

13th Practice class

Weak convergence of probability distributions

13.1 Let $S = [0, 1]$ and let μ_n be the discrete measure such that every point k/n has weight $1/(n+1)$ for $k = 0, 1, \dots, n$. Show that $\mu_n \Rightarrow \mu$, where μ is the Lebesgue measure on S . (Use the definition of the weak convergence, not the equivalent formalisations for distribution functions.)

Solution Let $f : S \mapsto \mathbb{R}$ be bounded and continuous. Then f is Riemann integrable on $[0, 1]$. Hence,

$$\begin{aligned} \int_S f(x) d\mu_n(x) &= \sum_{k=0}^n f\left(\frac{k}{n}\right) \cdot \mu_n\left(\left\{\frac{k}{n}\right\}\right) = \sum_{k=0}^n f\left(\frac{k}{n}\right) \frac{1}{n+1} \\ &= \frac{n}{n+1} \cdot \sum_{k=0}^n f\left(\frac{k}{n}\right) \frac{1}{n} \rightarrow \int_0^1 f(x) dx = \int_S f(x) d\mu(x). \end{aligned}$$

Since every Riemann integrable function is Lebesgue integrable and the integrals are the same, the last equality holds.

Remark. We actually showed that if

$$X_n \sim \text{Uni}\left\{0, \frac{1}{n}, \frac{2}{n}, \dots, \frac{n-1}{n}, 1\right\} \quad \text{and} \quad X \sim \text{Uni}(0, 1),$$

then

$$X_n \xrightarrow{d} X.$$

13.4 Let $c \in \mathbb{R}$ and X_n be a sequence of random variables defined on the same probability space $(\Omega, \mathcal{F}, \mathbb{P})$. Show that the following are equivalent.

(i) $X_n \xrightarrow{\mathbb{P}} c$ as $n \rightarrow \infty$.

(ii) $X_n \xrightarrow{d} c$ as $n \rightarrow \infty$.

Solution

(i) $\Rightarrow$ (ii) $X_n \xrightarrow{\mathbb{P}} c$ implies $X_n \xrightarrow{d} c$ in general.

(ii) $\Rightarrow$ (i) Let F be the distribution function of the constant c random variable. Then

$$F(x) = \begin{cases} 0 & \text{if } x \leq c, \\ 1 & \text{if } c < x. \end{cases}$$

Let $\delta > 0$. Since $c - \delta$ and $c + \delta$ are points of continuity of F , by weak convergence

$$\mathbb{P}(|X_n - c| > \delta) = \mathbb{P}(X_n < c - \delta) + \mathbb{P}(X_n > c + \delta) \rightarrow F(c - \delta) + 1 - F(c + \delta) = 0 + 1 - 1 = 0.$$

13.5 (Cramer-Slutsky theorem) Let $X_n, Y_n, Z_n, n \in \mathbb{N}$ and X be random variables over the same probability space $(\Omega, \mathcal{F}, \mathbb{P})$. Show if $X_n \xrightarrow{d} X$ and $Y_n \xrightarrow{\mathbb{P}} y, Z_n \xrightarrow{\mathbb{P}} z$ for some $y, z \in \mathbb{R}$ then $X_n Y_n + Z_n \xrightarrow{d} X \cdot y + z$.

Solution

(i) We first show that

$$X_n + Z_n \xrightarrow{d} X + z.$$

Let

$$F(x) = \mathbb{P}(X + z < x) \quad \text{and} \quad F_n(x) = \mathbb{P}(X_n + Z_n < x)$$

It is enough to show that for any $x \in \mathbb{R}$ and $\varepsilon > 0$

$$F(x - \varepsilon) \leq \liminf_{n \rightarrow \infty} F_n(x) \leq \limsup_{n \rightarrow \infty} F_n(x) \leq F(x + \varepsilon).$$

For the upper bound, let $0 < \delta_1 \leq \varepsilon$ be such that

$$\mathbb{P}(X + z = x + \delta_1) = 0.$$

Since the distribution of X has at most countably many discontinuity points, there is such a δ_1 . Then we have the bound

$$\begin{aligned} F_n(x) &= \mathbb{P}(X_n + Z_n < x) \\ &= \mathbb{P}(X_n + Z_n < x, |Z_n - z| > \delta_1) + \mathbb{P}(X_n + Z_n < x, |Z_n - z| \leq \delta_1) \\ &\leq \underbrace{\mathbb{P}(|Z_n - z| > \delta_1)}_{\rightarrow 0} + \underbrace{\mathbb{P}(X_n + z < x + \delta_1)}_{\rightarrow F(x + \delta_1)}. \end{aligned}$$

Hence,

$$\limsup_{n \rightarrow \infty} F_n(x) \leq F(x + \delta_1) \leq F(x + \varepsilon).$$

For the lower bound, let $0 < \delta_2 \leq \varepsilon$ be such that

$$\mathbb{P}(X + z = x - \delta_2) = 0.$$

Since the distribution of X has at most countably many discontinuity points, there is such a δ_2 . Then we have the bound

$$\begin{aligned} 1 - F_n(x) &= \mathbb{P}(X_n + Z_n \geq x) \\ &= \mathbb{P}(X_n + Z_n \geq x, |Z_n - z| > \delta_2) + \mathbb{P}(X_n + Z_n \geq x, |Z_n - z| \leq \delta_2) \\ &\leq \mathbb{P}(|Z_n - z| > \delta_1) + \mathbb{P}(X_n + z \geq x - \delta_2) \\ &= \mathbb{P}(|Z_n - z| > \delta_1) + 1 - \mathbb{P}(X_n + z < x - \delta_2). \end{aligned}$$

After rearranging

$$\underbrace{\mathbb{P}(X_n + z < x - \delta_2)}_{\rightarrow F(x - \delta_2)} - \underbrace{\mathbb{P}(|Z_n - z| > \delta_1)}_{\rightarrow 0} \leq F_n(x).$$

Hence,

$$F(x - \varepsilon) \leq F(x - \delta_2) \leq \liminf_{n \rightarrow \infty} F_n(x).$$

(ii) We now show that

$$X_n \cdot Y_n \xrightarrow{d} y \cdot X.$$

Notice

$$X_n \cdot Y_n = y \cdot X_n + X_n(Y_n - y).$$

Hence, by part (i) it is enough to show that

$$y \cdot X_n \xrightarrow{d} y \cdot X \quad \text{and} \quad X_n(Y_n - y) \xrightarrow{\mathbb{P}} 0.$$

Let $f : \mathbb{R} \rightarrow \mathbb{R}$ be bounded and continuous. Then $x \mapsto f(y \cdot x)$ is also bounded and continuous. Thus,

$$\mathbb{E}(f(y \cdot X_n)) \rightarrow \mathbb{E}(f(y \cdot X)).$$

This exactly means

$$y \cdot X_n \xrightarrow{d} y \cdot X.$$

We know that $(X_n)_{n=1}^\infty$ is tight. Thus, for any $\varepsilon > 0$ there exists $0 < M < \infty$ such that for all $n \geq 1$

$$\mathbb{P}(|X_n| > M) < \varepsilon.$$

Then for any $\delta > 0$

$$\begin{aligned} \mathbb{P}(|X_n(Y_n - y)| > \delta) &= \mathbb{P}(|X_n(Y_n - y)| > \delta, |X_n| > M) + \mathbb{P}(|X_n(Y_n - y)| > \delta, |X_n| \leq M) \\ &\leq \mathbb{P}(|X_n| > M) + \mathbb{P}\left(|Y_n - y| > \frac{\delta}{M}\right) \\ &\leq \underbrace{\varepsilon}_{\rightarrow 0} + \mathbb{P}\left(|Y_n - y| > \frac{\delta}{M}\right). \end{aligned}$$

Hence, for any $\varepsilon > 0$

$$\limsup_{n \rightarrow \infty} \mathbb{P}(|X_n(Y_n - y)| > \delta) \leq \varepsilon.$$

It follows

$$X_n(Y_n - y) \xrightarrow{\mathbb{P}} 0.$$

Combining (i) and (ii) we get $X_n \cdot Y_n + Z_n \xrightarrow{d} X_n \cdot y + z$.

13.7 (*Limit theorems for maximum of i.i.d random variables*) Let $X_1, X_2, \dots$ be i.i.d random variables with distribution function $F(x) := \mathbb{P}(X_i < x)$. Let $M_n := \max\{X_1, X_2, \dots, X_n\}$. The asymptotic behavior of the random variable M_n as $n \rightarrow \infty$ depends on then asymptotic of the upper tail of $F(x)$. Show the following limit theorems:

- (a) Suppose that $F(x) < 1$ for every $x < \infty$ and $\lim_{x \rightarrow \infty} x^\alpha(1 - F(x)) = b$ for some $\alpha, b \in (0, \infty)$. Show that the distribution of $n^{-1/\alpha}M_n$ converges weakly to:

$$\mathbb{P}(n^{-1/\alpha}M_n < x) \rightarrow \mathbb{1}_{\{x > 0\}} \exp(-bx^{-\alpha}).$$

- (b) Suppose that $F(x_0) = 1$ and $F(x) < 1$ for every $x < x_0$, moreover, $\lim_{x \rightarrow x_0} (x_0 - x)^{-\alpha}(1 - F(x)) = b$ for some $\alpha, b \in (0, \infty)$. Then the distribution of $n^{1/\alpha}(x_0 - M_n)$ converges weakly to:

$$\mathbb{P}(n^{1/\alpha}(x_0 - M_n) < x) \rightarrow \mathbb{1}_{\{x > 0\}} (1 - \exp(-bx^\alpha)).$$

- (c) Suppose that $F(x) < 1$ for every $x < \infty$ and $\lim_{x \rightarrow \infty} e^{\lambda x}(1 - F(x)) = b$ for some $\lambda, b \in (0, \infty)$. Then the distribution of $M_n - \lambda^{-1} \log n$ converges weakly to:

$$\mathbb{P}(M_n - \lambda^{-1} \log n < x) \rightarrow \exp(-be^{-\lambda x}).$$

Solution

- (a) Notice that $x \mapsto \exp(-bx^{-\alpha}) \cdot \mathbb{1}_{\{x > 0\}}$ is continuous at every $x \in \mathbb{R}$. Since the variables are i.i.d.,

$$\begin{aligned} \mathbb{P}(M_n < x) &= \mathbb{P}(\max\{X_1, \dots, X_n\} < x) \\ &= \mathbb{P}(X_1 < x, \dots, X_n < x) \\ &= \mathbb{P}(X_1 < x) \cdots \mathbb{P}(X_n < x) \\ &= F(x)^n. \end{aligned}$$

Hence, for any $x \in \mathbb{R}$

$$\mathbb{P}(n^{-\frac{1}{\alpha}}M_n < x) = \mathbb{P}(M_n < xn^{\frac{1}{\alpha}}) = F(xn^{\frac{1}{\alpha}})^n.$$

If $x \leq 0$, then since $0 \leq F(0) < 1$,

$$0 \leq F(xn^{\frac{1}{\alpha}})^n \leq F(0)^n \rightarrow 0.$$

If $x > 0$, then

$$F(xn^{\frac{1}{\alpha}})^n = (1 - (1 - F(xn^{\frac{1}{\alpha}}))^n) \rightarrow \exp(-bx^{-\alpha}),$$

since

$$\begin{aligned} \lim_{n \rightarrow \infty} n(1 - F(xn^{\frac{1}{\alpha}})) &= \lim_{n \rightarrow \infty} x^{-\alpha} \underbrace{\left(xn^{\frac{1}{\alpha}}\right)^\alpha (1 - F(xn^{\frac{1}{\alpha}}))}_{\rightarrow b, \text{ since } xn^{1/\alpha} \rightarrow \infty \text{ as } n \rightarrow \infty} = bx^{-\alpha}. \end{aligned}$$

Therefore, for any $x \in \mathbb{R}$

$$\lim_{n \rightarrow \infty} \mathbb{P}(n^{-\frac{1}{\alpha}}M_n < x) = \exp(-bx^{-\alpha}) \cdot \mathbb{1}_{\{x > 0\}}.$$