

1st Exercise Class

Convolutions I

1.1 Let X and Y be independent random variables with distribution

(a) $\text{Bin}(n, p)$ and $\text{Bin}(m, p)$ respectively, where $0 < p < 1$ and $n, m \in \mathbb{N}$;

What is the distribution of $X + Y$?

Solution

(a) For $0 \leq k \leq n + m$

$$\begin{aligned} \mathbb{P}(X + Y = k) &= \sum_{l=0}^{\infty} \mathbb{P}(X = k - l, Y = l) \\ &= \sum_{l=0}^k \mathbb{P}(X = k - l) \mathbb{P}(Y = l) \\ &= \sum_{l=0}^k \binom{n}{k-l} p^{k-l} (1-p)^{n-(k-l)} \binom{m}{l} p^l (1-p)^{m-l} \\ &= p^k (1-p)^{(n+m)-k} \sum_{l=0}^k \binom{n}{k-l} \binom{m}{l} \\ &= p^k (1-p)^{(n+m)-k} \binom{n+m}{k}. \end{aligned}$$

Therefore, $X + Y \sim \text{Bin}(n + m, p)$.

1.4 Let X and Y be independent random variables with distribution $\text{Exp}(\lambda)$ and $\text{Exp}(\mu)$. Find the density of $Z := X + Y$.

Solution The PDFs are $f(x) = \lambda e^{-\lambda x} \mathbb{1}_{[0, \infty)}(x)$ and $g(x) = \mu e^{-\mu x} \mathbb{1}_{[0, \infty)}(x)$. Let $h(x)$ denote the density of Z . Then for $x \leq 0$ we have $h(x) = 0$ and for $x > 0$

$$h(x) = \int_{-\infty}^{\infty} f(y)g(x-y) dy = \int_0^x f(y)g(x-y) dy = \lambda\mu e^{-\mu x} \int_0^x e^{-(\lambda-\mu)y} dy$$

Let us consider the following two cases.

(i) $\lambda = \mu$

$$h(x) = \lambda^2 e^{-\lambda x} \int_0^x 1 dy = \lambda^2 x e^{-\lambda x}$$

(ii) $\lambda \neq \mu$

$$h(x) = \lambda\mu e^{-\mu x} \int_0^x e^{-(\lambda-\mu)y} dy = \lambda\mu e^{-\mu x} \left[\frac{e^{-(\lambda-\mu)y}}{\mu - \lambda} \right]_{y=0}^{y=x} = \frac{\lambda\mu}{\mu - \lambda} (e^{-\lambda x} - e^{-\mu x})$$

1.8 Let X_1, X_2 and X_3 be i.i.d random variables with distribution $\text{Uni}(0, 1)$. Find the density functions of the random variables $Y := X_1 + X_2$ and $Z := X_1 + X_2 + X_3$.

Solution The density of a random variable with $\text{Uni}(0, 1)$ distribution is $f(x) = \mathbb{1}_{[0,1]}$. Let $g(x)$ denote the density of Y . Then

$$g(x) = \int_{-\infty}^{\infty} f(y)f(x-y) dy$$

Clearly for any $x < 0$ and $2 \leq x$ we have $g(x) = 0$. Considering the following cases

(i) $0 \leq x < 1$

$$g(x) = \int_0^x 1 dy = x$$

(ii) $1 \leq x < 2$

$$g(x) = \int_{x-1}^1 1 dy = 2 - x$$

We concluded

$$g(x) = x \mathbb{1}_{[0,1)}(x) + (2 - x) \mathbb{1}_{[1,2)}(x).$$

Clearly for any $x \leq 0$ and $3 \leq x$ we have $h(x) = 0$. Now let $h(x)$ denote the density of Z . Then we have

$$h(x) = \int_{-\infty}^{\infty} g(y) f(x - y) dy$$

Considering the following cases

(i) $0 \leq x < 1$

$$h(x) = \int_0^x g(y) dy = \int_0^x y dy = \frac{x^2}{2}$$

(ii) $1 \leq x < 2$

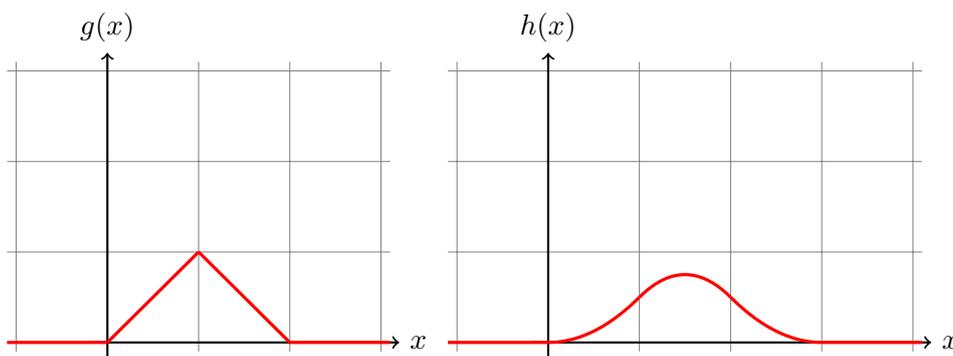
$$\begin{aligned} h(x) &= \int_{x-1}^x g(y) dy = \int_{x-1}^1 g(y) dy + \int_1^x g(y) dy \\ &= \int_{x-1}^1 y dy + \int_1^x (2 - y) dy = 3x - x^2 - \frac{3}{2} \end{aligned}$$

(iii) $2 \leq x < 3$

$$h(x) = \int_{x-1}^2 g(y) dy = \int_{x-1}^2 (2 - y) dy = \frac{9}{2} - 3x + \frac{x^2}{2}$$

We concluded

$$h(x) = \frac{x^2}{2} \mathbb{1}_{[0,1)}(x) + \left(3x - x^2 - \frac{3}{2}\right) \mathbb{1}_{[1,2)}(x) + \left(\frac{9}{2} - 3x + \frac{x^2}{2}\right) \mathbb{1}_{[2,3)}(x)$$



1.9 Let $X_1, X_2, \dots, X_n, \dots$ be i.i.d random variables with distribution $\text{Uni}(0, 1)$. Denote $f_n(x)$ the density function of $S_n := \sum_{k=1}^n X_k$. Prove that

$$f_n(x) = \frac{1}{(n-1)!} \sum_{k=0}^{\lfloor x \rfloor} (-1)^k \binom{n}{k} (x-k)^{n-1}.$$

Using a computer program (e.g. *Mathematica*, *Python*) plot the graph of the function

$$\tilde{f}_n(x) := \sqrt{\frac{n}{12}} f_n\left(\frac{n}{2} + \sqrt{\frac{n}{12}} x\right)$$

for $n = 1, 2, \dots, 10$. What do we see? Interpret the result!**Solution** We will prove it by induction on $n \in \mathbb{N}^+$. Clearly for $n = 1$ we have

$$f_1(x) = \mathbb{1}_{[0,1)}(x),$$

which is exactly the density of $S_1 = X_1 \sim \text{Uni}(0, 1)$.

Now suppose it holds for some $n \in \mathbb{N}^+$. Then

$$f_{n+1}(x) = \int_{-\infty}^{\infty} f_n(x-y) \mathbb{1}_{[0,1)}(y) dy = \int_0^1 f_n(x-y) dy$$

Thus by the induction hypothesis

$$f_{n+1}(x) = \int_0^1 \frac{1}{(n-1)!} \sum_{k=0}^{[x-y]} (-1)^k \binom{n}{k} (x-y-k)^{n-1} dy =$$

$$\underbrace{\int_0^{\{x\}} \frac{1}{(n-1)!} \sum_{k=0}^{[x]} (-1)^k \binom{n}{k} (x-y-k)^{n-1} dy}_{I_1} + \underbrace{\int_{\{x\}}^1 \frac{1}{(n-1)!} \sum_{k=0}^{[x]-1} (-1)^k \binom{n}{k} (x-y-k)^{n-1} dy}_{I_2}.$$

Where

$$I_1 = \frac{1}{(n-1)!} \sum_{k=0}^{[x]} (-1)^k \binom{n}{k} \left[-\frac{(x-y-k)^n}{n!} \right]_{y=0}^{y=\{x\}}$$

$$= \frac{1}{n!} \sum_{k=0}^{[x]} (-1)^{k+1} \binom{n}{k} (x - \{x\} - k)^n - \frac{1}{n!} \sum_{k=0}^{[x]} (-1)^{k+1} \binom{n}{k} (x - k)^n$$

and

$$I_2 = \frac{1}{(n-1)!} \sum_{k=0}^{[x]-1} (-1)^k \binom{n}{k} \left[-\frac{(x-y-k)^n}{n!} \right]_{y=\{x\}}^1$$

$$= \frac{1}{n!} \sum_{k=0}^{[x]-1} (-1)^{k+1} \binom{n}{k} (x - 1 - k)^n - \frac{1}{n!} \sum_{k=0}^{[x]-1} (-1)^{k+1} \binom{n}{k} (x - \{x\} - k)^n.$$

Then

$$f_{n+1}(x) = I_1 + I_2$$

$$= \frac{1}{n!} (-1)^{[x]+1} \underbrace{(x - \{x\} - [x])^n}_{=0}$$

$$+ \frac{1}{n!} \sum_{k=0}^{[x]-1} (-1)^{k+1} \binom{n}{k} (x - 1 - k)^n - \frac{1}{n!} \sum_{k=0}^{[x]} (-1)^{k+1} \binom{n}{k} (x - k)^n$$

$$= \frac{1}{n!} \sum_{k=1}^{[x]} (-1)^k \binom{n}{k-1} (x - k)^n + \frac{1}{n!} \sum_{k=0}^{[x]} (-1)^k \binom{n}{k} (x - k)^n$$

$$= \frac{x^n}{n!} + \frac{1}{n!} \sum_{k=1}^{[x]} (-1)^k \left(\binom{n}{k-1} + \binom{n}{k} \right) (x - k)^n$$

$$= \frac{x^n}{n!} + \frac{1}{n!} \sum_{k=1}^{[x]} (-1)^k \binom{n+1}{k} (x - k)^n$$

$$= \frac{1}{n!} \sum_{k=0}^{[x]} (-1)^k \binom{n+1}{k} (x - k)^n.$$

Therefore, the formula also holds for $n + 1$.

1.10 Let X and Y be independent r.v. with dist. $\text{Poi}(\lambda)$ and $\text{Uni}(0, 1)$. Find the distribution of $Z := X + Y$.

Solution Let F denote the CDF of X and g the density function of Y . Then for $z \geq 0$

$$\begin{aligned}
 \mathbb{P}(Z \leq z) &= \int_{-\infty}^{\infty} F(z-x) \cdot g(x) \, dx \\
 &= \int_0^1 F(z-x) \, dx \\
 &= \int_0^1 \mathbb{P}(X \leq z) \, dx \\
 &= \int_0^1 \sum_{k=0}^{\lfloor z-x \rfloor} e^{-\lambda} \frac{\lambda^k}{k!} \, dx \\
 &= \int_0^{\{z\}} \sum_{k=0}^{\lfloor z-x \rfloor} e^{-\lambda} \frac{\lambda^k}{k!} \, dx + \int_{\{z\}}^1 \sum_{k=0}^{\lfloor z-x \rfloor} e^{-\lambda} \frac{\lambda^k}{k!} \, dx \\
 &= \int_0^{\{z\}} \sum_{k=0}^{\lfloor z \rfloor} e^{-\lambda} \frac{\lambda^k}{k!} \, dx + \int_{\{z\}}^1 \sum_{k=0}^{\lfloor z \rfloor - 1} e^{-\lambda} \frac{\lambda^k}{k!} \, dx \\
 &= \{z\} \cdot \sum_{k=0}^{\lfloor z \rfloor} e^{-\lambda} \frac{\lambda^k}{k!} + (1 - \{z\}) \cdot \sum_{k=0}^{\lfloor z \rfloor - 1} e^{-\lambda} \frac{\lambda^k}{k!} \\
 &= \{z\} \cdot e^{-\lambda} \frac{\lambda^{\lfloor z \rfloor}}{\lfloor z \rfloor!} + \sum_{k=0}^{\lfloor z \rfloor - 1} e^{-\lambda} \frac{\lambda^k}{k!}.
 \end{aligned}$$