fn  $\psi_n$  and  $\psi(t) := \lim_{n \rightarrow \infty} \psi_n(t)$  exists for every  $t \in \mathbb{R}$ , AND  $\psi = \psi(t)$  is continuous at 0, ~~then~~  
then  $\exists$  a random variable  $X$  s.t.  $X_n \Rightarrow X$  and  $X$  has char. fn  $\psi$ .

2) Thm (CLT): If  $X_1, \dots, X_n, \dots$  are iid with  $\mathbb{E} X_i = m \in \mathbb{R}$ ,  $\text{Var } X_i = \sigma^2 < \infty$ ,

$$\text{then } \frac{X_1 + \dots + X_n - nm}{\sqrt{n\sigma^2}} \xrightarrow{\text{def}} N(0, 1).$$

Proof: Assume w.l.o.g. that  $m = 0$ . Let  $S_n = X_1 + \dots + X_n$ .

$$\text{Then } \psi_{X_i}(t) = 1 + i \mathbb{E} X_i t + \frac{1}{2} i^2 \mathbb{E}(X_i^2) t^2 + o(t^2) = 1 - \frac{\sigma^2}{2} t^2 + o(t^2)$$

$$\text{so } \psi_{S_n}(t) = (\psi_{X_i}(t))^n = \left(1 - \frac{\sigma^2}{2} t^2 + o(t^2)\right)^n$$

$$\text{so } \psi_{\frac{S_n}{\sqrt{n\sigma^2}}}(t) = \psi_{S_n}\left(\frac{t}{\sqrt{n\sigma^2}}\right) = \left[1 + \frac{-t^2/2}{n} + o\left(\frac{1}{n}\right)\right]^n \xrightarrow{n \rightarrow \infty} e^{-\frac{t^2}{2}}, \text{ which is exactly the char. fn of } N(0, 1), \text{ so the continuity thm gives the statement. } \square$$

③ a.)  $E[X_n^2] < \infty$  by assumption.

~~a)~~  $E[X_n^2] \in \mathcal{G}(\mathcal{Y}_n)$ , so  $X_n^2 \in \mathcal{F}_n$  because  $X_n \in \mathcal{F}_n$

$$E(X_{n+1}^2 | \mathcal{F}_n) \stackrel{\substack{\text{Jensen's} \\ \text{inequality}}}{\geq} \left[ E(X_{n+1} | \mathcal{F}_n) \right]^2 \stackrel{\substack{X_1 \text{ is a} \\ \text{martingale}}}{=} X_n^2 \quad \square$$

b.) Let  $X_n = -\frac{1}{n}$ . Then clearly  $E(X_{n+1} | X_n) = -\frac{1}{n+1} > -\frac{1}{n} = X_n$ .

But  $X_n^2 = \frac{1}{n^2}$  is decreasing.

④ Let  $X_n$  be the position of the frog after  $n$  steps. almost surely  
Then  $X_n$  is a nonnegative martingale  $\Rightarrow$  it is convergent  
by the martingale convergence theorem.

Since it is integer valued, it can only be convergent by being  
eventually constant.

However, the frog can not get stuck at any  $k \neq 0$ , because

$$\sum_{n \geq n_0} P(\text{it jumps away}) = \lim_{n \rightarrow \infty} \frac{k}{k+1} \sum_{n=n_0}^k \frac{1}{n} = \alpha \quad (\text{for } k \neq 0),$$

so the Borel-Cantelli lemma ensures that it would jump away.

Summary: The frog does get stuck, but it can only get  
stuck at 0 so  $P(\text{reaching 0}) = 1$ .

⑤ Let  $S_1, S_2, S_3, \dots$  be independent,  $P(S_i = -1) = \frac{2}{3} =: q$ ,  $P(S_i = +1) = \frac{1}{3} =: p$ .

Let  $S_0 = 0$ ;  $S_n = S_1 + \dots + S_n$ . ( $S_n$  is the simple asymmetric random walk, not trapped).

Then  $M_n := (\frac{q}{p})^{S_n} = 2^{S_n}$  is a martingale adapted to the natural filtration  $F_n := \sigma(S_1, \dots, S_n) =$

- Clearly  $M_n$  is bounded for every  $n$  (takes only finitely many values) so  $E(M_n) < \infty$
- adopted by construction

$$\begin{aligned} E(M_{n+1} | F_n) &= E\left((\frac{q}{p})^{S_{n+1}} | F_n\right) = E\left(M_n \left(\frac{q}{p}\right)^{S_{n+1}} | F_n\right) = \\ &\stackrel{M_n \in F_n}{=} M_n E\left(\left(\frac{q}{p}\right)^{S_{n+1}} | F_n\right) \stackrel{\text{$S_{n+1}$ is independent of } F_n}{=} M_n E\left(\left(\frac{q}{p}\right)^{S_{n+1}}\right) = \\ &= M_n \left[q \left(\frac{q}{p}\right)^{-1} + p \left(\frac{q}{p}\right)^{+1}\right] = M_n \left[q \frac{p}{q} + p \frac{q}{p}\right] = M_n \end{aligned}$$

Let  $T$  be the 1st hitting time of  $\{-10, 10\}$  by  $S_n$ :

$T = \inf\{n \mid S_n \in \{-10, 10\}\}$ . Then  $T$  is a stopping time,

so  $M_{nT}$  is also a martingale, which is also bounded:

$$M_{nT} = 2^{S_{nT}}, \text{ so } 2^{-10} \leq M_{nT} \leq 2^{10}. \text{ The martingale}$$

convergence thm says that  $M_{nT}$  is convergent, so  $P(T < \infty) = 1$ .

Let  $A = \{S_T = -10\}; B = \{S_T = +10\}$ . We saw that  $\boxed{P(A) + P(B) = 1}$ ,

so  ~~$E(S_T = 10)P(A) + 10P(B)$~~ .  $E M_T = P(A) 2^{-10} + P(B) 2^{10}$ .

The optional stopping thm says that  $E M_T = E M_0 = 1$ , so

~~$$\begin{aligned} &E(S_T = 10)P(A) + 10P(B) = 1 \\ &\cancel{E(S_T = 10)P(A)} + \cancel{10P(B)} = 1 \quad \text{simplifying} \\ &10P(B) = 1 \quad \text{canceling terms} \\ &P(B) = \frac{1}{10} \quad \text{dividing by 10} \\ &P(A) = 1 - P(B) = 1 - \frac{1}{10} = \frac{9}{10} \end{aligned}$$~~

③ (continued)

$$\text{So } \begin{cases} P(A) + P(B) = 1 \\ 2^{10}P(A) + 2^{10}P(B) = 1 \end{cases} \xrightarrow{\substack{x 2^{10} \\ \text{copy}}} \begin{cases} 2^{10}P(A) + 2^{10}P(B) = 2^{-10} \\ 2^{10}P(A) + 2^{10}P(B) = 1 \end{cases}$$

Subtract

$$P(B)[2^{10} - 2^{10}] = 2^{-10} - 1$$

$$\Rightarrow P(B) = \frac{1 - 2^{-10}}{2^{10} - 2^{10}} = \frac{2^{10} - 1}{2^{20} - 1} = \frac{2^{10} - 1}{(2^{10} - 1)(2^{10} + 1)} = \underline{\underline{\frac{1}{1025}}}$$